Bayesian analysis of non-Gaussian Long-Range Dependent processes

Tim Graves 1, Christian Franzke 2, Nick Watkins 2, and Bobby Gramacy 3

Statistical Laboratory, University of Cambridge, Cambridge, UK

British Antarctic Survey, Cambridge, UK

Booth School of Business, The University of Chicago, Chicago, USA

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2 Exact Bayesian analysis for Gaussian case

3 Approximate Bayesian analysis for general case





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Long Range Dependence

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... [T]he stationary long memory processes form a layer among the stationary processes that is "near the boundary" with non-stationary processes, or, alternatively, as the layer separating the non-stationary processes from the "usual" stationary processes. [Samorodnitsky, 2006]

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$$X_t = \sum_{k=0}^{\infty} \psi_k \epsilon_{t-k},$$

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ARFIMA processes

Definition

A process $\{X_t\}$ is an ARFIMA(p, d, q) process if it is the solution to:

$$\Phi(B)(1-B)^d X_t = \Theta(B)\varepsilon_t,$$

where $\Phi(z) = 1 + \sum_{j=1}^p \phi_j z^j$ and $\Theta(z) = 1 + \sum_{j=1}^q \theta_j z^j,$

and the innovations $\{\varepsilon_t\}$ are iid with 0 mean and variance $\sigma^2 < \infty$. We say that $\{X_t\}$ is an ARFIMA(p, d, q) process with mean μ , if $\{X_t - \mu\}$ is an ARFIMA(p, d, q) process.

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ARFIMA parameters

• μ – location parameter

- σ scale parameter
- $d \log$ memory parameter (long memory process iff 0 < d < 0.5)
- ϕp -dimensional short memory parameter
- θq -dimensional short memory parameter

Which parameters are of interest?

When considering *long memory* processes, we are usually primarily interested in the parameter d (and possibly μ). The parameters σ, ϕ, θ (and even p, q) are essentially nuisance parameters.

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- Even assuming Gaussianity, the likelihood for *d* is very complex impossible to find analytic posterior
- Must resort to MCMC methods in order to obtain samples from the posterior
- Don't want to assume form of short memory (i.e. p, q) must use Reversible-Jump (RJ) MCMC [Green, 1995]

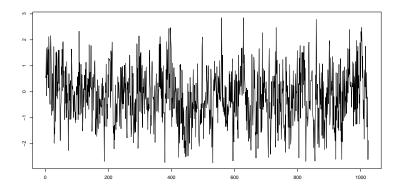
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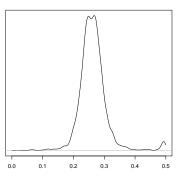
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- Correlation between parameters (e.g. ϕ and d) requires blocking.

Example: 'Pure' Gaussian Long Range Dependence



$$(1-B)^{0.25}X_t = \varepsilon_t$$

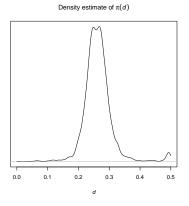
Example: 'Pure' Gaussian Long Range Dependence



Density estimate of $\pi(d)$

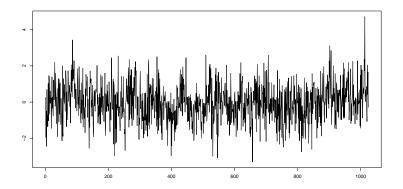
d

Example: 'Pure' Gaussian Long Range Dependence



- \bullet Similarly good results for μ and σ
- The posterior model probability for the (0, d, 0) model was 70%

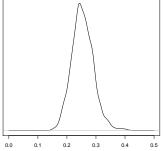
Example: 'Corrupted' Gaussian Long Range Dependence



 $(1+0.75B)(1-B)^{0.25}X_t = (1+0.5B)\varepsilon_t$

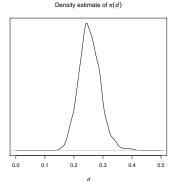
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d

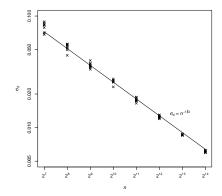
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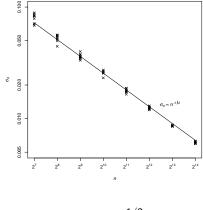
- The posterior model probability for the (1, d, 1) model was 77%
- The posterior model probability for the (0, d, 0) model was 0%

Dependence of posterior variance on n

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 $\sigma_d \propto n^{-1/2}$





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Assumptions and general method

- Drop the Gaussianity assumption
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- Infinite variance means that auto-covariance approach is no longer sound
- Lack of closed form for α -stable density implies lack of closed form for likelihood



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$$f(x_1,\ldots,x_t|\mathcal{H})=f(x_t|x_{t-1},\ldots,x_1,\mathcal{H})f(x_{t-1},\ldots,x_1|\mathcal{H})$$

where \mathcal{H} is the finite recent history of the process $x_0, x_{-1}, \ldots, x_{-n}$

Solution

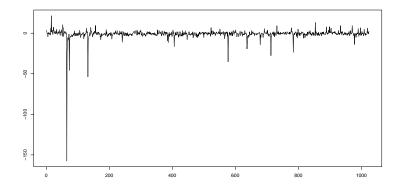
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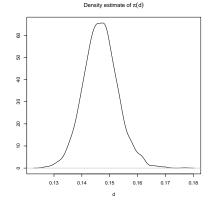
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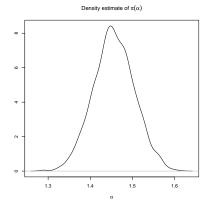
- \bullet Use auxiliary variables to integrate out the (unknown) history ${\cal H}$
- In practice, setting $\mathcal{H} = \bar{x}, \dots, \bar{x}$ suffices, providing enormous computational saving.

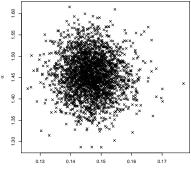
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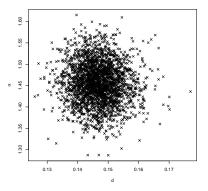






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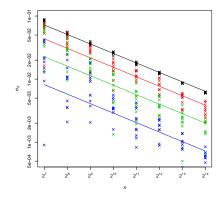
Graves et al. Non-Gaussian Long-Range Dependency



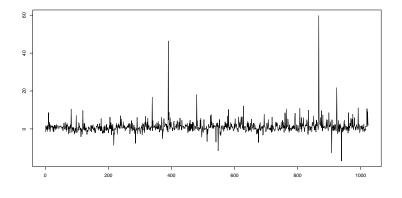
- Good estimation of all parameters
- The posteriors of d and α are independent

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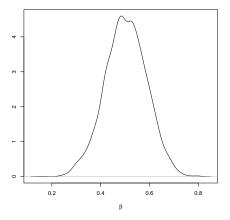


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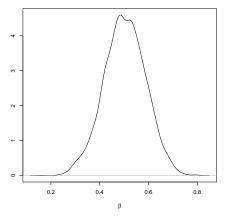


 $(1-B)^{0.1}X_t = \varepsilon_t, \qquad \alpha = 1.5 \qquad \beta = 0.5$

Density estimate of $\pi(\beta)$



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• Good estimation of all other parameters

References

Green, P. J. (1995).

Reversible jump Markov chain Monte Carlo computation and Bayesian model determination.

Biometrika, 82(4), 711–732.

