## CO902 Problem Set 4

- 1. Estimating "a priori" probabilities. This is Bishop Ex 4.9 in my own notation; see B §4.2.2 for hints. Consider a generative classification model for K classes defined by prior class probabilities  $P(Y = k) = \pi_k$  and class-conditional densities  $p_k(\mathbf{X}) =$  $p(\mathbf{X}|Y = k)$ , where  $\mathbf{X} \in \mathbb{R}^d$  is the input feature vector and  $Y \in \{1, 2, \ldots, K\}$  is the true class<sup>1</sup>. For data  $\{\mathbf{x}_i, y_i\}$ ,  $i = 1, \ldots, N$ , considering the joint likelihood for  $\boldsymbol{\pi} = (\pi_1, \ldots, \pi_K)$  and the (unspecified) parameters of the data  $\mathbf{x}$ ; find the maximum likelihood estimate of the class frequencies  $\boldsymbol{\pi}$ . Hint: Express the class variable as  $t_{ik} = \delta_{y_ik}$ , where  $\delta_{y_ik}$  is 1 if  $y_i = k$  and 0 otherwise, so that  $\mathbf{t}_i$  is a K-vector with K-1zeros and 1 one, indicating the true class for observation i.
- 2. Optimal decision rule for continuous data. Consider supervised learning based on  $\{\mathbf{X}_i, Y_i\}, i = 1, ..., n$ , for data  $\mathbf{X}_i \in \Re^d$  and a class membership  $Y_i \in \{1, 2, ..., K\}$ . Show that the optimal decision rule  $D_{\mathbf{x}}$  takes the form

$$D_{\mathbf{x}} = \operatorname*{argmax}_{k} P(Y = k | \mathbf{X} = \mathbf{x})$$

3. Regression. Consider observations of the form  $\{\mathbf{X}_i, Y_i\}$ , i = 1, ..., n, for predictors  $\mathbf{X}_i \in \mathbb{R}^d$  and response  $Y_i \in \mathbb{R}$ . Let  $\mathbf{X}$  be a  $n \times (d+1)$  matrix, where each row consists of  $[1 \mathbf{X}_i^{\top}]$ , and  $\mathbf{Y}$  be the *n*-vector of responses. The linear regression model approximates  $\mathbf{Y}$  with

$$\hat{\mathbf{Y}} = \mathbf{X}\mathbf{w}$$

where  $\mathbf{w}$  is a d + 1 vector of regression coefficients. The standard estimate of  $\mathbf{w}$  is  $\hat{\mathbf{w}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{Y}$ .

(a) Derive  $\hat{\mathbf{w}}$  as the minimizer of the residual sum of squares,

$$J(\mathbf{w}) = (\mathbf{Y} - \mathbf{X}\mathbf{w})^{\top} (\mathbf{Y} - \mathbf{X}\mathbf{w}).$$

- (b) Derive  $\hat{\mathbf{w}}$  on the assumption that  $\mathbf{Y} \sim \mathcal{N}_n(\mathbf{X}\mathbf{w}, \mathbf{I}_n\sigma^2)$ , where  $\mathcal{N}_n$  is a *n*-dimensional multivariate Normal distribution,  $\mathbf{I}_n$  is a  $n \times n$  identity matrix, and  $\sigma^2$  is the resdiual error variance.
- 4. *Ridge Regression.* Consider the same data matrix  $\mathbf{X}$  and response  $\mathbf{Y}$  as in the previous question.
  - (a) The following cost function is the residual sum of squares penalized by the sum of squares of the regression coefficients,

$$J(\mathbf{w}) = (\mathbf{Y} - \mathbf{X}\mathbf{w})^{\top} (\mathbf{Y} - \mathbf{X}\mathbf{w}) + \lambda \mathbf{w}^{\top} \mathbf{w}.$$

<sup>&</sup>lt;sup>1</sup>Here, capital Roman letters indicate (yet to be observed) random variables, while lower case Roman letters indicate particular (observed) values of the random variables. Boldface font indicates a vector-valued variable.

Show that the ridge regression estimator  $\hat{\mathbf{w}} = (\mathbf{X}^{\top}\mathbf{X} + \lambda \mathbf{I}_{d+1})^{-1}\mathbf{X}^{\top}\mathbf{Y}$  minimizes this  $J(\mathbf{w})$ .

(b) Derive the ridge regression estimator as the maximum a postiori (MAP) estimator of a Bayesian model with prior

$$\mathbf{w} \sim \mathcal{N}_{d+1}(\mathbf{0}, \mathbf{I}_{d+1}\sigma_0^2),$$

where  $\mathbf{I}_{d+1}$  is the  $(d+1) \times (d+1)$  identity matrix and  $\sigma_0^2$  is the prior variance. Explain the relationship between  $\lambda$ ,  $\sigma$  and  $\sigma_0$ ?

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