

2 95% of statistics (AJ Chapters 5-7)

1. (*) A p -value is a statistic $P : X \rightarrow [0, 1]$ such that under H_0 ,

$$\mathbb{P}(P \leq t) \leq t \text{ for } t \in [0, 1].$$

The p -value should be designed such that small values indicate a divergence between the data and the null hypothesis; the critical region for a test of size α is then $\{x : P(x) \in [0, \alpha]\}$.

- (a) Show that an independent Uniform(0,1) random variable is always a p -value. What is the power of the test?
- (b) Let X be a continuous random variable with strictly positive p.d.f. $f : \mathbb{R} \rightarrow (0, \infty)$. Show that the c.d.f.

$$F(x) = \mathbb{P}(X \leq x), \quad F : \mathbb{R} \rightarrow [0, 1]$$

is an invertible function and that the distribution of $F(X)$ is Uniform(0, 1).

- (c) Let $X_1, \dots, X_n \sim N(\theta, 1)$.
- Find a p -value for testing $H_0 : \theta = 0$ against $H_1 : \theta < 0$.
 - What is the critical region in terms of X_1, \dots, X_n ?
 - You are working at the 5% significance level. You want the probability of rejecting H_0 to be at least 90% if $\theta < -\epsilon$, where ϵ is a small, positive number. How large does n have to be as a function of ϵ ?
- (d) Let $X_1, \dots, X_n \sim N(\theta, 1)$. Find a p -value for testing $H_0 : \theta = 0$ against $H_1 : \theta \neq 0$.
- (e) Let $X_1, \dots, X_n \sim N(0, \sigma^2)$. Find a p -value for testing $H_0 : \sigma^2 = 1$ against $H_1 : \sigma^2 < 1$.
2. (*) Suppose $X_1, \dots, X_n \sim \text{Poisson}(\lambda)$ iidrv so $\mathbb{P}(X_i = k) = e^{-\lambda} \frac{\lambda^k}{k!}$.
- What, roughly speaking, is the distribution of $\bar{X} = \sum_{i=1}^n X_i/n$?
 - Given λ , find an interval such that \bar{X} belongs to that interval with probability 95%.
 - Re-arrange your answer to part (b) to produce an approximate 95% confidence interval for λ .
 - Let $\lambda = 2$ and $n = 5$. Using R, do the following a large number of times: Sample $X_1, \dots, X_n \sim \text{Poisson}(\lambda)$ and check if your confidence interval includes λ . What fraction of the time does your confidence interval include λ .
 - How about when $n = 100$?

3. Problems–Opportunities with Hypothesis Testing

- (a) The county of Midsomer has a population of 10,001. Or it did; there has been a murder. If we pick someone from the population uniformly at random, there is a one in 10,000 chance they are the murderer. A genetic test has been devised to test someone’s DNA with a DNA sample from the murder weapon. Relatives have similar DNA, so the machine will show a red light, if the test subject is the murderer, or any of the murderer’s 99 closest blood relatives. For the other 9,900 inhabitants of Midsomer, the machine will show a green light. Consider the hypothesis test being applied to a randomly chosen citizen of Midsomer. Consider H_0 : “not a murderer” against H_1 : “a murderer”.
- The machine show a red light. Should you (i) reject H_0 , or (ii) fail to reject H_0 / accept H_1 , at a 5% significance level? Interpret your test result.
 - What is the probability the test subject is the killer?
 - How, in the general context of hypothesis testing, can you reduce the number of false positives a test produces?
- (b) Suppose $X_1, \dots, X_{100} \sim N(\theta, 1)$. Consider a test $H_0 : \theta = 0$ against $H_1 : \theta = 100$. Calculate a critical region for a hypothesis test at a 5% significance level. What is the power of the test? You observe $\bar{X} = 10$. Do you reject H_0 or fail to do so? Compare the values of the likelihood function at $\theta = 0$ and $\theta = 100$. What do you think of the conclusion of the test?
- (c) Coins have two sides: heads or tails. They never fall on their side. (Well it did once in Midsomer, but we don’t talk about that any more).
- A coin might be biased. Consider $H_0 : \mathbb{P}(\text{heads}) = 0.6$ against $H_1 : \mathbb{P}(\text{heads}) = 0.4$. You toss the coin four times. Work at the 5% significance level.
 - Should you reject H_0 if you get four heads?
 - Should you reject H_0 if you get four tails?
 - A bag contains one biased coin ($\mathbb{P}(\text{heads}) = 0.6$) and 999 biased-the-other-way coins ($\mathbb{P}(\text{heads}) = 0.4$). A coin is picked from the bag uniformly at random. You toss it 4 times and get four heads. Using part (i), should you reject H_0 . Calculate the probability (conditional on the four tosses) that H_0 describes the coin.

4. The Hard EM algorithm is as its name suggest, much easier than the general EM-algorithm.

- Split the data $x = (x_o, x_m)$. You have only observed x_o . Start at some $\theta^{(0)}$.
- Iterate $\theta^{(t)} \rightarrow \theta^{(t+1)}$ by:
 - sampling $x_m = x_m(t)$ conditional on $\theta^{(t)}$, and then
 - setting $\theta^{(t+1)}$ to be the MLE for $x = (x_o, x_m)$.

A die has side-probabilities $p_1 + \dots + p_6 = 1$. Roll the die $2n$ times to get $Y_1, \dots, Y_{2n} \in \{1, \dots, 6\}$ independent. You observe

$$x_o = (Y_1 + Y_2, Y_3 + Y_4, \dots, Y_{2n-1} + Y_{2n})$$

but you do not observe the whole dataset, i.e.

$$x_m = (Y_1, \dots, Y_{2n}).$$

- (a) Download the file `hard_em.R` from the website. It contains x_o for $n = 1000$,

Total	2	3	4	5	6	7	8	9	10	11	12
#	1	11	24	38	82	133	138	166	169	147	91

and a function for estimating the M.L.E. for the (p_i) using the hard EM algorithm. Run the code to calculate the MLE approximately.

- (b) Try using some other method (perhaps moment generating functions, and the square root operation) to approximate the distribution of the die. Compare with your answer to part (a).
- (c) Reduce n to say 100 (`xo<-xo[1:100]`).

Total	2	3	4	5	6	7	8	9	10	11	12
#	0	1	2	4	11	9	12	17	22	15	7

Faced with a small amount of data, can you modify the algorithm in some simple way to improve performance.

5. Likelihood ratio test: Let $t \in (0, \infty)$ be some known quantity. Let $X_1, \dots, X_n \sim \text{Exponential}(r)$ with r unknown. Consider the simple test $H_0 : r = 1$ against $H_1 : r = t$. The simple likelihood ratio test has critical region

$$C_\alpha := \left\{ x : \frac{L(r = 1 | x)}{L(r = t | x)} \leq A_\alpha \right\} \quad L(\mu | x) = \prod_{i=1}^n r \exp(-rx_i).$$

- (a) Calculate the critical region as a function of t , and of the test size α , and of sample size n .
- (b) Calculate the power β of the simple likelihood ratio test.
- (c) Assuming β needs to be at least 0.9, what is the minimum sample size n as a function of t .

6. Let $X_1, \dots, X_n \sim \text{Bernoulli}(p)$ independent.
- (a) Is there a uniformly most powerful test for $H_0 : p = 1/2$ against $H_1 : p > 1/2$?
 - (b) Is there a uniformly most powerful test for $H_0 : p = 1/2$ against $H_1 : p \neq 1/2$?

7. Let $L(\theta | X)$ denote the likelihood. Assume regularity conditions are met...

Wilks' theorem states that when testing a composite null hypothesis $H_0 : \theta \in \Omega_0$ against a composite alternative hypothesis $H_1 : \theta \in \Omega_1 = \Omega \setminus \Omega_0$, the distribution of

$$-2 \log \frac{L(\hat{\theta}_0 | X)}{L(\hat{\theta} | X)} \sim \chi_{\dim \Omega - \dim \Omega_0}^2$$

where $\hat{\theta} = \arg \max_{\theta \in \Omega} L(\theta | X)$, $\hat{\theta}_0 = \arg \max_{\theta \in \Omega_0} L(\theta | X)$.

The probability distribution when rolling a die is described by 6 numbers, adding to one. You have two dice, and you roll each die n times. Denote the outcomes X_1, \dots, X_n for die 1, and Y_1, \dots, Y_n for die 2.

i	1	2	3	4	5	6
X	130	116	109	197	223	225
Y	100	108	128	209	210	245

Test H_0 : the dice have the same distribution against H_1 : the distributions differ.

8. Read Chapters 6 and 7 of AJ's notes.