

## Large deviation theory

Hendrik Weber  
hendrik.weber  
@ wwwwide.ac.  
uk

## Literature:

- Frank den Hollander Large deviations
  - Denuchel & Broek "n"
  - Dembo & Zeitouni LD, Techniques and applications

Lecture notes Wolfgang König (online)  
Feng & Kurtz "Large deviations for  
stochastic processes"

## Introduction

$$\text{LLN: } (x_i) \text{ i.i.d real valued random var} \\ S_n = \sum_{i=1}^n x_i \quad \frac{1}{n} S_n \xrightarrow{\text{a.s.}} \mathbb{E}[x_i] \\ \text{if } \mathbb{E}[|x_i|] < \infty$$

CLT: typical fluctuations are of order  $\sqrt{n}$ .  
 Question: How does  $P\left[\frac{1}{n} S_n \geq x\right]$  behave  
 for  $x \geq \mathbb{E}[Y]$ ?

Example: (Markov chain Monte Carlo)

$\mu$  probability measure on  $E$   
 Look for  $K(x, dy)$  on  $E$  such that  
 $\mu$  is unique invariant measure for  $M_{\alpha, \beta}^{\text{Markov}}(x)$

$$\Rightarrow \text{Ergodic}_{\mathcal{C}}^{\text{Thm.}}: \mu_N(\beta) = \frac{1}{N} \sum_{i=1}^N \mathbb{1}_{\{\beta\}}(x_i) \xrightarrow{\text{a.s.}} \mu(\beta)$$

$P\{\mu_n \approx v\}$ , how does this decay?

Example- Suppose we have

$$d\dot{x}_t^\varepsilon = -V(x_t^\varepsilon)dt + \varepsilon dw_t \quad x \in \mathbb{R}^n$$

To fix  $\varepsilon \rightarrow$   
 $x_t^\varepsilon \xrightarrow{\text{a.s.}} x_t$

Outline: First Classical "Crammer  
Schilder  
Samov  
Gardner's

Second, Markov processes  
Fréchet-Wentzell  
Donsker-Varadhan

Chapter 1 Large deviations for iid sequences  
 'Gauß Thm' '38

I A Coin tossing

$$X_i = \begin{cases} 1 & \text{with proba } \frac{1}{2} \\ 0 & \text{independent} \end{cases}$$

$$\mathbb{E}[X_i] = \frac{1}{2} \quad \text{fix } x > \frac{1}{2}$$

$$S_n = \sum_{i=1}^n X_i$$

Question: How does  $\mathbb{P}\left[\frac{1}{n} S_n \geq x\right]$  behave for large  $n$ .

Answer:

$$\mathbb{P}[S_n \geq nx]$$

$$= \sum_{k \geq nx} \binom{n}{k} \cdot 2^{-n}$$

$$2^{-n} \max_{k \geq nx} \binom{n}{k} \leq \mathbb{P}[S_n \geq nx] \leq (n+1) 2^{-n} \max_{k \geq nx} \binom{n}{k}$$

Max is attained at  $k = \lceil nx \rceil$

$$\max_{k \geq nx} \binom{n}{k} = \frac{n!}{(n-nx)! (nx)!} = *$$

$$n! \sim n^n 2^n \sqrt{2\pi n}$$

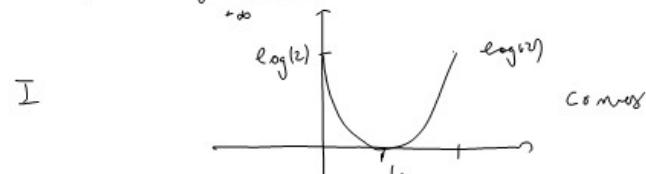
$$\Rightarrow * \sim \frac{\sqrt{2\pi n}}{\sqrt{2\pi(n-nx)} \sqrt{2\pi nx}} \cdot \underbrace{\frac{2^{-n}}{2^{-(n-nx)} \cdot 2^{nx}}}_{= 1}$$

$$\cdot \underbrace{\frac{n^n}{(n(1-x))^{n(1-x)} nx^{nx}}}_{= \frac{1}{(1-x)^{n(1-x)}} x^{nx}}$$

$$\Rightarrow \log(\mathbb{P}[S_n \geq nx]) = -n \log(2) - n(1-x) \log(1-x) - n(x) \log x + \text{lower order}$$

$$\Rightarrow \frac{1}{n} \log \mathbb{P}[S_n \geq nx] \xrightarrow{n \rightarrow \infty} -I(x)$$

$$I(x) = \lim_{n \rightarrow \infty} \left[ -n \log(2) - n(1-x) \log(1-x) - n(x) \log x \right]$$



Conclusion: ①  $\mathbb{P}[X_n \geq x]$  decay exponentially fast. Question is right exponential rate.

Procedure of taking

$$\frac{1}{n} \log \mathbb{P}(A) \rightarrow ?$$

Corresponds to throwing away lower order.

unique zero (at  $\frac{1}{2}$ )  
 $\Rightarrow \text{LLN}$   
 For every  $\delta > 0$   $P\left(\frac{1}{n} S_n \geq \frac{1}{2} + \delta\right)$   
 summable. By Borel-Cantelli ok.

③  $\nu, \mu$  are measures on some space  $E$

$$H(\nu || \mu) \quad \text{relative entropy}$$

$$= \begin{cases} \int \log \frac{d\nu}{d\mu} d\nu & \text{if } \nu \text{ ac. } \mu \\ +\infty & \text{else.} \end{cases} \quad \text{Kullback-Leibler divergence}$$

$$\mu_x \text{ on } \{0,1\} \text{ such that under this measure}$$

$$\int y \mu_x(dy) = x$$

$$I(x) = H(\mu_x || \frac{1}{2}\delta_0 + \frac{1}{2}\delta_1) \quad \underline{\text{Exo}}$$

[B] From now on  $X_i$ : i.i.d real valued random variables

$$\Psi(t) = \mathbb{E}[e^{tX}] < \infty \quad \forall t$$

$$\mathbb{E}(X_i) = m \quad \text{Var}(X_i) = \sigma^2 \quad S_n = \sum_{i=1}^n X_i$$

Question: How does the probability

$$P\left(\frac{1}{n} S_n \geq x\right) \text{ decay to 0 for } x > m^2$$

Calculation:  $t > 0$

$$P\left(\frac{1}{n} S_n \geq x\right) = P\left[\exp(tS_n) \geq \exp(tx)\right]$$

$$\stackrel{\text{Chebycheff}}{\leq} \exp(-tx) \mathbb{E}[\exp(tS_n)]$$

$$= \exp(-tx) \Psi(t)^n$$

$$= \exp(-n(tx - \log \Psi(t))).$$

$$\Rightarrow P\left(\frac{1}{n} S_n \geq x\right) \leq \exp(-nI(x))$$

$$I(x) = \sup_{t>0} (tx - \log \Psi(t)).$$

Comments: Legendre transform of  $\Psi$

$$\Psi^*(x) = \sup_{t \in \mathbb{R}} (tx - \Psi(t))$$

Rate function is the Legendre transform of log of moment generating function.

Exo. The Rate function  $I$  from [A] is also of that form.

Assume that  $t^*$  minimises

$$\Rightarrow x - \frac{\Psi'(t^*)}{\Psi(t^*)} = 0$$

$$0 = x - \frac{\varphi'(t)}{\varphi(t)} \quad \varphi(t) = \mathbb{E}(e^{tx})$$

$$\Rightarrow x = \frac{\mathbb{E}[X e^{tx}]}{\mathbb{E}(e^{tx})} \quad \textcircled{*}$$

$\mu$  = distribution of  $X$ .  
 $\nu = \frac{e}{\varphi(t)}$  defines a new proba. measure  
then (x) states that  $\mathbb{E}_\nu(X) = x$ .

### [C] Cramér Theorem:

Theorem:  $X_i$  i.i.d.,  $\varphi(t) = \mathbb{E}(e^{tx}) < \infty$  for all  $t$   
 $S_n = \sum_{i=1}^n X_i$ ,  $m = \mathbb{E}(X_i)$ ,  $\sigma^2 = \text{Var}(X_i)$

Then for all  $x \geq m$

$$\lim_{n \rightarrow \infty} n \log \mathbb{P}\left[\frac{1}{n} S_n \geq x\right] = -I(x)$$

$$I(x) = \sup_{t \in \mathbb{R}} (tx - \log \varphi(t))$$

proof: We assume  $X$  is not deterministic.

Step 1: We can reduce ourselves to case

$$x = 0 \quad m < 0.$$

proof of ①: Assume, we have proved that case.

Then for general  $X$  consider  $Y = X - x$

$$\mathbb{P}\left[\sum_{i=1}^n X_i \geq x\right] = \mathbb{P}\left[\frac{1}{n} \sum Y_i \geq 0\right]$$

taking log, multiply with  $n \longrightarrow -I_Y(0)$

$$I_Y(0) = \sup_{t \in \mathbb{R}} (-\log \varphi_Y(t)) \Rightarrow \varphi_Y(r) = \mathbb{E}(e^{tr})$$

$$= \mathbb{E}(e^{tx}) e^{t^2}$$

$$= \sup_{t \in \mathbb{R}} (tx - \log \varphi(t)) = I(x)$$

Hence, we want to proof

$$n \log \mathbb{P}[S_n \geq 0] \longrightarrow -I(0) = -\sup_{t \in \mathbb{R}} (-\log \varphi(t))$$

$$= \inf_{t \in \mathbb{R}} \varphi(t)$$

$$= \log \beta.$$

Step 2 Observations:  $\varphi \in C^\infty$

$$\varphi(t) = \mathbb{E}(e^{tx})$$

$$\varphi'(t) = \mathbb{E}(X e^{tx})$$

$$\varphi''(t) = \mathbb{E}(X^2 e^{tx}) > 0$$

$\varphi$  strictly convex, inf can be attained at at most 1 point.

Trivial cases:

$$(i) \quad P[X_i \geq 0] = 0 \implies P[S_n \geq 0] = 0$$

in this case

$$\varphi(t) = E[e^{tX}] > 0$$

and  $\lim_{t \rightarrow \infty} \varphi(t) = E[e^{tX}] = 0$

$$(ii) \text{ Assume that } P[X_i \leq 0] = 1 \text{ but}$$

$$P[X_i = 0] > 0$$

$$\begin{aligned} \text{Then } \varphi(t) &= E[e^{tX}] \\ &= \underbrace{E[e^{tX} \mathbb{1}_{\{X < 0\}}]}_{\rightarrow 0 \text{ as } t \rightarrow \infty} + \underbrace{E[e^{tX} \mathbb{1}_{\{X \geq 0\}}]}_{P[X \geq 0]} \\ \Rightarrow \varphi &= P[X \geq 0] \\ P[S_n \geq 0] &= \varphi^n. \end{aligned}$$

$$\underline{\text{Step 3}} \quad P[X > 0] > 0, \quad P[X < 0] > 0.$$

$$\varphi(t) \rightarrow \pm \infty \quad t \rightarrow \pm \infty$$

$\varphi$  strictly convex

$\Rightarrow$  there is a unique minimum attained at  $t_*$ .

$$t_* > 0 \quad \text{because} \quad \varphi'(0) = E[X] = m < 0.$$

Upper bound: Already proved that  $t > 0$

$$P[S_n \geq 0] = P[\exp(t_n S_n) \geq 1]$$

$$\leq E[\exp(t_n S_n)] = \varphi^n(t)$$

minimum is attained for a positive  $t$ .

Lower bound: Recall  $\mu$  was the mean the distribution of  $X$  and

$$\nu \text{ as } e^{tX} \cdot \varphi^{-1} \mu(dx)$$

under  $\nu \quad E[X] = 0$ .

$$\begin{aligned} P_\mu[S_n \geq 0] &= \mathbb{E}_\mu \left[ e^{-t_n S_n} \frac{\prod_{i=1}^n e^{t_i X_i}}{\varphi^n} \mathbb{1}_{\{S_n \geq 0\}} \right] \\ &= \varphi^n \mathbb{E}_\nu \left[ e^{-t_n S_n} \mathbb{1}_{\{S_n \geq 0\}} \right] \end{aligned}$$

$$\text{It only remains } \mathbb{E}_\nu \left[ e^{-t_n S_n} \mathbb{1}_{\{S_n \geq 0\}} \right]$$

$$\text{under } \nu \quad \frac{S_n}{\sqrt{n}} \Rightarrow \mathcal{N}(0, \sigma^2)$$

$$\mathbb{E}_\nu \left[ e^{-t_n S_n} \mathbb{1}_{\{S_n \geq 0\}} \right] \geq$$

$$\begin{aligned} &\mathbb{E} \left[ e^{-t_n \sqrt{n} Z} \mathbb{1}_{\{Z \geq 0\}} \right] \\ &\geq e^{-t_n \sqrt{n} C} \quad \square \end{aligned}$$

### D) Discussion of rate function

$$I(x) = \sup_t (tx - \log \varphi(t))$$

Lemma: (i)  $\log \varphi$  convex;  $I(x)$  convex  
(ii)  $I \geq 0$   $I(x) = 0$  iff  $x = m = E(X)$   
(iii) Sublevels sets of  $I$  are compact, in particular  
 $I$  is l.s.c.

Proof of (i)

$$\begin{aligned} & \log \varphi(tx_1 + (1-\lambda)x_2) \\ &= \log \mathbb{E} \left[ e^{tx_1 + (1-\lambda)x_2} X \right] \\ &\stackrel{\text{Hölder}}{\leq} \log \mathbb{E} \left[ e^{x_1 X} \right]^\lambda \mathbb{E} \left[ e^{x_2 X} \right]^{(1-\lambda)} \\ &= \lambda \log \varphi(x_1) + (1-\lambda) \varphi(x_2) \\ I(\lambda x_1 + (1-\lambda)x_2) &= \sup_t (t(\lambda x_1 + (1-\lambda)x_2) - \log \varphi(t)) \\ &\leq \lambda \sup_t (tx_1 - \log \varphi(t)) + (1-\lambda) \sup_t (tx_2 - \log \varphi(t)) \end{aligned}$$

ref exercise.

Remark: This implies LLN.

- $\varphi(t) < \infty$  but  $\Rightarrow$  too strong. We will discuss general case later.

### E) Relative entropy

$\exists$  some measure space,  $\nu, \mu$  proba-measures

$$H(\nu || \mu) = \begin{cases} \int \log \frac{d\nu}{d\mu} d\nu & \text{if } \nu \ll \mu \\ +\infty & \text{else} \end{cases}$$

relative entropy - Kullback-Leibler divergence.

Lemma:  $H(\nu || \mu) \geq 0$ , equality iff  $\nu = \mu$ .

Proof:  $f = \frac{d\nu}{d\mu} \rightarrow H(\nu || \mu) = \int (\log f) f d\mu$

we know that  $\int f d\mu = 1$

$$(\log f) f \geq f - 1 \quad \text{equality if } f = 1$$

Lemma:  $X_i$ : iid random variables,  $\varphi(t) = \text{moment gen. fn.}$

$\mu = \text{distribution of } X$   $I(x) = \sup_t (tx - \log \varphi(t))$

$\nu_x = \text{"shifted measure"}$

(i)  $H(\nu_x || \mu)$  is minimized by  $\nu_x$  among all

proba. measures that have expectation  $x$ .

(ii)  $H(\nu_x || \mu) = I(x)$

prob] Ex 0.