

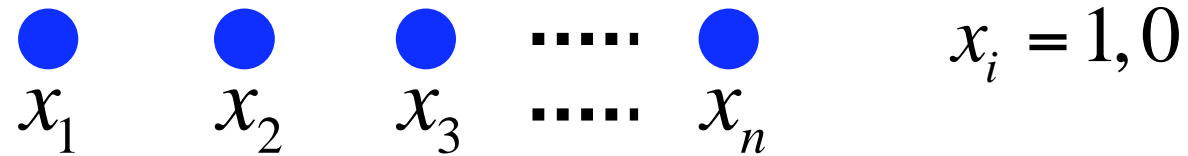
**Computational Neuroscience ESPRC Workshop
--Warwick**

**Information-Geometric
Studies on Neuronal Spike
Trains**

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Mathematical Neuroscience Unit**

Neural Firing



$p(\mathbf{x}) = p(x_1, x_2, \dots, x_n)$: joint probability

$r_i = E[x_i]$ ---- firing rate $S = \{p(x_1, x_2, \dots, x_n)\}$

$v_{ij} = Cov[x_i, x_j]$ ---- covariance: correlation

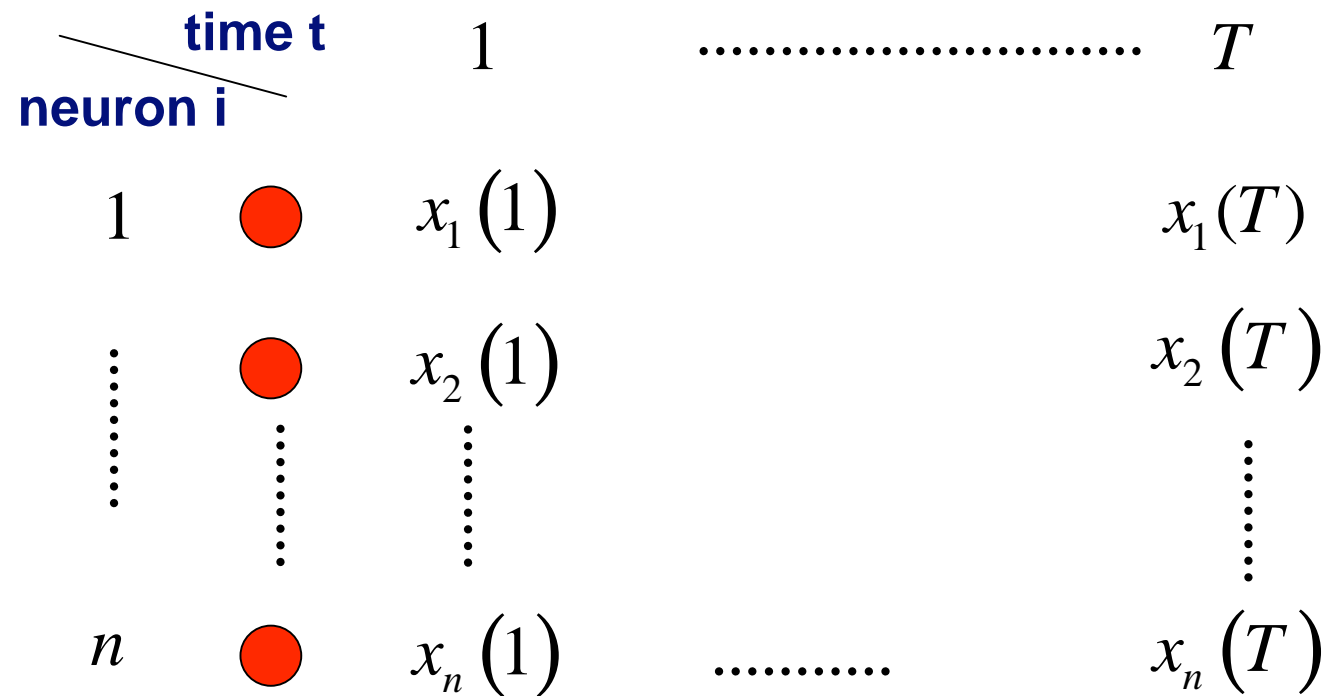
higher-order correlations

orthogonal decomposition

Multiple spike sequence:

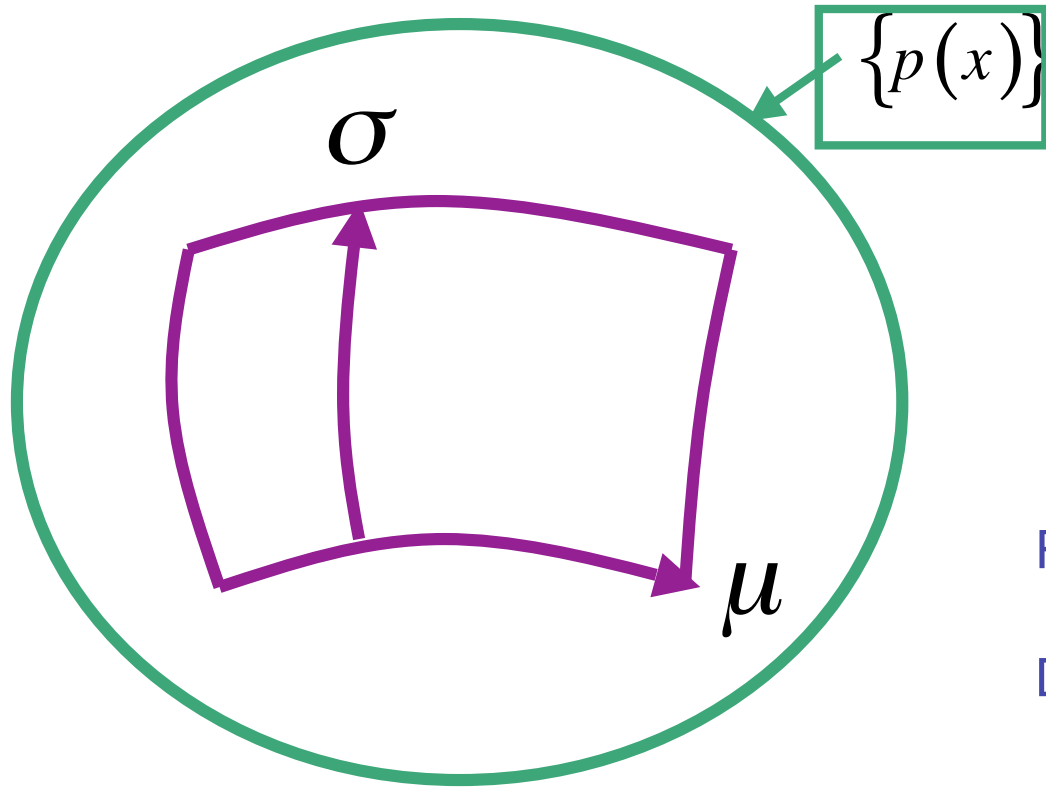
$$\{x_i(t), i = 1, \dots, n; t = 1, \dots, T\}$$

$$x_i(t) = 0, 1$$



Information Geometry ?

$$\mathcal{S} = \{p(x; \mu, \sigma)\} \quad p(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left\{-\frac{(x - \mu)^2}{2\sigma^2}\right\}$$



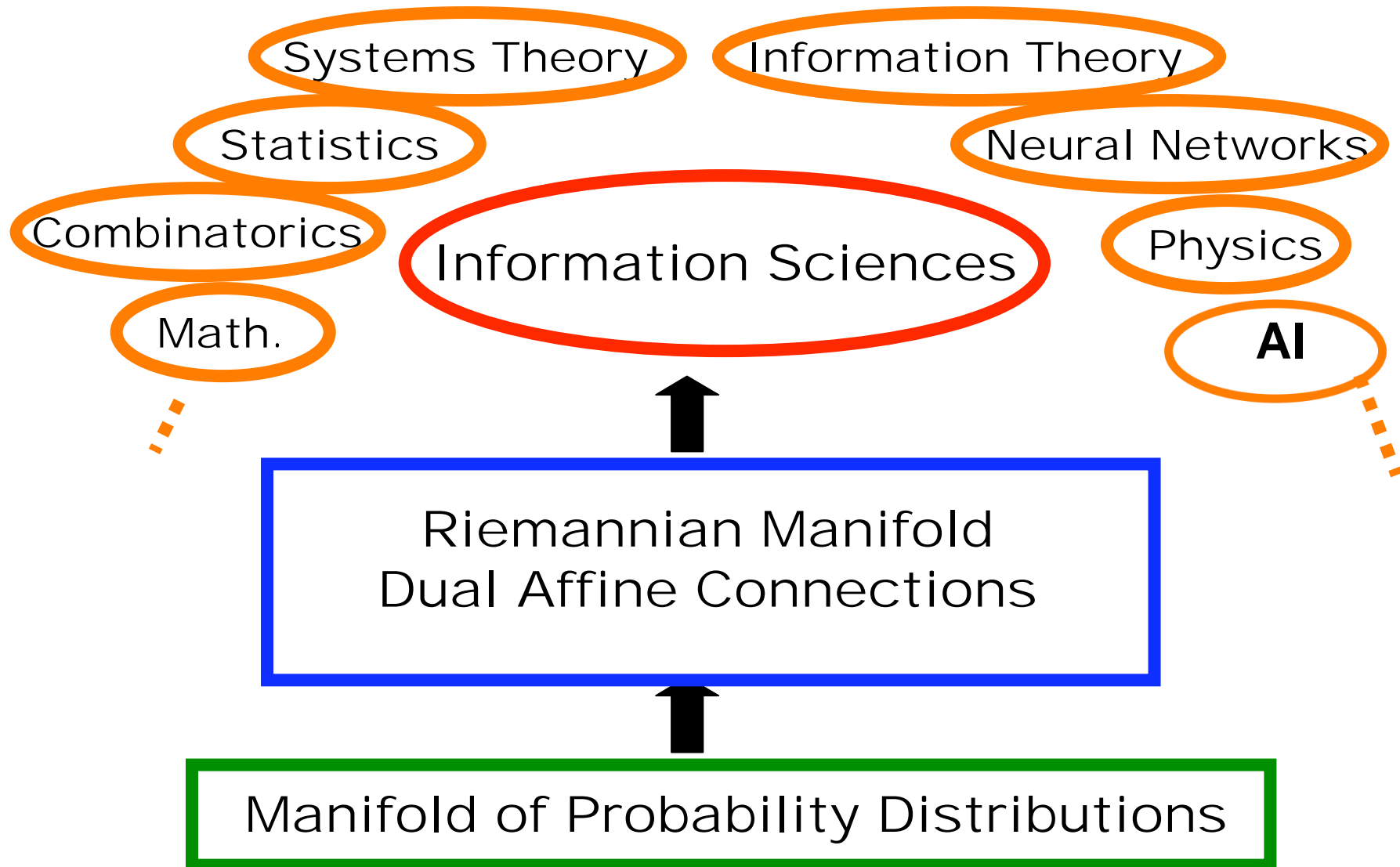
$$\mathcal{S} = \{p(x; \mathbf{\hat{e}})\}$$

$$\mathbf{\hat{e}} = (\mu, \sigma)$$

Riemannian metric

Dual affine connections

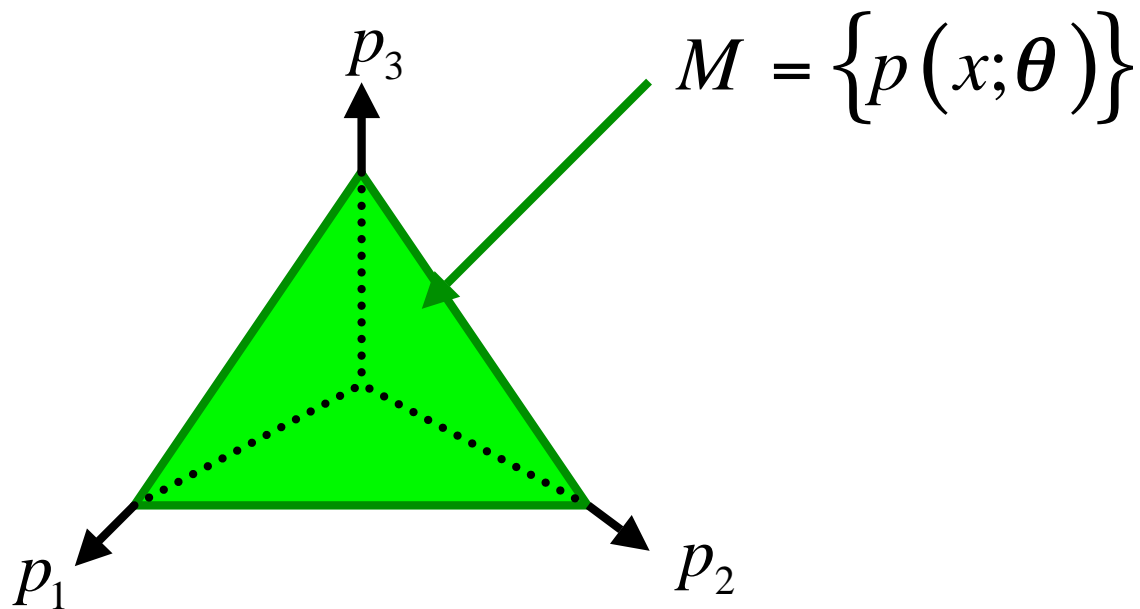
Information Geometry



Manifold of Probability Distributions

$$x = 1, 2, 3 \quad \{p(x)\}$$

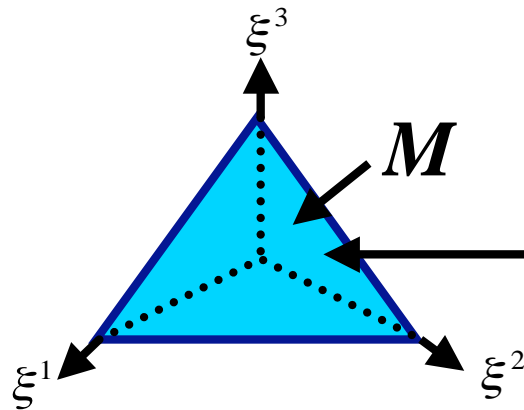
$$\mathbf{p} = (p_1, p_2, p_3) \quad p_1 + p_2 + p_3 = 1$$



Manifold of Probability Distributions

$$\mathcal{M} = \{P(x, \mathbf{p})\}$$

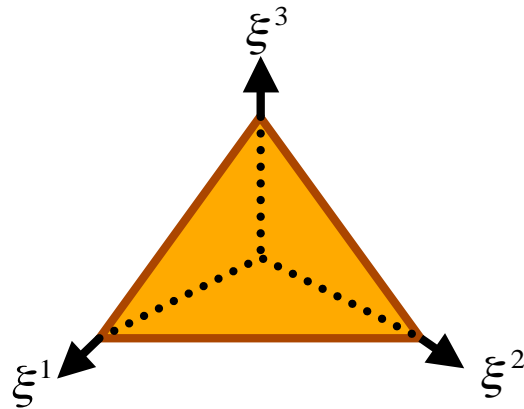
$$x = 1, 2, 3$$



$$P(x, \mathbf{p}) = \sum_{i=1}^3 p_i \delta_i(x) \quad (p_3 = 1 - p_1 - p_2)$$

$$\mathbf{p} = (p_1, p_2, p_3)$$

$$p_1 + p_2 + p_3 = 1$$



$$\xi_i = \sqrt{p_i}$$

$$\xi_1^2 + \xi_2^2 + \xi_3^2 = 1$$

$$\begin{cases} \theta_1 = \log \frac{p_1}{p_3} \\ \theta_2 = \log \frac{p_2}{p_3} \end{cases}$$

$$\boldsymbol{\theta} = (\theta^1, \theta^2)$$

Invariance

$$S = \{p(x, \theta)\}$$

1. Invariant under reparameterization

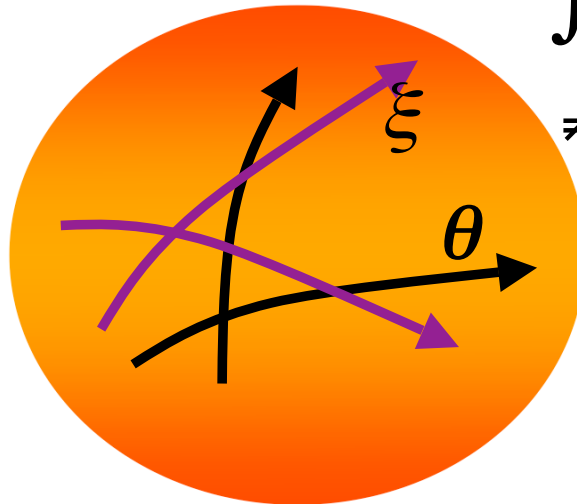
$$p(x, \theta) = \bar{p}(x, \xi) \quad D = \sum \theta_i^2 \neq \sum \xi_i^2$$

2. Invariant under different representation

$$y = y(x), \quad \{\bar{p}(y, \theta)\} = \{p(x, \theta)\}$$

$$\int |p(x, \theta_1) - p(x, \theta_2)|^2 dx$$

$$\neq \int |\bar{p}(y, \theta_1) - \bar{p}(y, \theta_2)|^2 dy$$

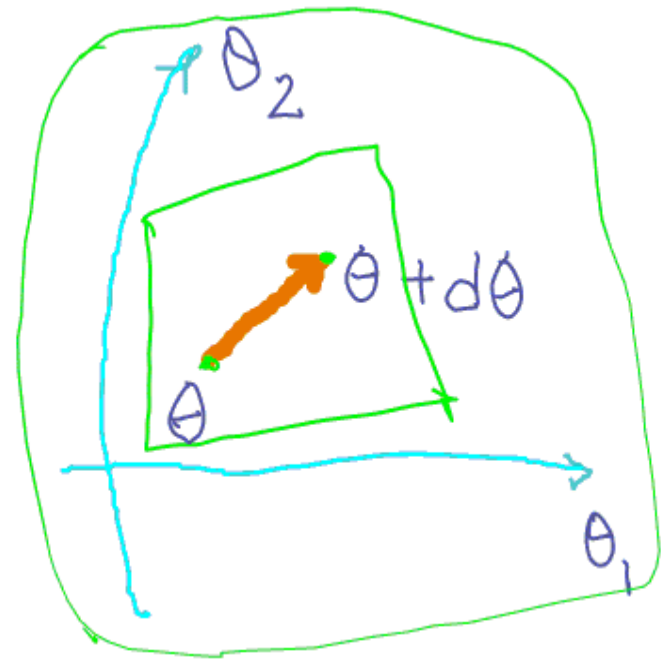


Two Structures

Riemannian metric
affine connection --- geodesic

$$g_{ij} = E \left[\frac{\partial}{\partial \theta_i} \log p \frac{\partial}{\partial \theta_j} \log p \right]$$

Fisher information



$$ds^2 = \sum g_{ij}(\theta) d\theta_i d\theta_j = \langle d\theta, d\theta \rangle$$

Affine Connection

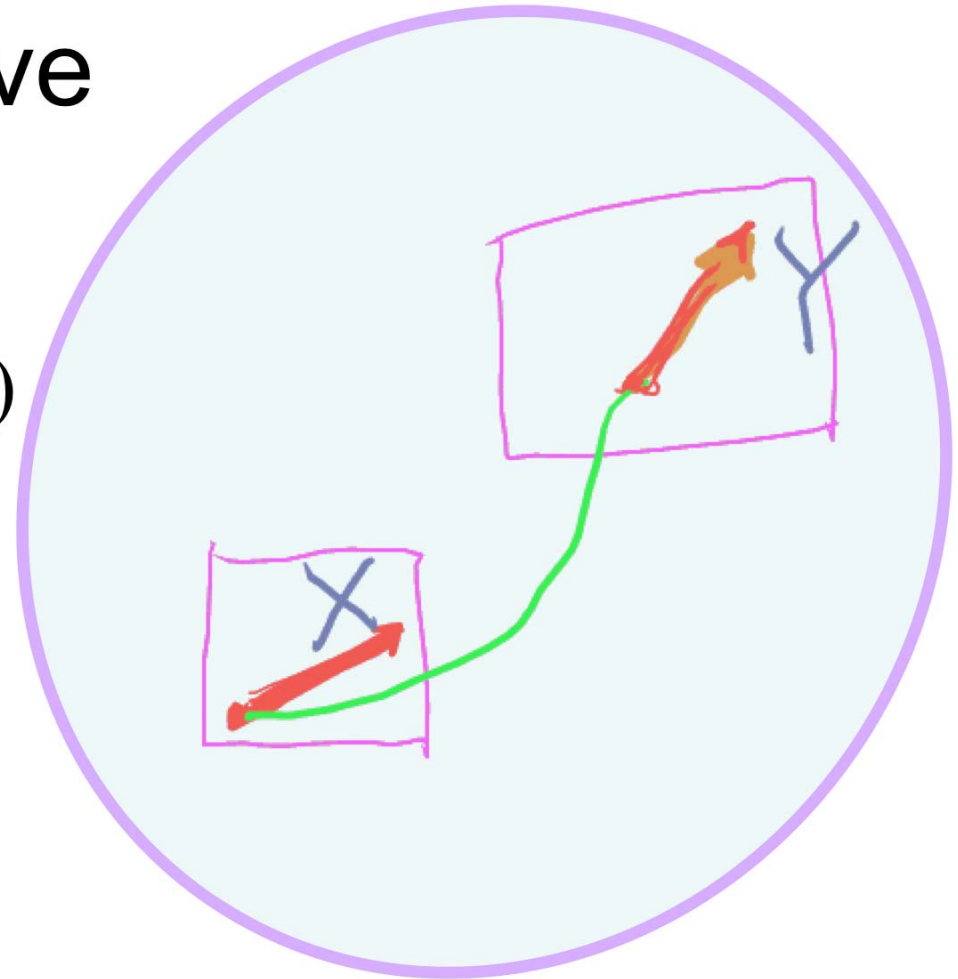
covariant derivative

$$\Pi_c X = Y$$

geodesic $\Pi \dot{X} = \dot{X}$ $X = X(t)$

$$s = \int \sqrt{\sum g_{ij}(\theta) d\theta^i d\theta^j}$$

minimal distance
straight line



Affine Connection

covariant derivative; parallel transport

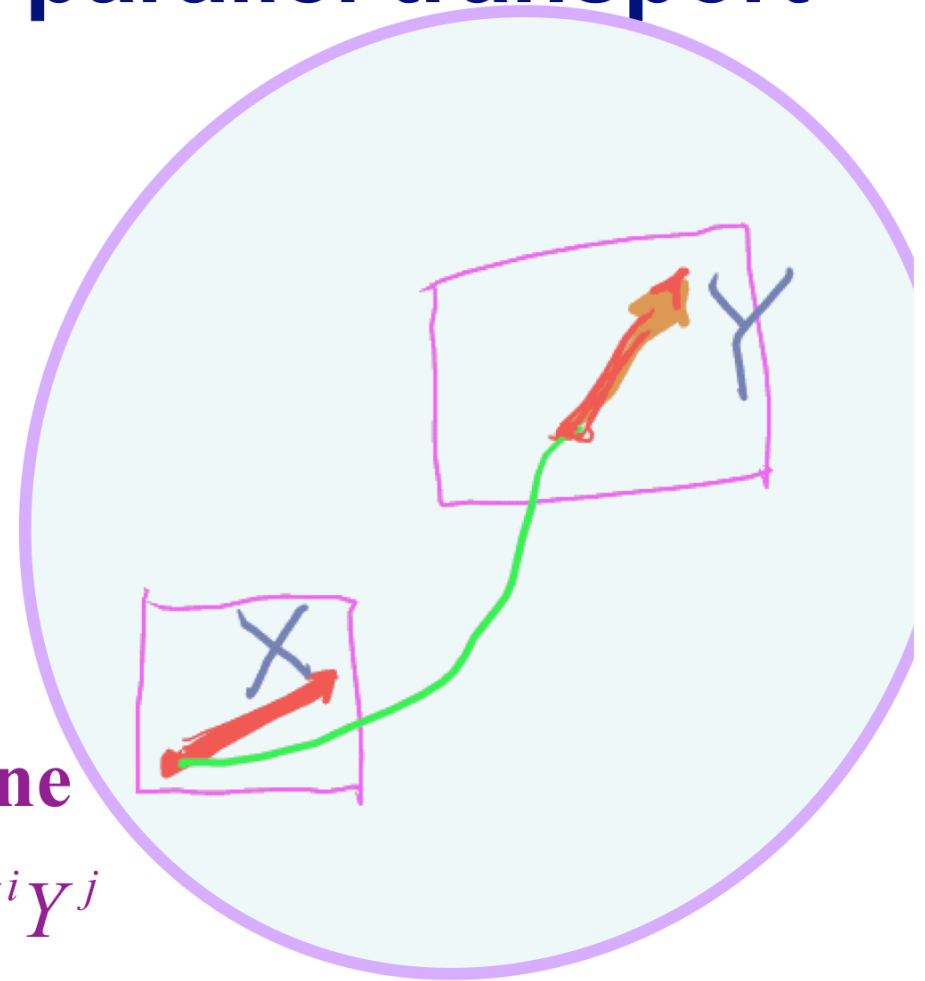
$$\nabla_X Y, \quad \Pi_c X = Y$$

geodesic $\Pi \dot{x} = \dot{x} \quad x = x(t)$

$$s = \int \sqrt{\sum g_{ij}(\theta) d\theta^i d\theta^j}$$

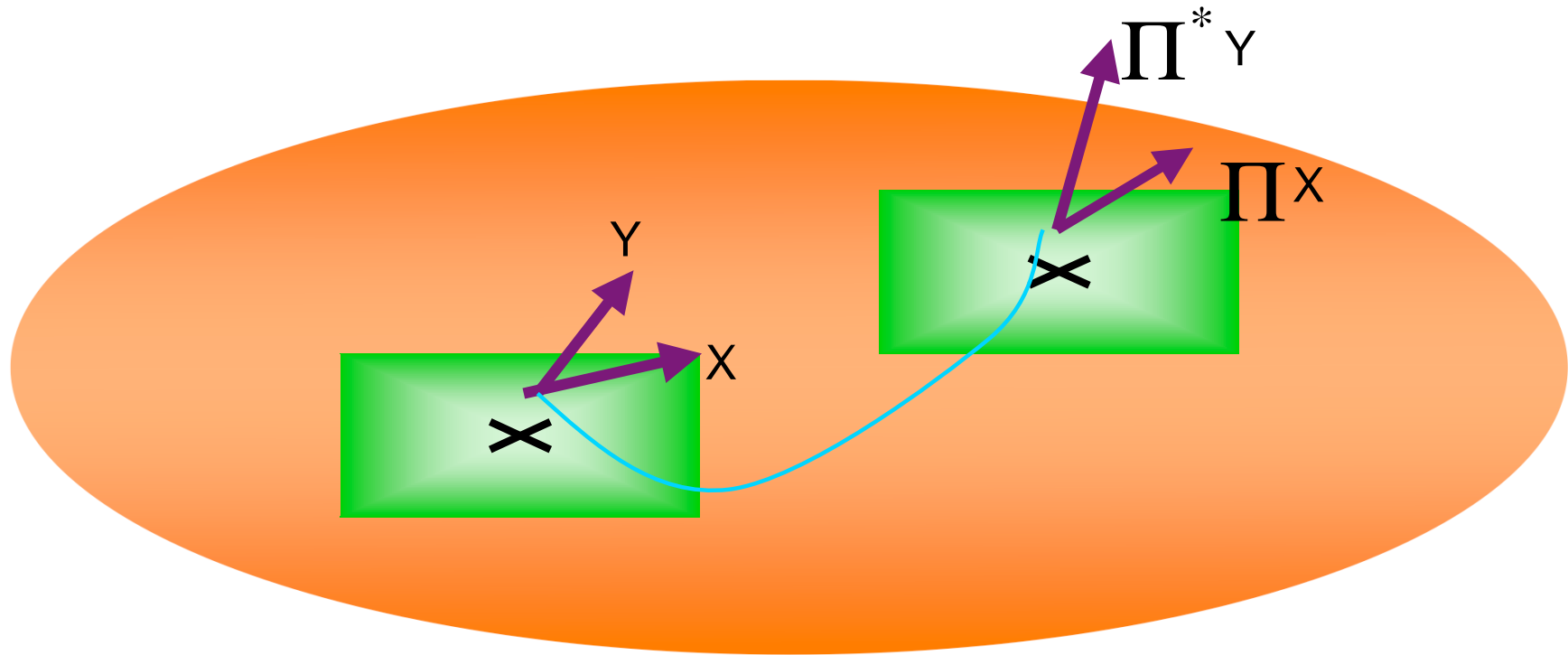
minimal distance: straight line

$$\langle \Pi X, \Pi Y \rangle = \langle X, Y \rangle \quad g_{ij} X^i Y^j$$



Duality: two affine connections $\{S, g, \nabla, \nabla^*\}$

$$\langle X, Y \rangle = \langle \Pi X, \Pi^* Y \rangle \quad \langle X, Y \rangle = \sum g_{ij} X^i Y^j$$



Riemannian geometry:

$$\Pi = \Pi^*$$

Dual Affine Connections

$$(\nabla, \nabla^*)$$

$$(\Pi, \Pi^*)$$

e-geodesic

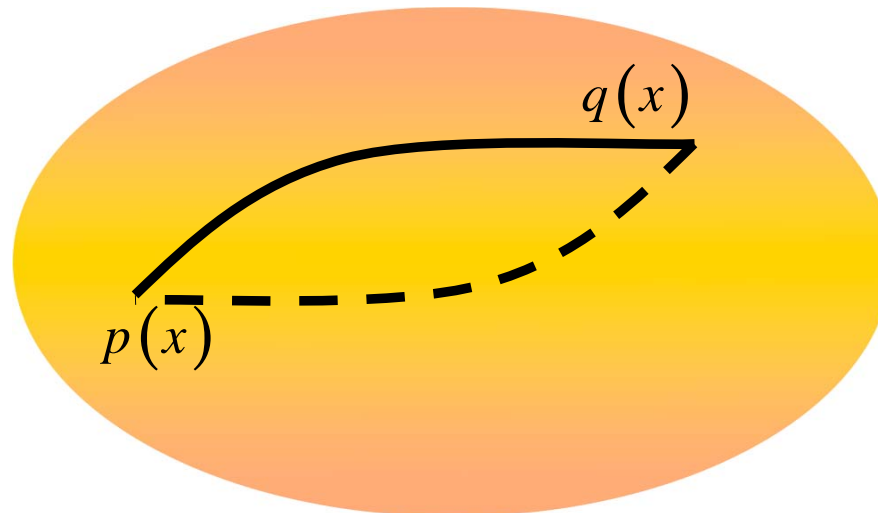
$$\log r(x, t) = t \log p(x) + (1-t) \log q(x) + c(t)$$

m-geodesic

$$r(x, t) = tp(x) + (1-t)q(x)$$

$$\nabla_{\dot{x}} \dot{x}(t) = 0$$

$$\nabla^*_{\dot{x}} \dot{x}(t) = 0$$



Information Geometry

-- Dually Flat Manifold

Convex Analysis

Legendre transformation

Divergence

Pythagorean theorem

I-projection

Dually Flat Manifold

1. Potential Functions

---convex (Bregman, Legendre transformation)

2. Divergence

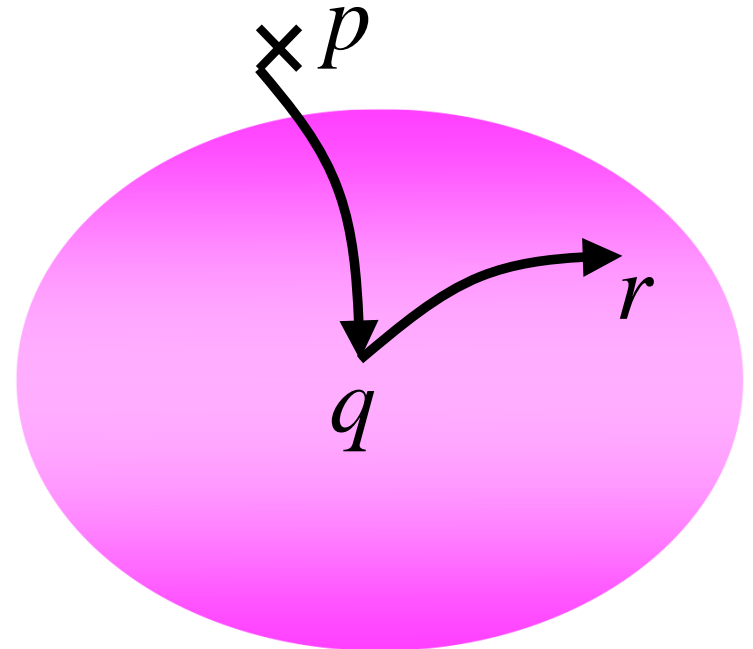
$$D[p : q]$$

3. Pythagoras Theorem

$$D[p : q] + D[q : r] = D[p : r]$$

4. Projection Theorem

5. Dual foliation



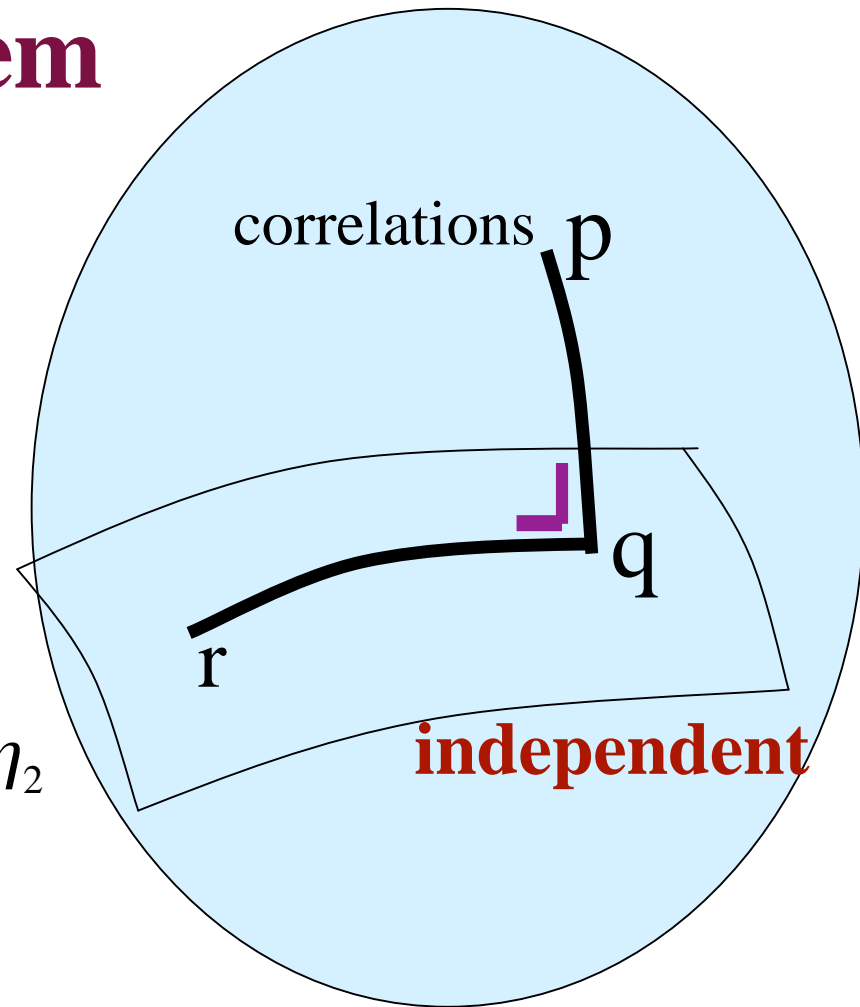
Pythagoras Theorem

$$D[p:r] = D[p:q] + D[q:r]$$

p, q : same marginals η_1, η_2

r, q : same correlations θ

$$D[p:r] = \sum_x p(x) \log \frac{p(x)}{q(x)}$$



estimation correlation
testing

invariant under firing rates

Projection Theorem

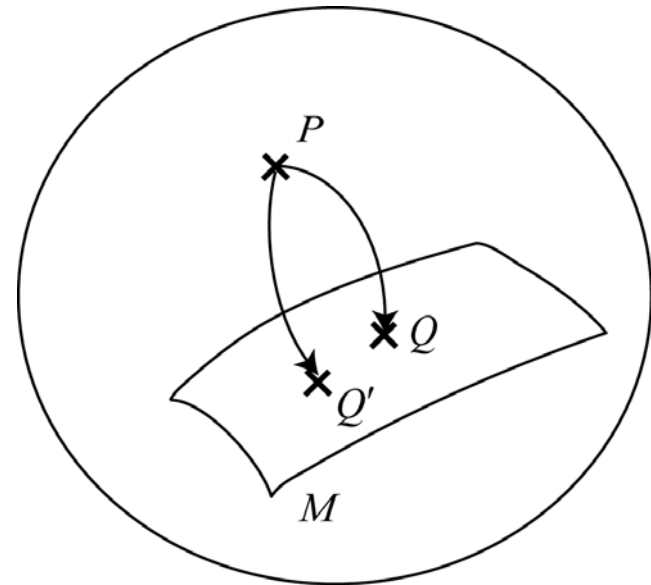
$$\min_{Q \in M} D[P : Q]$$

$Q =$ m-geodesic

projection of P to M

$$\min_{Q \in M} D[Q : P]$$

$Q =$ e-geodesic projection of P to M



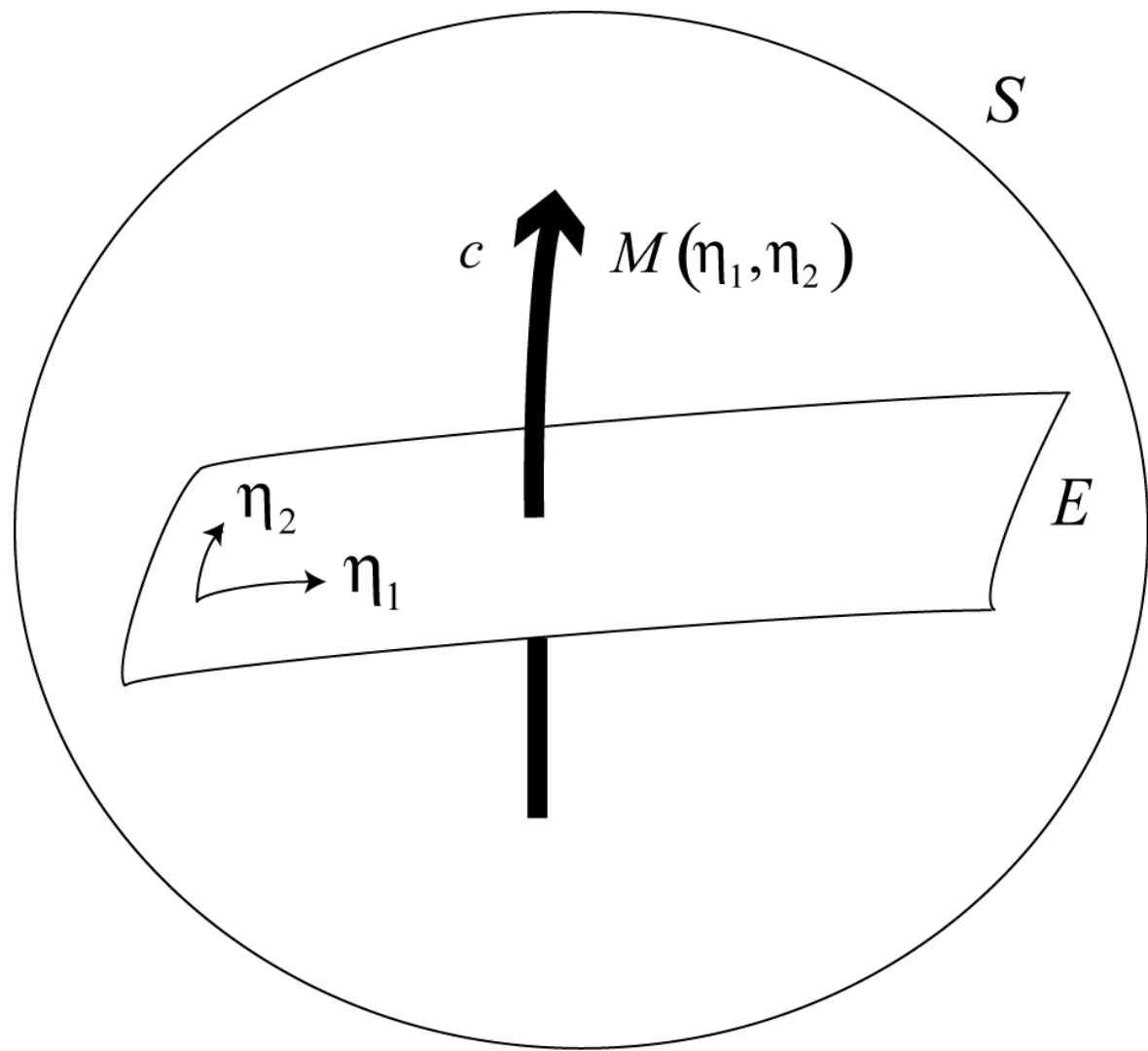


Fig. 1a

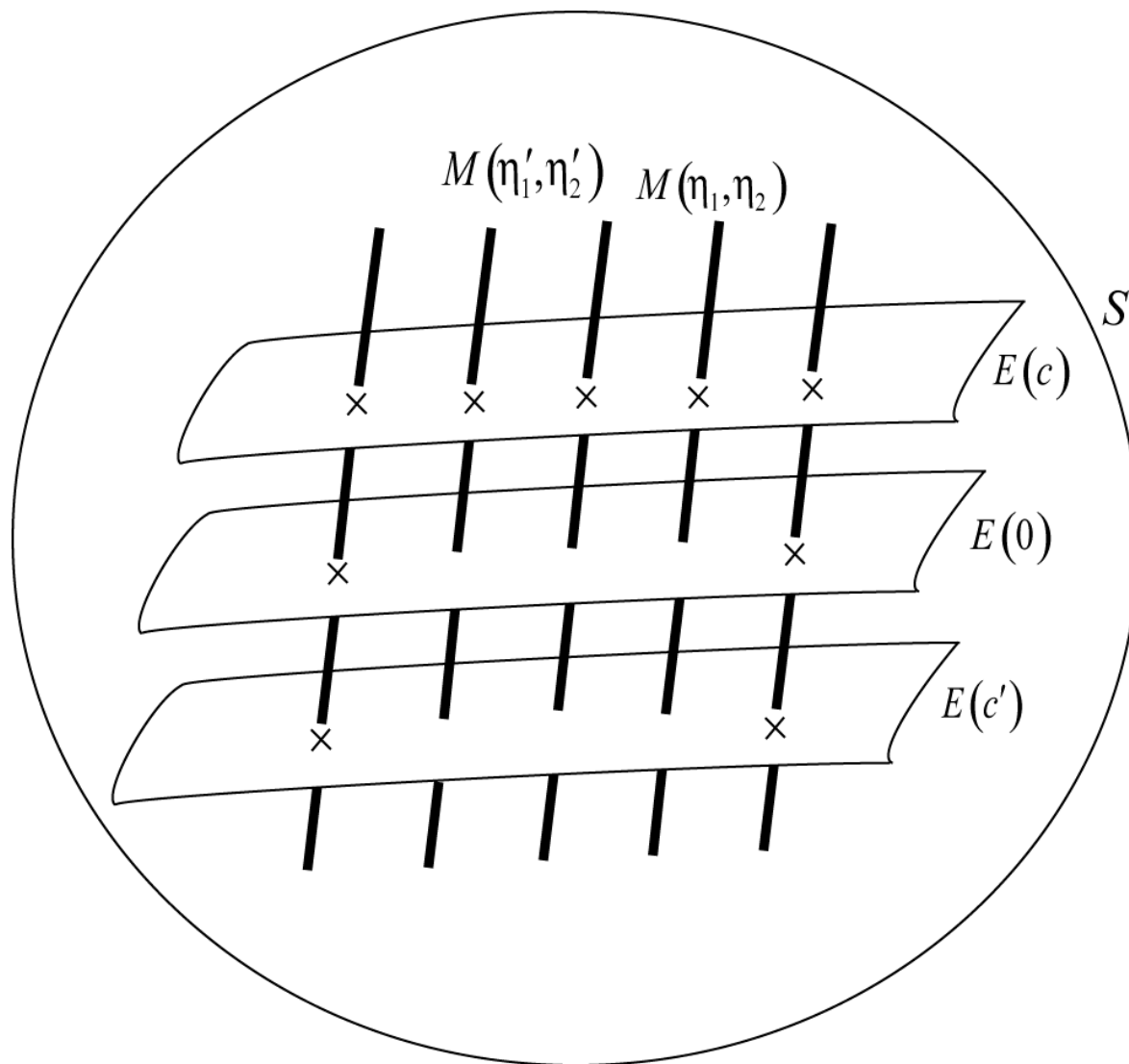
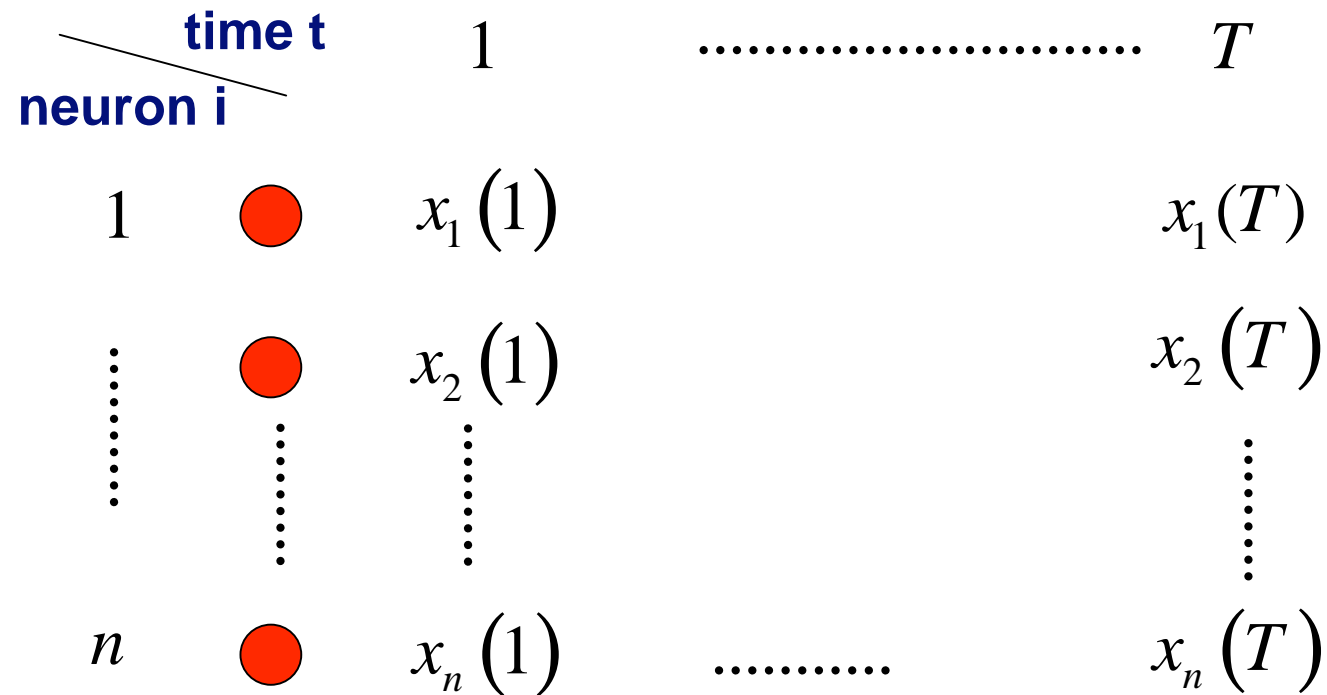


Fig. 1b

Multiple spike sequence:

$$\{x_i(t), i = 1, \dots, n; t = 1, \dots, T\}$$

$$x_i(t) = 0, 1$$



spatial correlations

$$\mathbf{x} = (x_1, \dots, x_n)$$

$$p(\mathbf{x}) = \exp \left\{ \sum \theta_i x_i + \sum_{i < j} \theta_{ij} x_i x_j + \dots + \theta_{1 \dots n} x_1 \dots x_n - \psi(\boldsymbol{\theta}) \right\}$$

$$r_i = E[x_i] = \text{Prob} \{x_i = 1\}$$

$$r_{ij} = E[x_i x_j] = \text{Prob} \{x_i = x_j = 1\}$$

⋮

$$\boldsymbol{\theta} = (\theta_i, \theta_{ij}, \dots, \theta_{1 \dots n})$$

$$\mathbf{r} = (r_i, r_{ij}, \dots, r_{1 \dots n})$$

$S = \{p(\mathbf{x})\}$: coordinates

orthogonal structure

Spatio-temporal correlations

correlated Poisson

$p(x)$: temporally independent

correlated renewal process

$p(x_t)$: firing rate $r_i(t)$ modified

spatial correlations fixed

Two neurons: $\{p_{00}, p_{01}, p_{10}, p_{11}\}$

x_1 0011000101101

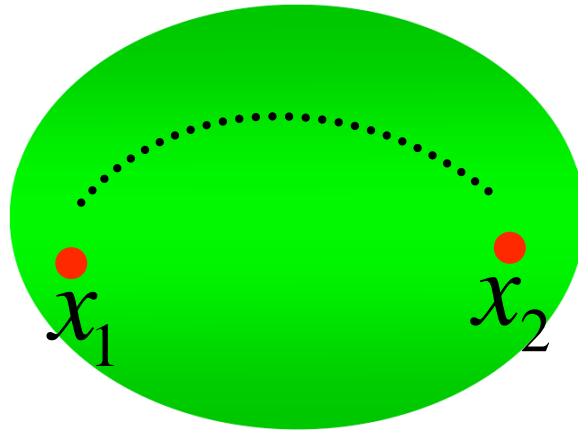
x_2 0100100110100

x_3 0101101001010

firing rates: $r_1, r_2; r_{12}$

correlation—covariance?

Correlations of Neural Firing



$$\{p(x_1, x_2)\}$$

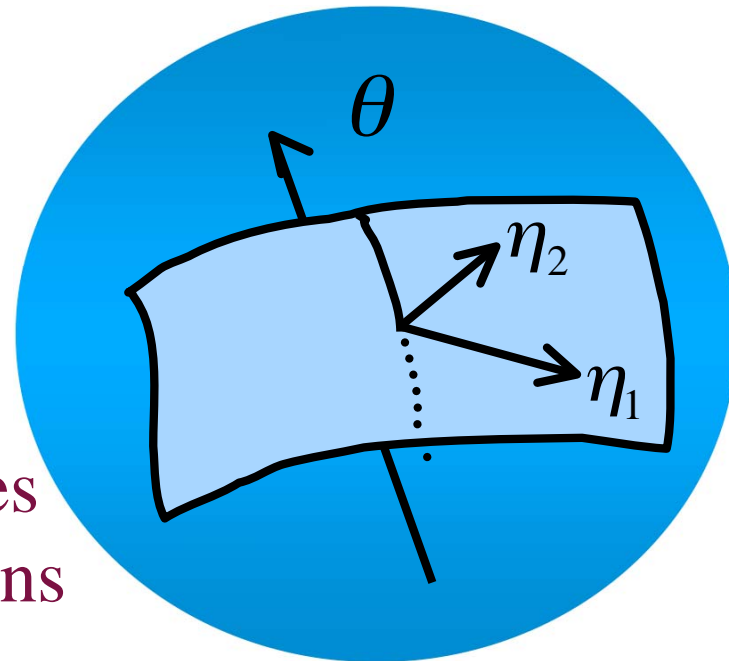
$$\{p_{00}, p_{10}, p_{01}, p_{11}\}$$

$$r_1 = p_{1\cdot} = p_{10} + p_{11}$$

$$r_2 = p_{\cdot 1} = p_{01} + p_{11}$$

$$\theta = \log \frac{p_{11} p_{00}}{p_{10} p_{01}}$$

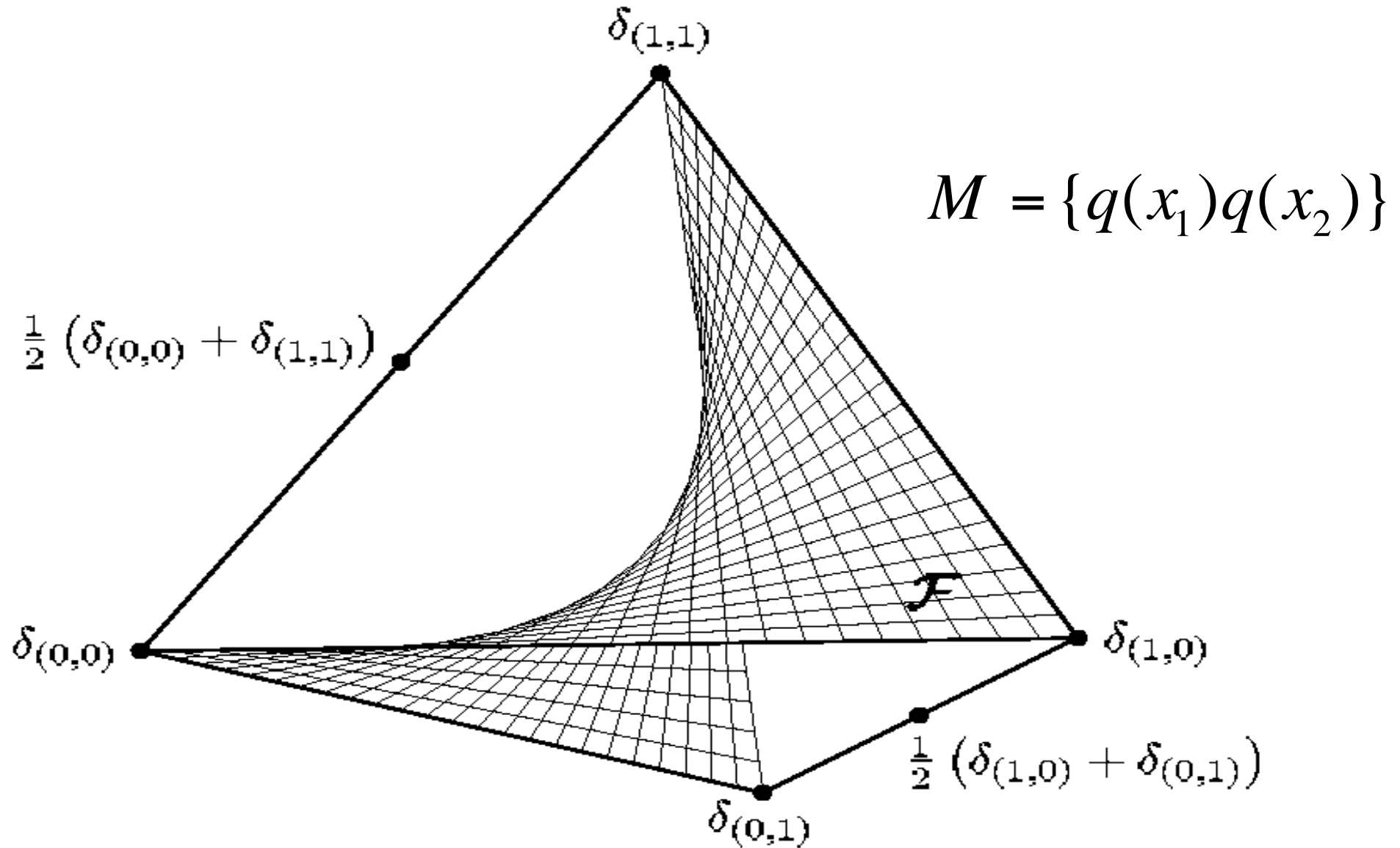
firing rates
correlations



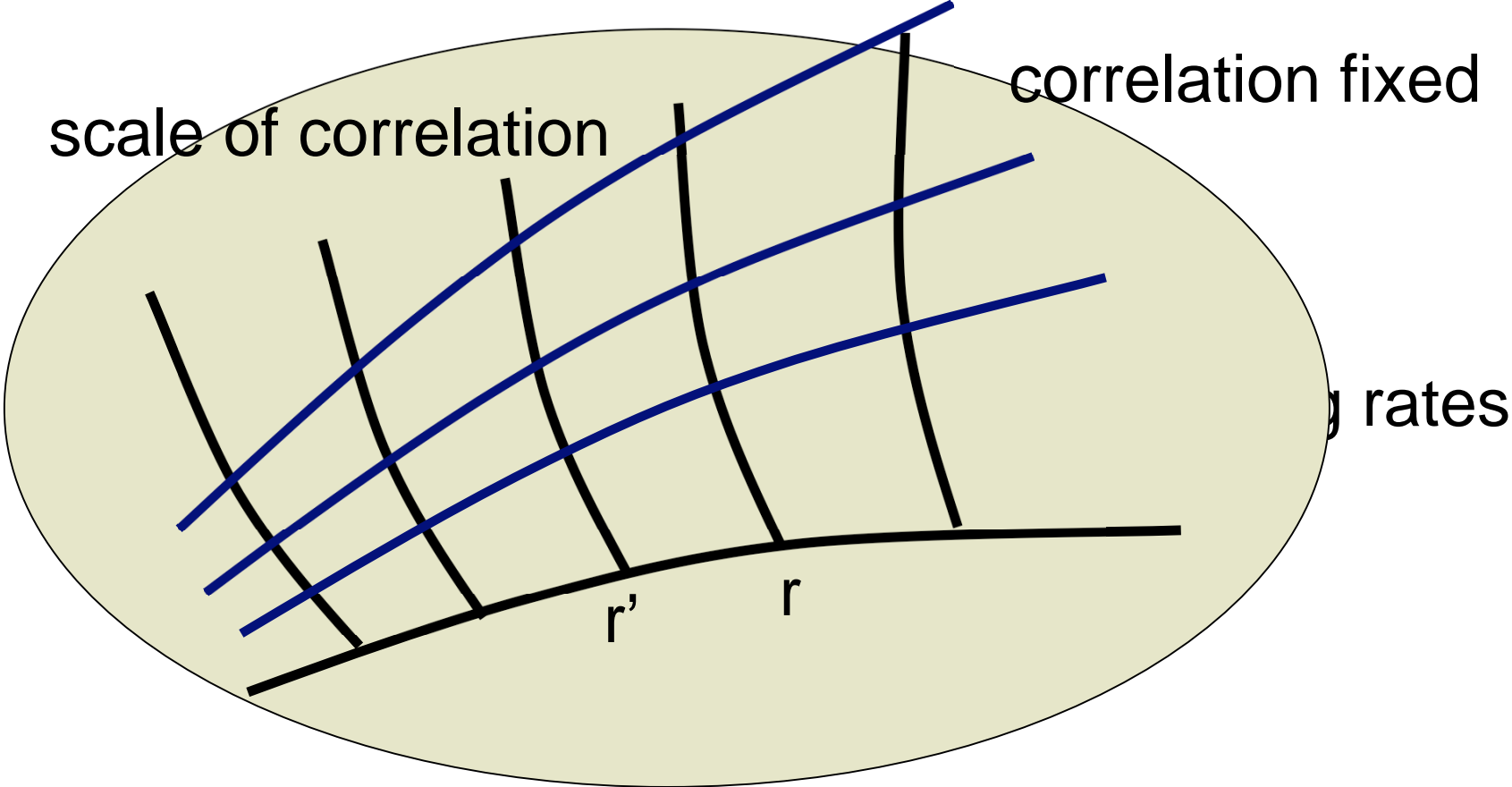
$$\{(r_1, r_2), \theta\}$$

orthogonal coordinates

Independent Distributions



Orthogonal Coordinates:



two neuron case

$$r_1, r_2, r_{12}; \theta_1, \theta_2, \theta_{12}$$

$$\theta_{12} = \log \frac{p_{00} p_{11}}{p_{01} p_{10}} = \log \frac{r_{12} (1 + r_{12} - r_1 - r_2)}{(r_1 - r_{12})(r_2 - r_{12})}$$

$$r_{12} = f(r_1, r_2, \theta)$$

$$r_{12}(t) = f(r_1(t), r_2(t), \theta)$$

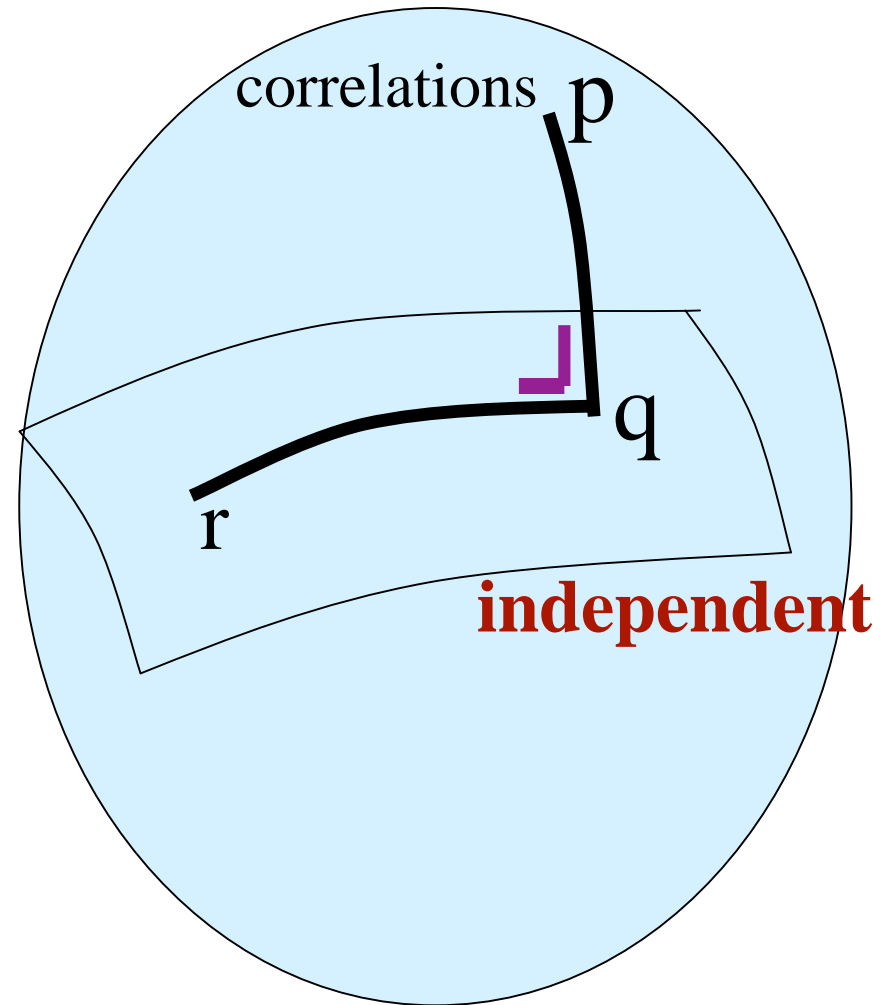
Decomposition of KL-divergence

$$D[p:r] = D[p:q] + D[q:r]$$

p, q : same marginals η_1, η_2

r, q : same correlations θ

$$D[p:r] = \sum_x p(x) \log \frac{p(x)}{q(x)}$$



pairwise correlations

covariance: $C_{ij} = r_{ij} - r_i r_j$ not orthogonal

independent distributions

$$r_{ij} = r_i r_j, \quad r_{ijk} = r_i r_j r_k, \dots$$

How to generate correlated spikes?
(Niebur, Neural Computation [2007])

higher-order correlations

Orthogonal higher-order correlations

$$\boldsymbol{\theta} = \left(\theta_i, \theta_{ij}; \quad L, \theta_{1L n} \right)$$

$$\boldsymbol{r} = \left(r_i, r_{ij}; \quad L, r_{1L n} \right)$$

tractable models of distributions $M = \{p(\mathbf{x})\}$

Full model: $2^n - 1$ parameters

Homogeneous model: n parameters

Boltzmann machine: only pairwise
correlations

Mixture model

Models of Correlated Spike Distributions (1)

Reference	$y(t)$	0110001011.....
	$x_0(t)$	1001011001.....
	$x_1(t)$	0101101101.....
	$x_2(t)$

Models of Correlated Spike Distributions (2)

$$x_i(t) = \%_i(t) \circ y(t)$$

additive interaction

eliminating interaction

Niebur replacement

$y(t)$ 0110001011

$\%_1(t)$ 1001001101

$\%_2(t)$ 1100010011

Mixture Model

$y(t)$: 0110001011 mixing sequence

$$p_1(\mathbf{x}; \mathbf{u}) = \prod p(x_i, u_i) \quad \text{when } y = 1$$

$$p_2(\mathbf{x}; \mathbf{u}) = \prod p(x_i, v_i) \quad \text{when } y = 0$$

$p(x, u)$: $x = 1$ with probability u

$$p(\mathbf{x}) = \sum p(y) p(\mathbf{x} | y)$$

Mixture model

$p(\mathbf{x}; \mathbf{u}) = \prod p(x_i, u_i)$: independent distribution

$$p(\mathbf{x}; \mathbf{u}, \mathbf{v}, m) = mp(\mathbf{x}; \mathbf{u}) + \bar{m}p(\mathbf{x}; \mathbf{v})$$

$$\mathbf{u} = (u_1, \dots, u_n) \quad \bar{m} = 1 - m$$

$$\mathbf{v} = (v_1, \dots, v_n)$$

$$M_n = \{p(\mathbf{x}; \mathbf{u}, \mathbf{v}, m)\}, \quad S_n = \{p(\mathbf{x})\}$$

firing rates

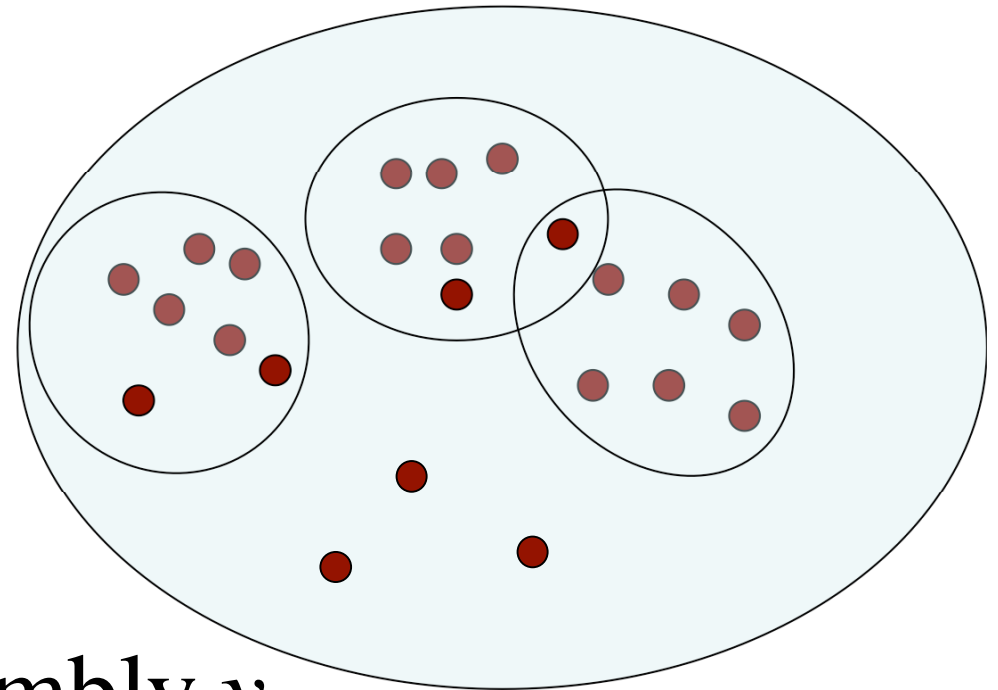
$$r_i = mu_i + \bar{m}v_i$$

$$r_{ij} = mu_i u_j + \bar{m}v_i v_j$$

$$r_{ijk} = mu_i u_j u_k + \bar{m}v_i v_j v_k$$

L

Conditional distributions and mixture



$$y = 1, 2, 3, \dots$$

$p(\mathbf{x} | y)$: Hebb assembly y

$$p(\mathbf{x}) = \sum p(y) p(\mathbf{x} | y)$$

State transition of Hebb assemblies mixture and covariances

$$p(\mathbf{x}, t) = (1-t)p_u(\mathbf{x}) + tp_v(\mathbf{x})$$

$$c_{ij}(t) = (1-t)c_{ij}(u) + tc_{ij}(v) + t(1-t)(u_i - v_i)(u_j - v_j)$$

Extra increase (decrease) of covariances

within Hebb assemblies---- increase

between Hebb assemblies ---decrease

Temporal correlations

$$\{x(t)\}, t = 1, 2, \dots, T$$

$$\text{Prob}\{x(1), \dots, x(T)\}$$

Poisson sequence:

Markov chain

Renewal process

Orthogonal parameter of correlation ---- Markov case

$$p(x_{t+1} | x_t)$$

$$a_{ij} = \Pr\{x_{t+1} = i | x_t = j\}$$

$$\Pr\{x_t\} = \Pr(x_1) \prod p(x_{t+1} | x_t)$$

$$r_1 = \Pr\{x_t = 1\}$$

$$c = r_{11} - r_1^2$$

$$\theta = \log \frac{a_{11}a_{00}}{a_{10}a_{01}}$$

Spatio-temporal correlation

correlated Poisson

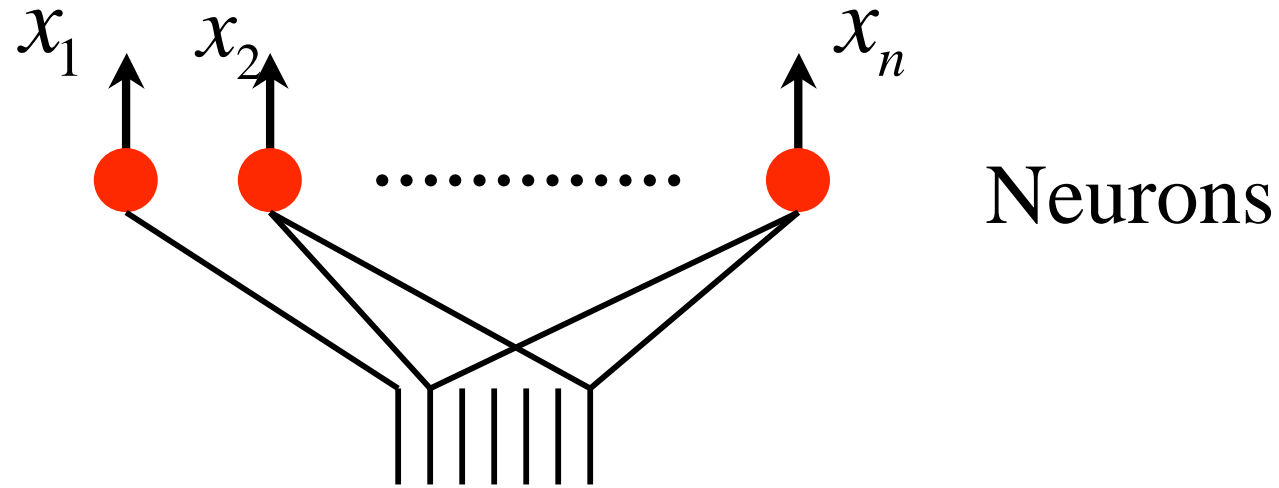
$$p(\{\mathbf{x}_t\}) = \prod_{t=1}^T p(\mathbf{x}_t)$$

Correlated renewal process

$\Pr(x_t) = k(t - t')$: t' last spike time

$$\Pr(x_{i,t}) = k(t - t') p(x_i | \mathbf{x})$$

Population and Synfire



$$x_i = 1(u_i)$$

$$u_i = \text{Gaussian}$$

$$E[u_i u_j] = \alpha$$

Population and Synfire

$$u_i = \sum w_{ij} s_j - h$$

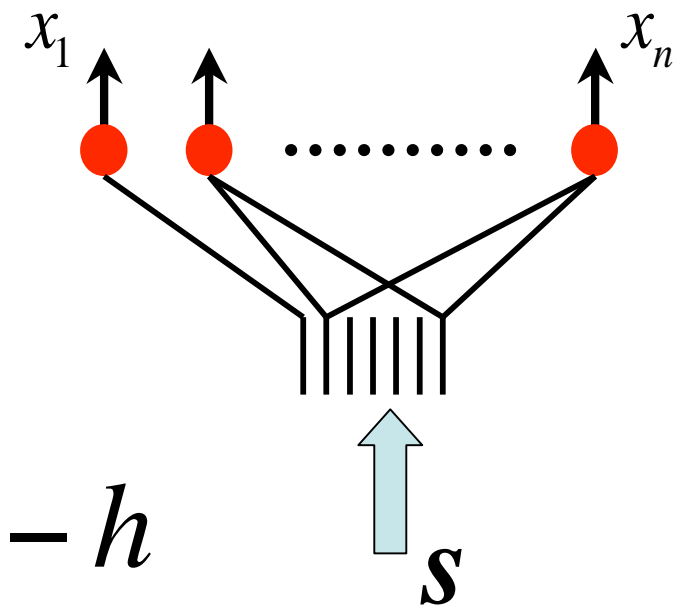
$$x_i = 1[u_i]$$

$$u_i = \sqrt{(1-\alpha)}\delta_i + \sqrt{\alpha}\varepsilon - h$$

$$\delta_i, \varepsilon \approx N(0, 1)$$

$$E[u_i u_j] = \alpha$$

$$E[u_i^2] = 1$$



$$p_i = \text{Prob} \{ i \text{ neurons fire at the same time} \}$$

$$r = \frac{i}{n} \quad P_r = \text{Pr}\{nr \text{ neurons fire}\}$$

$$q(r, \alpha) = e^{nH(r)} \int e^{-nz(\varepsilon)} d\varepsilon$$

$$z(\varepsilon) = \frac{\varepsilon^2}{2n} - r \log F - (1-r) \log (1-F)$$

$$F = F(a\varepsilon - h) = \frac{1}{\sqrt{2\pi}} \int_0^{a\varepsilon - h} e^{-\frac{t^2}{2}} dt$$

$$q(r, \alpha) = c \exp\left[\frac{2\alpha - 1}{2(1 - \alpha)} \left\{ F^{-1}(\alpha) - \frac{\sqrt{\alpha}}{2\alpha - 1} h \right\}^2\right]$$

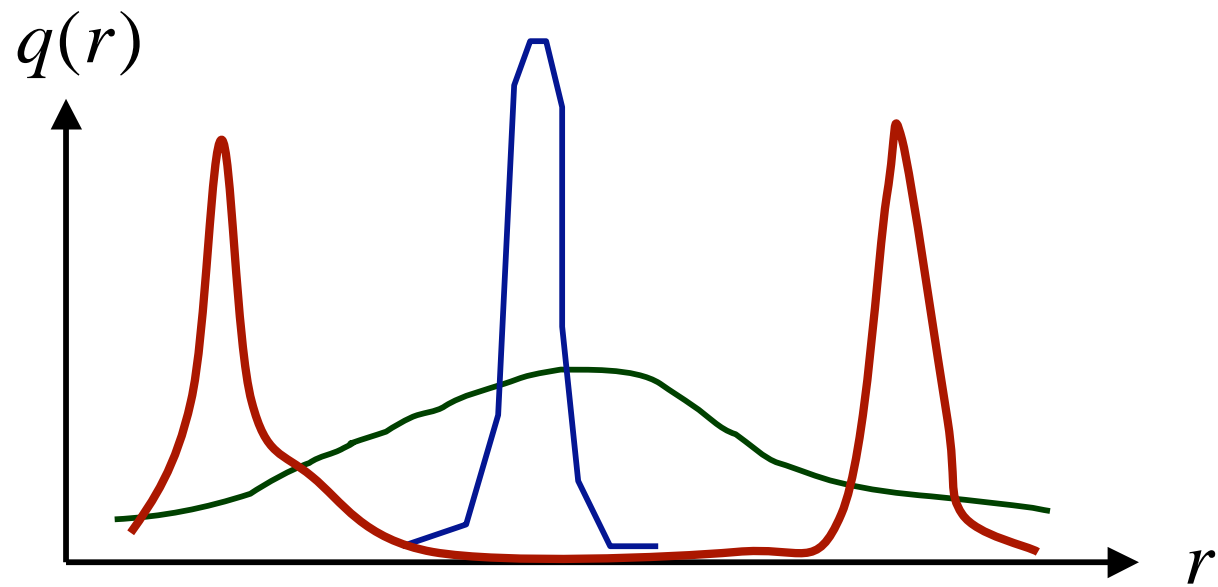
$$p(\mathbf{x}, \theta) = \exp\left\{ \sum \theta_i x_i + \sum \theta_{ij} x_i x_j + \sum \theta_{ijk} x_i x_j x_k + \dots \right\}$$

$$\theta_{i_1 i_2 \dots i_k} = O(1/n^{k-1})$$

Synfiring

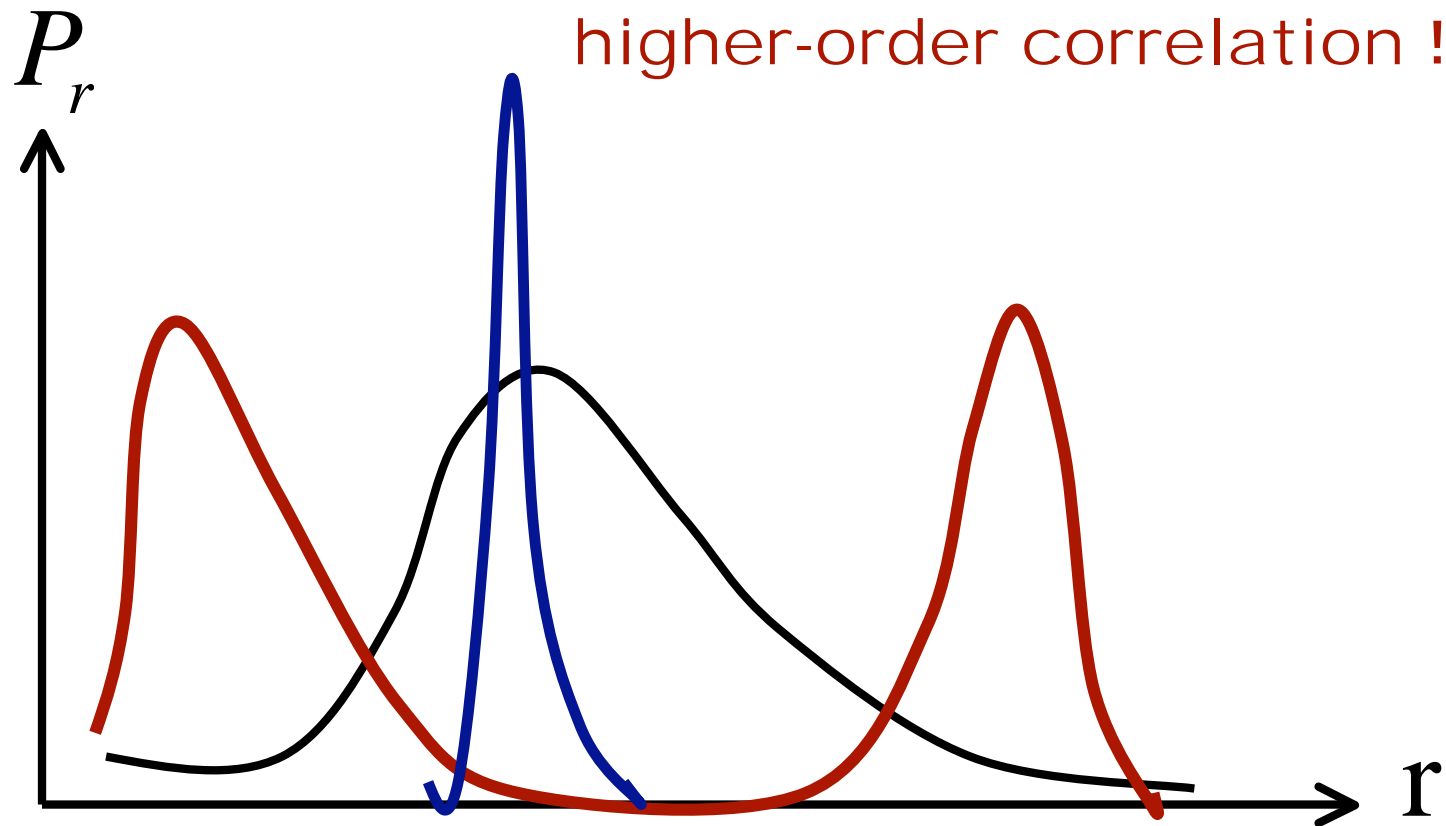
$$p(\mathbf{x}) = p(x_1, \dots, x_n)$$

$$r = \frac{1}{n} \sum x_i \quad q(r)$$



Bifurcation

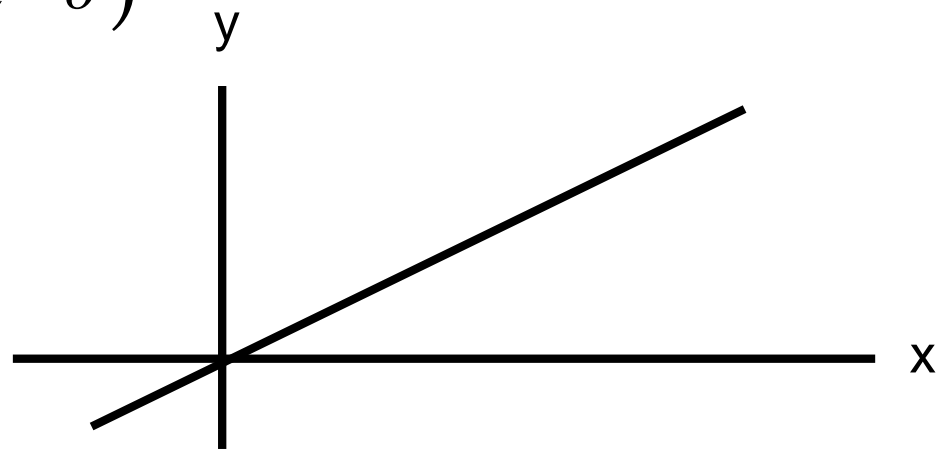
x_i : independent---single delta peak
pairwise correlated



Semiparametric Statistics

$$M = \{p(x, \theta, r)\}$$

$$p(x, \theta) = r(x - \theta)$$



$$y = \theta x$$

$$\begin{cases} y_i = \theta \xi_i + \varepsilon_i \\ x_i = \xi_i + \varepsilon_i' \end{cases} \quad p(x, y; \theta) = \int p(x, y; \xi, \theta) r(\xi) d\xi$$

mle, least square, total least square

Linear Regression: Semiparametrics

$$(x_1, y_1)$$

$$x_i = \xi_i + \varepsilon_i$$

$$(x_2, y_2)$$

$$y_i = \theta \xi_i + \varepsilon_i'$$

$$\vdots$$

$$(x_n, y_n)$$

$$\varepsilon_i, \varepsilon_i' \sim N(0, \sigma^2)$$

$$y = \theta x$$

Statistical Model

$$p(x, y | \theta, \xi) = c \exp \left\{ -\frac{1}{2} (x - \xi)^2 - \frac{1}{2} (y - \theta \xi)^2 \right\}$$

$$p(x_i, y_i | \theta, \xi_i) : \theta, \xi_1, \dots, \xi_n$$

$$p(x, y | \theta) = \int p(x, y | \theta, \xi) \mathcal{Z}(\xi) d\xi$$

———— semiparametric

Least squares?

$$L(\theta) = \sum (y_i - \theta x_i)^2 \rightarrow \min \quad : \hat{\theta} = \frac{\sum x_i y_i}{\sum x_i^2}$$

$$\frac{1}{n} \sum \frac{y_i}{x_i}, \quad \frac{\sum y_i}{\sum x_i}$$

mle, TLS

$$\sum (y_i - \theta x_i)(\theta y_i + x_i) = 0$$

Neyman-Scott

Semiparametric statistical model

$$x_1, x_2, \dots \quad p(x, \theta, Z)$$

Estimating function

$$y(x, \theta) \quad E_{\theta, Z} [y(x, \theta)] = 0$$

$$E_{\theta', Z} [y(x, \theta)] \neq 0$$

Estimating equation

$$\sum y(x_i, \theta) = 0 \quad \Rightarrow \hat{\theta}$$

$$u(x, y; \theta, Z) = \frac{\partial}{\partial \theta} \log p$$

$$= \partial_{\theta} s E[\xi | s] \dots$$

$$v(x, y; \theta, Z) = E[f(\xi) | s]$$

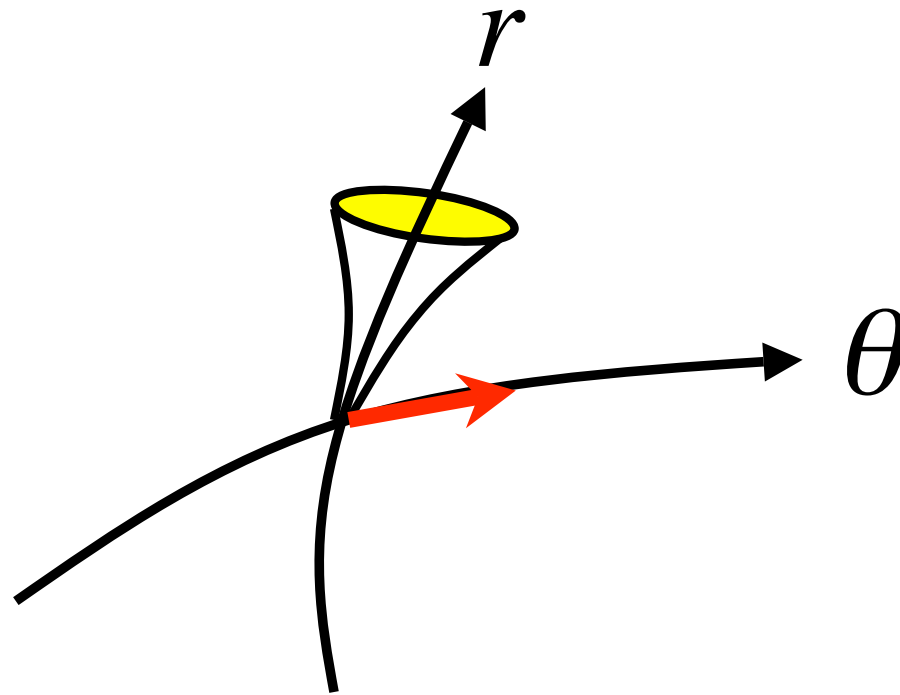
$$= k \{s(x, y; \theta)\}$$

$$u^I(x, y, \theta) = u - E[u | s]$$

$$= k(x + \theta y)(y - \theta x)$$

Fibre bundle

function space



$$z(\xi) \sim N(\mu_\xi, \sigma_\xi^2)$$

$$u^I(x, \theta, Z) = (x + \theta y + c)(y - \theta x)$$

$$c = \mu_\xi / \sigma_\xi^2$$

$$f(x, \theta; c) = (x + \theta y + c)(y - \theta x)$$

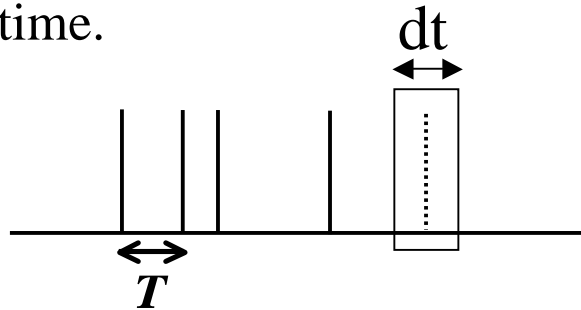
$$c = \mu_\xi / \sigma_\xi^2$$

$$\mu_\xi = \frac{1}{n} \sum x_i$$

$$\sigma_\xi^2 = \frac{1}{n} \sum x_i^2 - (\mu_\xi)^2 - \sigma^2$$

Poisson process

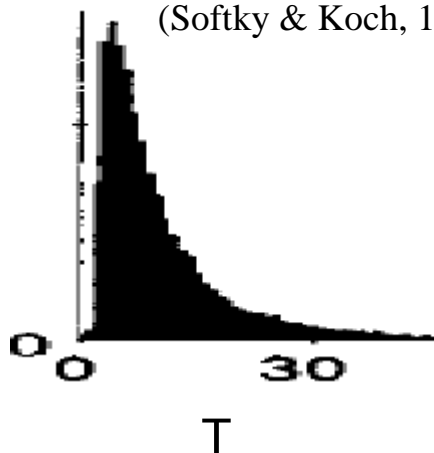
Poisson Process: Instantaneous firing rate is constant over time.



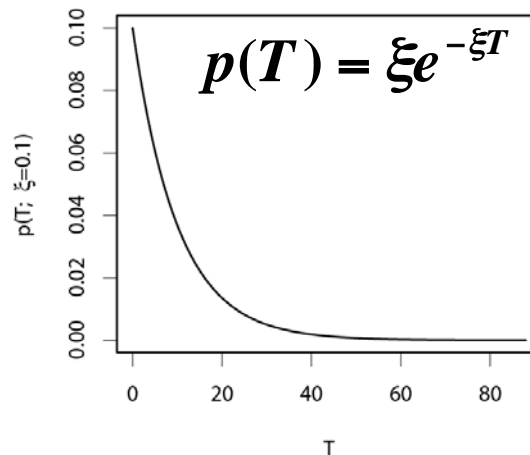
For every small time window dt , generate a spike with probability ξdt .

Cortical Neuron

(Softky & Koch, 1993)



Poisson Process



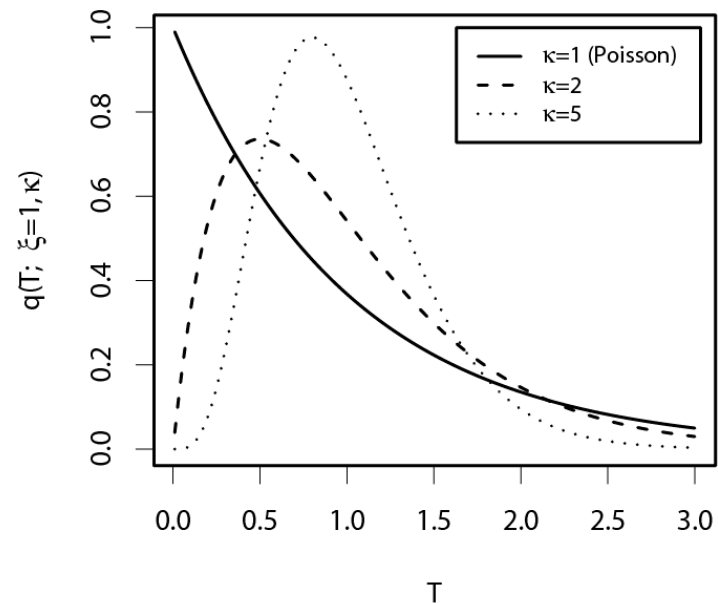
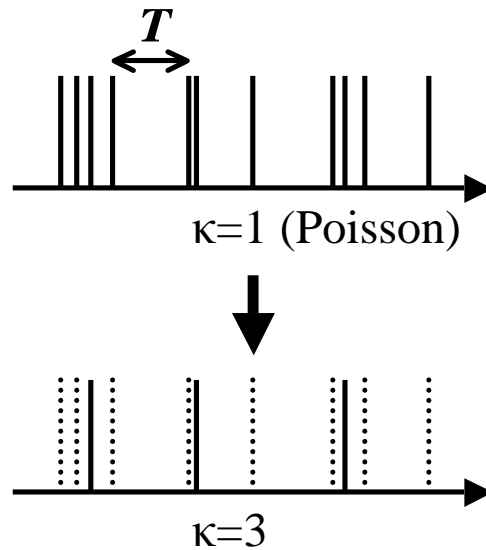
Poisson process
cannot explain inter-
spike interval
distributions.

Gamma distribution

Gamma Distribution: Every κ -th spike of the Poisson process is left.

$$q(T; \xi, \kappa) = \frac{(\xi \kappa)^\kappa}{\Gamma(\kappa)} T^{\kappa-1} e^{-\xi \kappa T}.$$

Two parameters $\begin{cases} \xi: \text{Firing rate} \\ \kappa: \text{Irregularity} \end{cases}$



Gamma distribution

$$f(T) = \frac{(r\kappa)^\kappa}{\Gamma(\kappa)} T^{\kappa-1} \exp\{-r\kappa T\}$$

$\kappa = 1$: Poisson

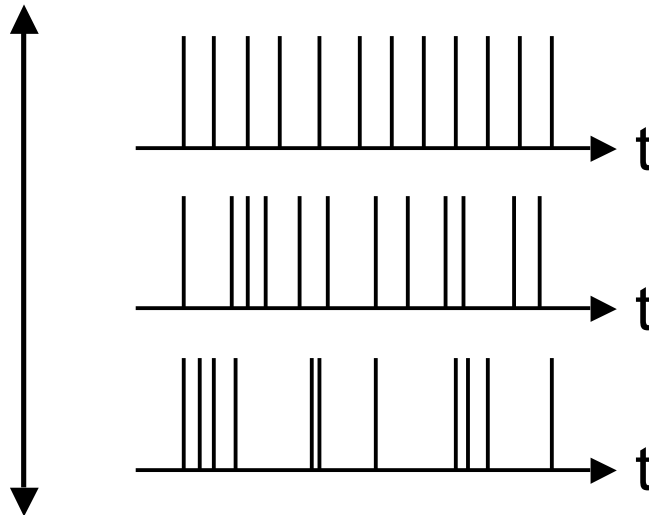
$\kappa \rightarrow \infty$: regular

Integrate-and fire

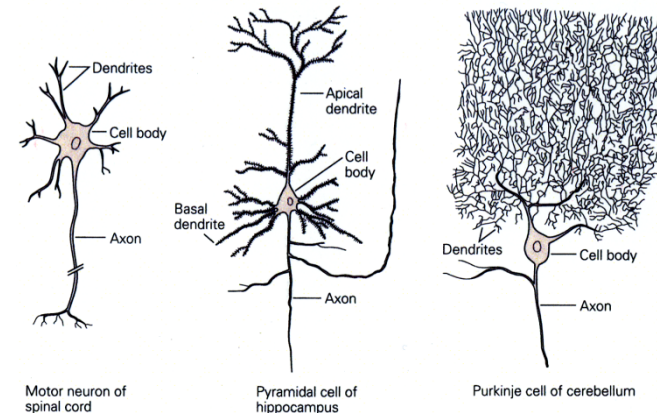
Markov model

Irregularity κ is unique to individual neurons.

Regular (large κ)



Irregular (small κ)



Irregularity varies among neurons.

→ **We assume that κ is independent of time.**

Information geometry of estimating function

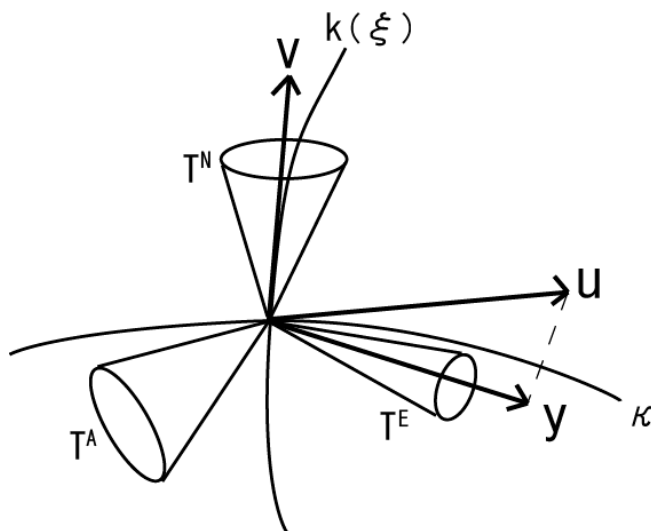
• Estimating function $y(T, \kappa)$:

$$\rightarrow \sum_{l=1}^N y(T_l; \kappa) = 0$$

$$E[y(T, \kappa)] = 0$$

• Maximum likelihood Method:

$$\leftrightarrow \frac{d}{d\kappa} \log p(T_1) \cdots p(T_N) = \sum_{l=1}^N u(T_l; \kappa) = 0$$



(See poster for details)

How to obtain an estimating function y :

$$\text{Score function } s \begin{cases} u(T; \hat{\epsilon}, k) \equiv \frac{d \log p(T; \hat{\epsilon}, k)}{d\hat{\epsilon}} \\ v(T; \hat{\epsilon}, k) \equiv \frac{\delta \log p(T; \hat{\epsilon}, k)}{\delta k(\hat{\epsilon})} \end{cases}$$

$$\rightarrow y = u - \frac{\langle u \cdot v \rangle}{\langle v \cdot v \rangle} v$$

temporal correlations

$$\circ: x(1)x(2)\dots x(N)$$

independent with firing probability $r(t)$

→ spike counts: Poisson
ISI: exponential

renewal:

$$r(t) = k(t - t_i) \quad : \text{last spike}$$

ISI distribution $f(t)$

$$f(T) = ck(T) \exp \left\{ \int_0^T k(t) dt \right\}$$

Estimation of κ by estimating functions

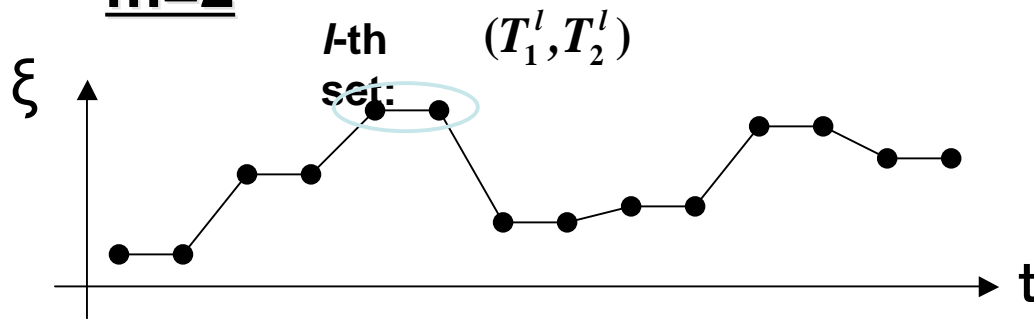
 No estimating function exists if the neighboring firing rates are different.

 $m(\geq 2)$ consecutive observations must have the same firing rate.

Example:

$m=2$

$$\text{Model: } p(T_1, T_2; \kappa, k) = \int_0^{\infty} q(T_1; \xi, \kappa) q(T_2; \xi, \kappa) k(\xi) d\xi$$

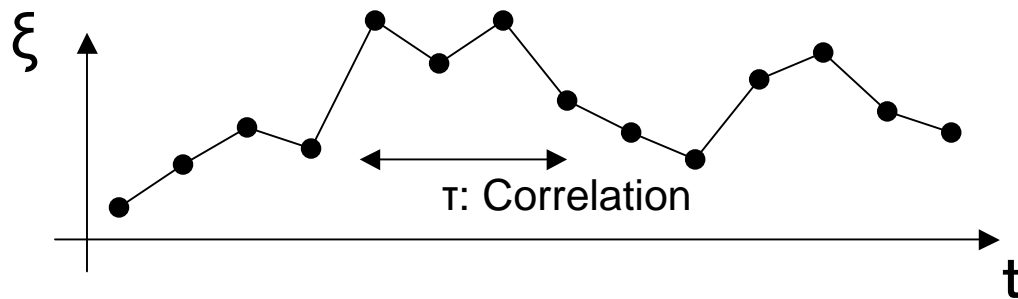


Estimating function:
($E[y] = 0$)

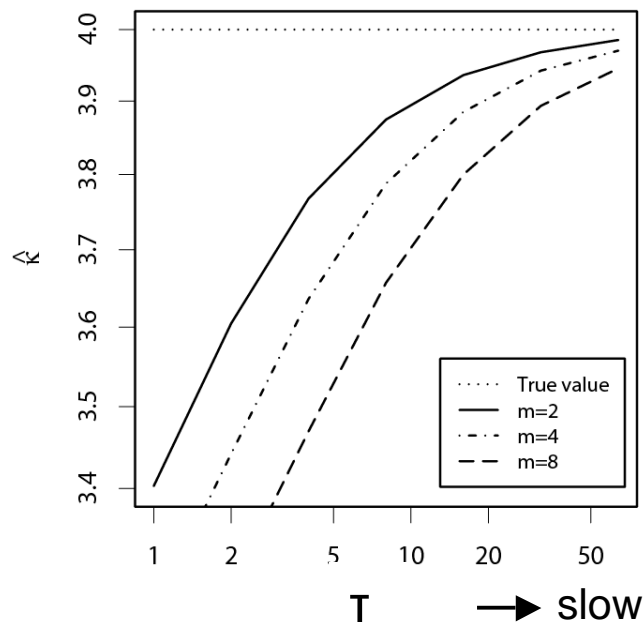
$$y = \log \frac{T_1 T_2}{(T_1 + T_2)^2} + 2\ddot{o}(2\hat{e}) - 2\ddot{o}(\hat{e})$$

$$\bar{y} = \frac{1}{N} \sum_{l=1}^m \log \frac{T_1^l T_2^l}{(T_1^l + T_2^l)^2} + 2\ddot{o}(2\hat{e}) - 2\ddot{o}(\hat{e}) = 0$$

Case of $m=1$ (spontaneous discharge)



Firing rate continuously changes and is driven by Gaussian noise.



κ can be approximated if the firing rate changes slowly.

Estimating function ($m=2$):

$$\bar{y} = \frac{1}{N} \sum_{l=1}^m \log \frac{T_1^l T_2^l}{(T_1^l + T_2^l)} + 2\ddot{o}(2\hat{e}) - 2\ddot{o}(\hat{e}) = 0$$