The exactly solvable log-gamma polymer

Timo Seppäläinen

Department of Mathematics University of Wisconsin-Madison

2011







Collaborators: Nicos Georgiou (Utah), Ivan Corwin (Microsoft/MIT), Neil O'Connell (Warwick), Nikos Zygouras (Warwick)

Context. Among directed polymer models there are currently two 1+1 dimensional exactly solvable cases:

 polymer in a Brownian environment with continuous-time random walk paths, discovered by O'Connell-Yor (2001), subsequently worked on by O'Connell, O'C–Moriarty, and Valkó–S. **Context.** Among directed polymer models there are currently two 1+1 dimensional exactly solvable cases:

- polymer in a Brownian environment with continuous-time random walk paths, discovered by O'Connell-Yor (2001), subsequently worked on by O'Connell, O'C–Moriarty, and Valkó–S.
- lattice polymer whose weights are i.i.d. log-gamma distributed

Context. Among directed polymer models there are currently two 1+1 dimensional exactly solvable cases:

- polymer in a Brownian environment with continuous-time random walk paths, discovered by O'Connell-Yor (2001), subsequently worked on by O'Connell, O'C–Moriarty, and Valkó–S.
- lattice polymer whose weights are i.i.d. log-gamma distributed

This talk is about the log-gamma polymer.





 Π_{m,n} = set of up-right lattice paths x. = (x_k) from (1, 1) to (m, n)



- $\Pi_{m,n}$ = set of up-right lattice paths $x_{.} = (x_k)$ from (1,1) to (m, n)
- environment of i.i.d. weights under \mathbb{P} : $\{\omega(x): x \in \mathbb{N}^2\}$



- $\Pi_{m,n}$ = set of up-right lattice paths $x_{\cdot} = (x_k)$ from (1,1) to (m, n)
- environment of i.i.d. weights under \mathbb{P} : $\{\omega(x): x \in \mathbb{N}^2\}$
- quenched probability measure on $\Pi_{m,n}$:

$$Q_{m,n}(x_{\cdot}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$



- $\Pi_{m,n}$ = set of up-right lattice paths $x_{.} = (x_k)$ from (1,1) to (m, n)
- environment of i.i.d. weights under \mathbb{P} : $\{\omega(x): x \in \mathbb{N}^2\}$
- quenched probability measure on $\Pi_{m,n}$:

$$Q_{m,n}(x_{\cdot}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$

• inverse temperature $\beta > 0$



- $\Pi_{m,n}$ = set of up-right lattice paths $x_{\cdot} = (x_k)$ from (1,1) to (m, n)
- environment of i.i.d. weights under \mathbb{P} : $\{\omega(x): x \in \mathbb{N}^2\}$
- quenched probability measure on $\Pi_{m,n}$:

$$Q_{m,n}(x_{\cdot}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$

• inverse temperature $\beta > 0$

• partition function
$$Z_{m,n} = \sum_{x_{\star} \in \Pi_{m,n}} \exp \left\{ \beta \sum_{k} \omega(x_{k}) \right\}$$

• Quenched measure
$$Q_{m,n}(x_{.}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$

• Partition function $Z_{m,n} = \sum_{x_{.} \in \Pi_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$

Questions:

• Behavior of walk x_{\cdot} under $Q_{m,n}$ on large scales.

• Quenched measure
$$Q_{m,n}(x_{\cdot}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$

• Partition function $Z_{m,n} = \sum_{x_{\cdot} \in \Pi_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$

Questions:

- Behavior of walk x_{\cdot} under $Q_{m,n}$ on large scales.
- Behavior of log $Z_{m,n}$ (now also random as a function of ω).

• Quenched measure
$$Q_{m,n}(x_{\cdot}) = \frac{1}{Z_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$$

• Partition function $Z_{m,n} = \sum_{x_{\cdot} \in \Pi_{m,n}} \exp\left\{\beta \sum_{k} \omega(x_{k})\right\}$

Questions:

- Behavior of walk x_{\cdot} under $Q_{m,n}$ on large scales.
- Behavior of log $Z_{m,n}$ (now also random as a function of ω).

In this talk focus is on $\log Z$.

- Fix $\beta = 1$.
- Pick a parameter $0 < \mu < \infty$.

• Fix $\beta = 1$.

- Pick a parameter $0 < \mu < \infty$.
- Let $Y_{i,j} = e^{\omega(i,j)}$ be i.i.d. Gamma⁻¹(μ) distributed,

in other words, $Y_{i,j}^{-1} \sim \text{Gamma}(\mu)$.

• Fix $\beta = 1$.

- Pick a parameter $0 < \mu < \infty$.
- Let Y_{i,j} = e^{ω(i,j)} be i.i.d. Gamma⁻¹(μ) distributed,
 in other words, Y⁻¹_{i,j} ~ Gamma(μ).
- Gamma(μ) density: $f(x) = \Gamma(\mu)^{-1} x^{\mu-1} e^{-x}$ on \mathbb{R}_+

• Fix $\beta = 1$.

- Pick a parameter $0 < \mu < \infty$.
- Let Y_{i,j} = e^{ω(i,j)} be i.i.d. Gamma⁻¹(μ) distributed,
 in other words, Y⁻¹_{i,j} ~ Gamma(μ).
- Gamma(μ) density: $f(x) = \Gamma(\mu)^{-1} x^{\mu-1} e^{-x}$ on \mathbb{R}_+
- $\mathbb{E}(\log Y) = -\Psi_0(\mu)$ (digamma function) $\mathbb{V}ar(\log Y) = \Psi_1(\mu)$ (trigamma function) where $\Psi_n(s) = (d^{n+1}/ds^{n+1})\log\Gamma(s)$

• Fix $\beta = 1$.

- Pick a parameter $0 < \mu < \infty$.
- Let Y_{i,j} = e^{ω(i,j)} be i.i.d. Gamma⁻¹(μ) distributed,
 in other words, Y⁻¹_{i,j} ~ Gamma(μ).
- Gamma(μ) density: $f(x) = \Gamma(\mu)^{-1} x^{\mu-1} e^{-x}$ on \mathbb{R}_+
- $\mathbb{E}(\log Y) = -\Psi_0(\mu)$ (digamma function) $\mathbb{V}ar(\log Y) = \Psi_1(\mu)$ (trigamma function) where $\Psi_n(s) = (d^{n+1}/ds^{n+1})\log\Gamma(s)$

What is special about this choice?

• Parameters $0 < \theta < \mu$.

- Parameters $0 < \theta < \mu$.
- Bulk weights $Y_{i,j}$ for $i,j \in \mathbb{N} = \{1,2,3,\dots\}$

- Parameters $0 < \theta < \mu$.
- Bulk weights $Y_{i,j}$ for $i,j \in \mathbb{N} = \{1,2,3,\dots\}$
- Boundary weights $U_{i,0} = Y_{i,0}$ and $V_{0,j} = Y_{0,j}$.

- Parameters $0 < \theta < \mu$.
- Bulk weights $Y_{i,j}$ for $i,j \in \mathbb{N} = \{1,2,3,\dots\}$
- Boundary weights $U_{i,0} = Y_{i,0}$ and $V_{0,j} = Y_{0,j}$.



- Parameters $0 < \theta < \mu$.
- Bulk weights $Y_{i,j}$ for $i,j \in \mathbb{N} = \{1,2,3,\dots\}$
- Boundary weights $U_{i,0} = Y_{i,0}$ and $V_{0,j} = Y_{0,j}$.



$$U_{m,n} = \frac{Z_{m,n}}{Z_{m-1,n}} \qquad V_{m,n} = \frac{Z_{m,n}}{Z_{m,n-1}} \qquad X_{m,n} = \left(\frac{Z_{m,n}}{Z_{m+1,n}} + \frac{Z_{m,n}}{Z_{m,n+1}}\right)^{-1}$$

$$U_{m,n} = \frac{Z_{m,n}}{Z_{m-1,n}} \qquad V_{m,n} = \frac{Z_{m,n}}{Z_{m,n-1}} \qquad X_{m,n} = \left(\frac{Z_{m,n}}{Z_{m+1,n}} + \frac{Z_{m,n}}{Z_{m,n+1}}\right)^{-1}$$

For an undirected edge
$$f$$
: $T_f = \begin{cases} U_x & f = \{x - e_1, x\} \\ V_x & f = \{x - e_2, x\} \end{cases}$

$$U_{m,n} = \frac{Z_{m,n}}{Z_{m-1,n}} \qquad V_{m,n} = \frac{Z_{m,n}}{Z_{m,n-1}} \qquad X_{m,n} = \left(\frac{Z_{m,n}}{Z_{m+1,n}} + \frac{Z_{m,n}}{Z_{m,n+1}}\right)^{-1}$$

.

For an undirected edge
$$f$$
: $T_f = \begin{cases} U_x & f = \{x - e_1, x\} \\ V_x & f = \{x - e_2, x\} \end{cases}$



- down-right path (z_k) with edges $f_k = \{z_{k-1}, z_k\}, k \in \mathbb{Z}$
- interior points \mathcal{I} of path (z_k)

$$U_{m,n} = \frac{Z_{m,n}}{Z_{m-1,n}} \qquad V_{m,n} = \frac{Z_{m,n}}{Z_{m,n-1}} \qquad X_{m,n} = \left(\frac{Z_{m,n}}{Z_{m+1,n}} + \frac{Z_{m,n}}{Z_{m,n+1}}\right)^{-1}$$

.

For an undirected edge
$$f$$
: $T_f = \begin{cases} U_x & f = \{x - e_1, x\} \\ V_x & f = \{x - e_2, x\} \end{cases}$



- down-right path (z_k) with edges $f_k = \{z_{k-1}, z_k\}, k \in \mathbb{Z}$
- interior points \mathcal{I} of path (z_k)

Burke property

Theorem

Variables $\{T_{f_k}, X_z : k \in \mathbb{Z}, z \in \mathcal{I}\}$ are independent with marginals $U^{-1} \sim \text{Gamma}(\theta), \quad V^{-1} \sim \text{Gamma}(\mu - \theta), \text{ and } X^{-1} \sim \text{Gamma}(\mu).$

Burke property

Theorem

Variables $\{T_{f_k}, X_z : k \in \mathbb{Z}, z \in \mathcal{I}\}$ are independent with marginals $U^{-1} \sim \text{Gamma}(\theta), \quad V^{-1} \sim \text{Gamma}(\mu - \theta), \text{ and } X^{-1} \sim \text{Gamma}(\mu).$

"Burke property" because the analogous property for last-passage percolation with exponential weights is a generalization of Burke's Theorem for M/M/1 queues.

Partition functions

Partition functions

• Limits:
$$p(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}}{n}$$
 and $p^{\theta}(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}^{\theta}}{n}$

Partition functions

• Limits:
$$p(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}}{n}$$
 and $p^{\theta}(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}^{\theta}}{n}$



$$\log Z_{ns,nt}^{\theta} = \sum_{j=1}^{nt} \log V_{0,j} + \sum_{i=1}^{ns} \log U_{i,nt}$$

Partition functions

• Limits:
$$p(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}}{n}$$
 and $p^{\theta}(s,t) = \lim_{n \to \infty} \frac{\log Z_{ns,nt}^{\theta}}{n}$



$$\log Z_{ns,nt}^{\theta} = \sum_{j=1}^{nt} \log V_{0,j} + \sum_{i=1}^{ns} \log U_{i,nt}$$

hence $p^{\theta}(s,t) = -t\Psi_0(\mu - \theta) - s\Psi_0(\theta)$

Now to compute p(s, t):






$$p^{\theta}(s,t) = \sup_{0 \le a \le s} \{-a\Psi_0(\theta) + p(s-a,t)\} \bigvee \sup_{0 \le b \le t} \{-b\Psi_0(\mu-\theta) + p(s,t-b)\}$$



$$p^{\theta}(s,t) = \sup_{0 \le a \le s} \{-a\Psi_0(\theta) + p(s-a,t)\} \bigvee \sup_{0 \le b \le t} \{-b\Psi_0(\mu-\theta) + p(s,t-b)\}$$

Vary θ . Specializes to a convex duality that can be inverted to give



$$p^{\theta}(s,t) = \sup_{0 \leq a \leq s} \{-a\Psi_0(\theta) + p(s-a,t)\} \bigvee \sup_{0 \leq b \leq t} \{-b\Psi_0(\mu-\theta) + p(s,t-b)\}$$

Vary θ . Specializes to a convex duality that can be inverted to give

$$p(s,t) = \inf_{0 < \theta < \mu} \{-s\Psi_0(\theta) - t\Psi_0(\mu - \theta)\}$$

Burke property also serves as basis for deriving fluctuation exponents.

Burke property also serves as basis for deriving fluctuation exponents.

• If endpoint $(m, n) \rightarrow \infty$ in characteristic direction

 $|m - N\Psi_1(\mu - heta)| \leq CN^{2/3}$ and $|n - N\Psi_1(heta)| \leq CN^{2/3}$

then fluctuations have conjectured order of magnitude:

 $N^{1/3}$ for log $Z_{m,n}^{\theta}$ and $N^{2/3}$ for the path.

Burke property also serves as basis for deriving fluctuation exponents.

• If endpoint $(m, n) \rightarrow \infty$ in characteristic direction

 $|m - N\Psi_1(\mu - heta)| \leq CN^{2/3}$ and $|n - N\Psi_1(heta)| \leq CN^{2/3}$

then fluctuations have conjectured order of magnitude:

 $N^{1/3}$ for log $Z_{m,n}^{\theta}$ and $N^{2/3}$ for the path.

• Further away from the characteristic log $Z_{m,n}^{\theta}$ satisfies CLT.

Burke property also serves as basis for deriving fluctuation exponents.

• If endpoint $(m, n) \rightarrow \infty$ in characteristic direction

 $|m - N\Psi_1(\mu - heta)| \leq CN^{2/3}$ and $|n - N\Psi_1(heta)| \leq CN^{2/3}$

then fluctuations have conjectured order of magnitude:

 $N^{1/3}$ for log $Z_{m,n}^{\theta}$ and $N^{2/3}$ for the path.

- Further away from the characteristic log $Z_{m,n}^{\theta}$ satisfies CLT.
- Upper bounds hold for i.i.d. model without boundaries.

Explicit large deviations for $\log Z$

L.m.g.f. of log Y, $Y \sim \Gamma^{-1}(\mu)$:

$$M_{\mu}(\xi) = \log \mathbb{E}(e^{\xi \log Y}) = \begin{cases} \log \Gamma(\mu - \xi) - \log \Gamma(\mu) & \xi \in (-\infty, \mu) \\ \infty & \xi \in [\mu, \infty). \end{cases}$$

Explicit large deviations for $\log Z$

L.m.g.f. of log Y, $Y \sim \Gamma^{-1}(\mu)$:

$$M_{\mu}(\xi) = \log \mathbb{E}(e^{\xi \log Y}) = \begin{cases} \log \Gamma(\mu - \xi) - \log \Gamma(\mu) & \xi \in (-\infty, \mu) \\ \infty & \xi \in [\mu, \infty). \end{cases}$$

For i.i.d. $\Gamma^{-1}(\mu)$ model, let

$$\Lambda_{s,t}(\xi) = \lim_{n \to \infty} n^{-1} \log \mathbb{E}(e^{\xi \log Z_{ns,nt}}), \qquad \xi \in \mathbb{R}$$

Explicit large deviations for $\log Z$

L.m.g.f. of log Y, $Y \sim \Gamma^{-1}(\mu)$:

$$M_{\mu}(\xi) = \log \mathbb{E}(e^{\xi \log Y}) = \begin{cases} \log \Gamma(\mu - \xi) - \log \Gamma(\mu) & \xi \in (-\infty, \mu) \\ \infty & \xi \in [\mu, \infty). \end{cases}$$

For i.i.d. $\Gamma^{-1}(\mu)$ model, let

$$\Lambda_{s,t}(\xi) = \lim_{n \to \infty} n^{-1} \log \mathbb{E}(e^{\xi \log Z_{ns,nt}}), \qquad \xi \in \mathbb{R}$$

Theorem.

$$egin{aligned} \Lambda_{s,t}(\xi) &= egin{cases} p(s,t)\xi & \xi < 0 \ &\inf_{ heta \in (\xi,\mu)} ig\{ t M_ heta(\xi) - s M_{\mu- heta}(-\xi) ig\} & 0 \leq \xi < \mu \ &\infty & \xi \geq \mu. \end{aligned}$$

• $\Lambda_{s,t}$ linear on \mathbb{R}_{-} because for r < p(s,t)

$$\lim_{n\to\infty} n^{-1}\log \mathbb{P}\{\log Z_{ns,nt} \leq nr\} = -\infty.$$

- $\Lambda_{s,t}$ linear on \mathbb{R}_{-} because for r < p(s,t) $\lim_{n \to \infty} n^{-1} \log \mathbb{P}\{\log Z_{ns,nt} \le nr\} = -\infty.$
- Right tail LDP: for $r \ge p(s, t)$

$$J_{s,t}(r) \equiv -\lim_{n \to \infty} n^{-1} \log \mathbb{P}\{\log Z_{ns,nt} \ge nr\} = \Lambda_{s,t}^*(r)$$

- $\Lambda_{s,t}$ linear on \mathbb{R}_- because for r < p(s,t) $\lim_{n \to \infty} n^{-1} \log \mathbb{P}\{\log Z_{ns,nt} \le nr\} = -\infty.$
- Right tail LDP: for $r \ge p(s, t)$

$$J_{s,t}(r) \equiv -\lim_{n \to \infty} n^{-1} \log \mathbb{P}\{\log Z_{ns,nt} \ge nr\} = \Lambda_{s,t}^*(r)$$

• Proof of formula for $\Lambda_{s,t}$ goes by first finding $J_{s,t}$ and then convex conjugation.

Starting point for proof of large deviations



$$Z_{ns,nt}^{\theta} = \sum_{\ell=1}^{nt} \left(\prod_{j=1}^{\ell} V_{0,j} \right) Z_{(1,\ell),(ns,nt)} + \sum_{k=1}^{ns} \left(\prod_{j=1}^{k} U_{j,0} \right) Z_{(k,1),(ns,nt)}$$

Starting point for proof of large deviations



Divide by $\prod_{j=1}^{nt} V_{0,j}$:

$$\prod_{i=1}^{ns} U_{i,nt} = \sum_{\ell=1}^{nt} \left(\prod_{j=\ell+1}^{nt} V_{0,j}^{-1} \right) Z_{(1,\ell),(ns,nt)} + \sum_{k=1}^{ns} \left(\prod_{j=1}^{nt} V_{0,j}^{-1} \right) \left(\prod_{i=1}^{k} U_{i,0} \right) Z_{(k,1),(ns,nt)}$$

Starting point for proof of large deviations



Divide by $\prod_{j=1}^{nt} V_{0,j}$:

$$\prod_{i=1}^{ns} U_{i,nt} = \sum_{\ell=1}^{nt} \left(\prod_{j=\ell+1}^{nt} V_{0,j}^{-1} \right) Z_{(1,\ell),(ns,nt)} + \sum_{k=1}^{ns} \left(\prod_{j=1}^{nt} V_{0,j}^{-1} \right) \left(\prod_{i=1}^{k} U_{i,0} \right) Z_{(k,1),(ns,nt)}$$

Now we know LDP for log(l.h.s) and can extract log Z from the r.h.s.





Fix N, let $1 \le k \le N$ and $n \ge 1$ vary.



Fix N, let
$$1 \le k \le N$$
 and $n \ge 1$ vary.

$$\Pi^1_{n,k} = \{ \text{ admissible paths } (1,1) \rightarrow (n,k) \}$$



Fix N, let $1 \le k \le N$ and $n \ge 1$ vary. $\Pi_{n,k}^{1} = \{ \text{ admissible paths } (1,1) \to (n,k) \}$ $z_{k,1}(n) = \sum_{\pi \in \Pi_{n,k}^{1}} wt(\pi) \text{ where}$

weight
$$wt(\pi) = \prod_{(i,j)\in\pi} Y_{i,j}$$



Fix N, let $1 \le k \le N$ and $n \ge 1$ vary. $\Pi_{n,k}^{1} = \{ \text{ admissible paths } (1,1) \to (n,k) \}$ $z_{k,1}(n) = \sum_{\pi \in \Pi_{n,k}^{1}} wt(\pi) \text{ where}$ weight $wt(\pi) = \prod_{(i,i) \in \pi} Y_{i,j}$

$$\Pi_{n,k}^{\ell} = \{ \ \ell \text{-tuples } \pi = (\pi_1, \dots, \pi_{\ell}) \text{ of disjoint}$$

paths $\pi_j : (1, j) \rightarrow (n, k - j + 1) \}$



Fix
$$N$$
, let $1 \le k \le N$ and $n \ge 1$ vary.
 $\Pi_{n,k}^1 = \{ \text{ admissible paths } (1,1) \to (n,k) \}$
 $z_{k,1}(n) = \sum_{\pi \in \Pi_{n,k}^1} wt(\pi) \text{ where}$
weight $wt(\pi) = \prod_{(i,j)\in\pi} Y_{i,j}$

$$\Pi_{n,k}^{\ell} = \{ \ \ell \text{-tuples } \pi = (\pi_1, \dots, \pi_{\ell}) \text{ of disjoint}$$

paths $\pi_j : (1,j) \to (n,k-j+1) \}$





Fix N, let
$$1 \le k \le N$$
 and $n \ge 1$ vary.

$$\Pi_{n,k}^{1} = \{ \text{ admissible paths } (1,1) \to (n,k) \}$$

$$z_{k,1}(n) = \sum_{\pi \in \Pi_{n,k}^{1}} wt(\pi) \text{ where}$$
weight $wt(\pi) = \prod_{(i,j)\in\pi} Y_{i,j}$

$$\Pi_{n,k}^{\ell} = \{ \ \ell \text{-tuples } \pi = (\pi_1, \dots, \pi_\ell) \text{ of disjoint}$$

$$\text{paths } \pi_j : (1,j) \to (n, k - j + 1) \}$$

$$\text{weight} \quad wt(\pi) = \prod_{(i,j) \in \pi} Y_{i,j}$$



Fix N, let
$$1 \le k \le N$$
 and $n \ge 1$ vary.
 $\Pi_{n,k}^1 = \{ \text{ admissible paths } (1,1) \to (n,k) \}$
 $z_{k,1}(n) = \sum_{\pi \in \Pi_{n,k}^1} wt(\pi) \text{ where}$
weight $wt(\pi) = \prod_{(i,j) \in \pi} Y_{i,j}$

$$\Pi_{n,k}^{\ell} = \{ \ \ell\text{-tuples } \pi = (\pi_1, \dots, \pi_{\ell}) \text{ of disjoint} \\ \text{paths } \pi_j : (1,j) \to (n,k-j+1) \ \}$$

weight $wt(\pi) = \prod_{(i,j)\in\pi} Y_{i,j}$
 $\tau_{k,\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi)$

$$au_{k,\ell}(n) = \sum_{\pi \in \Pi_{n,k}^\ell} wt(\pi) \qquad \quad ext{for } 1 \leq k \leq N, \ 1 \leq \ell \leq n \wedge k.$$

$$au_{k,\ell}(n) = \sum_{\pi \in \Pi_{n,k}^\ell} wt(\pi) \qquad \quad ext{for } 1 \leq k \leq N, \ 1 \leq \ell \leq n \wedge k.$$

Define array $z(n) = \{z_{k,\ell}(n): \ 1 \leq k \leq N, \ 1 \leq \ell \leq k \land n\}$ by

$$z_{k,1}(n) \cdots z_{k,\ell}(n) = \tau_{k\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi).$$

$$au_{k,\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi) \qquad \quad ext{for } 1 \leq k \leq N, \ 1 \leq \ell \leq n \wedge k.$$

Define array $z(n) = \{z_{k,\ell}(n): 1 \le k \le N, 1 \le \ell \le k \land n\}$ by

$$z_{k,1}(n) \cdots z_{k,\ell}(n) = \tau_{k\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi).$$

N = 4 array $z_{11}(n)$

$$\begin{array}{ccc} & z_{22}(n) & z_{21}(n) & polymer\\ & z_{33}(n) & z_{32}(n) & z_{31}(n) \\ & z_{44}(n) & z_{43}(n) & z_{42}(n) & z_{41}(n) \end{array}$$

$$au_{k,\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi) \qquad \quad ext{for } 1 \leq k \leq N, \ 1 \leq \ell \leq n \wedge k.$$

Define array $z(n) = \{z_{k,\ell}(n): 1 \le k \le N, 1 \le \ell \le k \land n\}$ by

$$z_{k,1}(n) \cdots z_{k,\ell}(n) = \tau_{k\ell}(n) = \sum_{\pi \in \Pi_{n,k}^{\ell}} wt(\pi).$$

N = 4 array z_1

 $z_{11}(n)$

$$z_{22}(n) \qquad z_{21}(n) \qquad polymer$$

$$z_{33}(n) \qquad z_{32}(n) \qquad z_{31}(n)$$

$$z_{44}(n) \qquad z_{43}(n) \qquad z_{42}(n) \qquad z_{41}(n)$$

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

Look at case $\ell = 1$:



RSk

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

Look at case $\ell = 1$:



Add a new column *n* in weight matrix.

RSł

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

Look at case $\ell = 1$:



Add a new column *n* in weight matrix.

$$z_{1,1}(n) = Y_{n,1} z_{1,1}(n-1)$$

RSł
The mapping

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

Look at case $\ell = 1$:



Add a new column *n* in weight matrix. $z_{1,1}(n) = Y_{n,1} z_{1,1}(n-1)$ For k = 2, ..., N $z_{k,1}(n) = Y_{n,k} (z_{k,1}(n-1) + z_{k-1,1}(n))$ The mapping

weight matrix
$$(Y_{i,j}) \mapsto \operatorname{array} z(n)$$

is Kirillov's tropical RSK correspondence (2001).

Next develop the Markovian evolution in terms of tropical or geometric row insertion.

Look at case $\ell = 1$:



Add a new column *n* in weight matrix. $z_{1,1}(n) = Y_{n,1} z_{1,1}(n-1)$ For k = 2, ..., N $z_{k,1}(n) = Y_{n,k} (z_{k,1}(n-1) + z_{k-1,1}(n))$

After this, transformed weights are passed on to diagonal $z_2 = (z_{22} \dots, z_{N2})$ and that diagonal is updated. And so on.

Let $1 \le \ell \le N$. Geometric row insertion

of the word $b = (b_\ell, \dots, b_N)$ into the word $\xi = (\xi_\ell, \dots, \xi_N)$

produces two new words

$$\xi' = (\xi'_\ell, \dots, \xi'_N) \quad \text{ and } b' = (b'_{\ell+1}, \dots, b'_N).$$

Notation and definition:

$$\xi \xrightarrow{b} \xi' \quad ext{where} \quad \left\{ egin{array}{c} \xi'_\ell = b_\ell \xi_\ell \ \xi'_k = b_k (\xi'_{k-1} + \xi_k) & \ell+1 \leq k \leq N \ b'_k = b_k rac{\xi_k \xi'_{k-1}}{\xi_{k-1} \xi'_k} & \ell+1 \leq k \leq N. \end{array}
ight.$$

Words have strictly positive real entries.

Let z be an array with diagonals z_1, \ldots, z_N , and $b \in (0, \infty)^N$ a word.

Let z be an array with diagonals z_1, \ldots, z_N , and $b \in (0, \infty)^N$ a word. **Geometric row insertion** of b into z produces a new array $z' = z \leftarrow b$ defined by iterating basic row insertion N times. Let z be an array with diagonals z_1, \ldots, z_N , and $b \in (0, \infty)^N$ a word. **Geometric row insertion** of b into z produces a new array $z' = z \leftarrow b$ defined by iterating basic row insertion N times.

Let $a_1 = b$, and then

$$\begin{array}{cccc} a_1 & & \\ z_1 & \stackrel{a_2}{\longrightarrow} & z'_1 \\ z_2 & \stackrel{a_2}{\longrightarrow} & z'_2 \\ a_3 & & \\ \vdots & \\ z_N & \stackrel{a_N}{\longrightarrow} & z'_N \end{array}$$

Let z be an array with diagonals z_1, \ldots, z_N , and $b \in (0, \infty)^N$ a word. **Geometric row insertion** of b into z produces a new array $z' = z \leftarrow b$ defined by iterating basic row insertion N times.

Let $a_1 = b$, and then

$$\begin{array}{c} a_{1} \\ z_{1} \xrightarrow{a_{2}} z'_{1} \\ z_{2} \xrightarrow{a_{2}} z'_{2} \\ a_{3} \\ \vdots \\ z_{N} \xrightarrow{a_{N}} z'_{N} \end{array}$$

The process exhausts the input, a_{N+1} is empty.

Let z be an array with diagonals z_1, \ldots, z_N , and $b \in (0, \infty)^N$ a word. **Geometric row insertion** of b into z produces a new array $z' = z \leftarrow b$ defined by iterating basic row insertion N times.

Let $a_1 = b$, and then

The process exhausts the input, a_{N+1} is empty.

Diagonals z'_1, \ldots, z'_N make up the new array z'.



Evolution of the array z(n) over time n = 0, 1, 2, ...



Evolution of the array z(n) over time n = 0, 1, 2, ...Initial state z(0) is on the left edge in terms of diagonals $z_1(0), ..., z_N(0)$.



Evolution of the array z(n) over time n = 0, 1, 2, ...Initial state z(0) is on the left edge in terms of diagonals $z_1(0), ..., z_N(0)$. Time n input from weight matrix: $a_1(n) = Y^{[n]} = (Y_{n,1}, ..., Y_{n,N})$. At each time, geometric row insertion is iterated N times to update each diagonal. The previous process is for evolving the full array. We also need the variant for starting with an empty array $z(0) = \emptyset$.

The previous process is for evolving the full array. We also need the variant for starting with an empty array $z(0) = \emptyset$.

$$e_{1}^{(N)} \xrightarrow{a_{1}(1)} a_{1}(2) \qquad a_{1}(3)$$

$$e_{1}^{(N)} \xrightarrow{\downarrow} z_{1}(1) \xrightarrow{\downarrow} z_{1}(2) \xrightarrow{\downarrow} z_{1}(3) \cdots$$

$$a_{2}(2) \qquad a_{2}(3)$$

$$e_{1}^{(N-1)} \xrightarrow{\downarrow} z_{2}(2) \xrightarrow{\downarrow} z_{2}(3) \cdots$$

$$a_{3}(3)$$

$$e_{1}^{(N-2)} \xrightarrow{\downarrow} z_{3}(3) \cdots$$

Resulting array: $z(n) = \emptyset \leftarrow Y^{[1]} \leftarrow Y^{[2]} \leftarrow \cdots \leftarrow Y^{[n]}$.

(Noumi and Yamada (2004), proof by a matrix technique.)

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

Assumption. Weights $\{Y_{n,j}\}$ are independent with marginals $Y_{n,j} \sim \Gamma^{-1}(\hat{\theta}_n + \theta_j)$ where $\{\hat{\theta}_n, \theta_j\}$ are fixed real parameters such that each $\hat{\theta}_n + \theta_j > 0$.

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

Assumption. Weights $\{Y_{n,j}\}$ are independent with marginals $Y_{n,j} \sim \Gamma^{-1}(\hat{\theta}_n + \theta_j)$ where $\{\hat{\theta}_n, \theta_j\}$ are fixed real parameters such that each $\hat{\theta}_n + \theta_j > 0$.

Can we say anything about partition function $z_{N,1}(n)$?

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

Assumption. Weights $\{Y_{n,j}\}$ are independent with marginals $Y_{n,j} \sim \Gamma^{-1}(\hat{\theta}_n + \theta_j)$ where $\{\hat{\theta}_n, \theta_j\}$ are fixed real parameters such that each $\hat{\theta}_n + \theta_j > 0$.

Can we say anything about partition function $z_{N,1}(n)$?

Markov kernel Π_n for time *n* transition $z(n-1) \rightarrow z(n)$ is complicated.

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

Assumption. Weights $\{Y_{n,j}\}$ are independent with marginals $Y_{n,j} \sim \Gamma^{-1}(\hat{\theta}_n + \theta_j)$ where $\{\hat{\theta}_n, \theta_j\}$ are fixed real parameters such that each $\hat{\theta}_n + \theta_j > 0$.

Can we say anything about partition function $z_{N,1}(n)$?

Markov kernel Π_n for time *n* transition $z(n-1) \rightarrow z(n)$ is complicated.

Bottom row $y(n) = (z_{N,1}(n), z_{N,2}(n), \dots, z_{N,N}(n))$ of the array turns out to be a more tractable Markov chain.

(Noumi and Yamada (2004), proof by a matrix technique.)

Now let us make the input random.

Assumption. Weights $\{Y_{n,j}\}$ are independent with marginals $Y_{n,j} \sim \Gamma^{-1}(\hat{\theta}_n + \theta_j)$ where $\{\hat{\theta}_n, \theta_j\}$ are fixed real parameters such that each $\hat{\theta}_n + \theta_j > 0$.

Can we say anything about partition function $z_{N,1}(n)$?

Markov kernel Π_n for time *n* transition $z(n-1) \rightarrow z(n)$ is complicated.

Bottom row $y(n) = (z_{N,1}(n), z_{N,2}(n), \dots, z_{N,N}(n))$ of the array turns out to be a more tractable Markov chain.

Theory of Markov functions shows this.

 \exists Markov kernel Π for z(n) on space T.



Т

Т

П

 \exists Markov kernel Π for z(n) on space T.

 $\exists map \phi : T \rightarrow Y.$



 \exists Markov kernel Π for z(n) on space T.

 $\exists map \phi : T \rightarrow Y.$

When is $y(n) = \phi(z(n))$ Markov with kernel \overline{P} ?



 \exists Markov kernel Π for z(n) on space T.

 $\exists \mathsf{map} \phi: T \to Y.$

When is $y(n) = \phi(z(n))$ Markov with kernel \overline{P} ?



Sufficient condition. Suppose \exists (positive but not necessary stochastic) kernels $P: Y \rightarrow Y$ and $K: Y \rightarrow T$ such that

 $K(y, \phi^{-1}(y)) = 1$ and $K \circ \Pi = P \circ K$

 \exists Markov kernel Π for z(n) on space T.

 $\exists \mathsf{map} \phi : T \to Y.$

When is $y(n) = \phi(z(n))$ Markov with kernel \overline{P} ?



Sufficient condition. Suppose \exists (positive but not necessary stochastic) kernels $P: Y \rightarrow Y$ and $K: Y \rightarrow T$ such that

$$K(y, \phi^{-1}(y)) = 1$$
 and $K \circ \Pi = P \circ K$

Set w(y) = K(y, T). Intertwining gives Pw = w. Define stochastic kernels

$$ar{K}(y,dz) = rac{1}{w(y)}K(y,dz) \quad ext{and} \quad ar{P}(y,d ilde{y}) = rac{w(ilde{y})}{w(y)}P(y,d ilde{y})$$

Markov functions idea, continued

Then $\bar{K} \circ \Pi = \bar{P} \circ \bar{K}$



Markov functions idea, continued



If z(n) starts with distribution $\overline{K}(y, dz)$, then y(n) is Markov in its own filtration with transition \overline{P} and initial state y(0) = y.

Markov functions idea, continued



If z(n) starts with distribution $\overline{K}(y, dz)$, then y(n) is Markov in its own filtration with transition \overline{P} and initial state y(0) = y.

Furthermore

$$E[f(z(n)) | y(0), ..., y(n-1), y(n) = y] = \bar{K}f(y)$$

(Rogers and Pitman, 1981)

Spaces: \mathbb{T}_N = space of arrays of size N,

 $\mathbb{Y}_N = (0, \infty)^N =$ space of positive *N*-vectors.

Spaces:
$$\mathbb{T}_N =$$
 space of arrays of size N ,
 $\mathbb{Y}_N = (0, \infty)^N =$ space of positive N -vectors.

Define a (substochastic) kernel P_n on \mathbb{Y}_N by

$$P_n(y, d\tilde{y}) = \prod_{i=1}^{N-1} \exp\left\{-\frac{\tilde{y}_{i+1}}{y_i}\right\} \prod_{j=1}^N \left(\Gamma(\gamma_{n,j})^{-1} \left(\frac{y_j}{\tilde{y}_j}\right)^{\gamma_{n,j}} \exp\left\{-\frac{y_j}{\tilde{y}_j}\right\} \frac{d\tilde{y}_j}{\tilde{y}_j}\right)$$

Spaces:
$$\mathbb{T}_N$$
 = space of arrays of size N ,
 $\mathbb{Y}_N = (0, \infty)^N$ = space of positive N -vectors.

Define a (substochastic) kernel P_n on \mathbb{Y}_N by

$$P_n(y, d\tilde{y}) = \prod_{i=1}^{N-1} \exp\left\{-\frac{\tilde{y}_{i+1}}{y_i}\right\} \prod_{j=1}^N \left(\Gamma(\gamma_{n,j})^{-1} \left(\frac{y_j}{\tilde{y}_j}\right)^{\gamma_{n,j}} \exp\left\{-\frac{y_j}{\tilde{y}_j}\right\} \frac{d\tilde{y}_j}{\tilde{y}_j}\right)$$

and intertwining kernel $\mathcal{K}:\mathbb{Y}_N\to\mathbb{T}_N$ by

$$\begin{split} \mathcal{K}(y,dz) &= \prod_{1 \leq \ell \leq k < N} \left(\frac{z_{k,\ell}}{z_{k+1,\ell}} \right)^{\theta_{k+1} - \theta_{\ell}} \\ &\times \exp\left(- \frac{z_{k,\ell}}{z_{k+1,\ell}} - \frac{z_{k+1,\ell+1}}{z_{k,\ell}} \right) \frac{dz_{k,\ell}}{z_{k,\ell}} \prod_{\ell=1}^{N} \delta_{y_{\ell}}(dz_{N,\ell}) \end{split}$$

Then $K \circ \prod_n = P_n \circ K$.

Then $K \circ \prod_n = P_n \circ K$. Bottom row y(n) is a MC with kernel

$$ar{P}_n(y,d ilde{y}) = rac{w(ilde{y})}{w(y)} P_n(y,d ilde{y})$$

where $w(y) = K(y, \mathbb{T}_N)$.

Then $K \circ \Pi_n = P_n \circ K$. Bottom row y(n) is a MC with kernel

$$\bar{P}_n(y,d\tilde{y}) = \frac{w(\tilde{y})}{w(y)} P_n(y,d\tilde{y})$$

where $w(y) = K(y, \mathbb{T}_N)$.

Conditional distribution of array, given evolution of bottom row:

$$E[f(z(n)) | y(0), ..., y(n-1), y(n) = y] = \bar{K}f(y)$$

Then $K \circ \Pi_n = P_n \circ K$. Bottom row y(n) is a MC with kernel

$$\bar{P}_n(y,d\tilde{y}) = \frac{w(\tilde{y})}{w(y)} P_n(y,d\tilde{y})$$

where $w(y) = K(y, \mathbb{T}_N)$.

Conditional distribution of array, given evolution of bottom row:

$$E[f(z(n)) | y(0), ..., y(n-1), y(n) = y] = \bar{K}f(y)$$

Now a closer look at the eigenfunctions we have found. All the previous makes sense also for complex parameters. This is beneficial because then we can use known special functions to diagonalize the transition kernel.
Whittaker functions

 $GL(N, \mathbb{R})$ -Whittaker function for $y \in \mathbb{Y}_N$, with $\lambda \in \mathbb{C}^N$

$$\Psi_{\lambda}(y) = \prod_{i=1}^{N} y_i^{-\lambda_i} \int_{\mathbb{T}_N} K_{\lambda}(y, dz)$$

where K_{λ} is the previous intertwining kernel with θ replaced by λ .

Whittaker functions

 $GL(N, \mathbb{R})$ -Whittaker function for $y \in \mathbb{Y}_N$, with $\lambda \in \mathbb{C}^N$

$$\Psi_{\lambda}(y) = \prod_{i=1}^{N} y_i^{-\lambda_i} \int_{\mathbb{T}_N} K_{\lambda}(y, dz)$$

where K_{λ} is the previous intertwining kernel with θ replaced by λ .

Intertwining works also with complex parameters and gives

$$\int_{(0,\infty)^N} \frac{\Psi_{\theta+\lambda}(\tilde{y})}{\Psi_{\theta}(\tilde{y})} \, \bar{P}_n(y,d\tilde{y}) = \left(\prod_{i=1}^N \frac{\Gamma(\gamma_{n,i}+\lambda_i)}{\Gamma(\gamma_{n,i})}\right) \frac{\Psi_{\theta+\lambda}(y)}{\Psi_{\theta}(y)}$$

Our goal is an expression for the distribution of $z_{N,1}(n)$, the polymer partition function. Getting there involves two steps.

Our goal is an expression for the distribution of $z_{N,1}(n)$, the polymer partition function. Getting there involves two steps.

Step 1. Row insertion reveals that if we start with bottom row

$$y^M = (e^{M(i-(k+1)/2)})_{1 \le i \le N},$$

let initial array have distribution $\overline{K}(y^N, dz)$, and let $M \to \infty$, the distribution of the array z(n) converges to the array started from the empty array $z(0) = \emptyset$.

So this limit recovers the distribution of the partition function from the path construction.

Outline Introduction Burke property Tropical RSK

Step 2. Invert the eigenfunction relation to write an expression for the distribution of bottom row y(n) when started from $y \in (0, \infty)^N$. Involves analytical properties of Whittaker functions (analogous to Fourier analysis). Take $y = y^M$ and let $M \to \infty$.

The result is a formula for the Laplace transform of the partition function:

$$\mathbb{E}(e^{-s \, z_{N,1}(n)}) = \int_{\iota \mathbb{R}^N} s^{\sum_{i=1}^N (\theta_i - \lambda_i)} \prod_{1 \le i, j \le N} \Gamma(\lambda_i - \theta_j) \\ \times \prod_{m=1}^n \prod_{i=1}^N \frac{\Gamma(\lambda_i + \hat{\theta}_m)}{\Gamma(\theta_i + \hat{\theta}_m)} \, s_N(\lambda) \, d\lambda$$

where the Sklyanin measure is given by

$$s_N(\lambda) = rac{1}{(2\pi\iota)^N N!} \prod_{j \neq k} \Gamma(\lambda_j - \lambda_k)^{-1}$$

A future goal: asymptotics for distribution of $\log z_{N,1}(n)$.

Proof of Burke property

Induction on ${\mathcal I}$ by flipping a growth corner:

$$\bigvee \underbrace{\overset{\mathsf{U}'}{\bullet}}_{\mathsf{U}} \qquad \underbrace{\overset{\mathsf{U}'}{\times}}_{\mathsf{V}} \qquad \underbrace{U' = Y(1 + U/V)}_{\mathsf{X} = (U^{-1} + V^{-1})^{-1}}$$

Proof of Burke property

Induction on \mathcal{I} by flipping a growth corner:

$$V \stackrel{V'}{\frown} V \stackrel{U'}{\frown} V' \qquad U' = Y(1 + U/V) \quad V' = Y(1 + V/U) \\ X \stackrel{\bullet}{\bullet} V' \qquad X = (U^{-1} + V^{-1})^{-1}$$

Lemma. Given that (U, V, Y) are independent positive r.v.'s, $(U', V', X) \stackrel{d}{=} (U, V, Y)$ iff (U, V, Y) have the gamma distr's.

Proof. "if" part by computation, "only if" part from a characterization of gamma due to Lukacs (1955). \Box

Proof of Burke property

Induction on \mathcal{I} by flipping a growth corner:

$$V \stackrel{\bullet}{\underset{U}{\bullet}} \stackrel{Y}{\underset{V}{\bullet}} V' \qquad \begin{array}{c} U' = Y(1 + U/V) \\ X \stackrel{\bullet}{\bullet} V' \\ X = (U^{-1} + V^{-1})^{-1} \end{array}$$

Lemma. Given that (U, V, Y) are independent positive r.v.'s, $(U', V', X) \stackrel{d}{=} (U, V, Y)$ iff (U, V, Y) have the gamma distr's.

Proof. "if" part by computation, "only if" part from a characterization of gamma due to Lukacs (1955). \Box

This gives all (z_k) with finite \mathcal{I} . General case follows.

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau T.

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau T.

Rows weakly, columns strictly increasing.

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau *T*. Rows weakly, columns strictly increasing.

 $\lambda_i =$ length of *i*th row

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau T. Rows weakly, columns strictly increasing. $\lambda_i =$ length of *i*th row

Shape $sh(T) = \lambda = (\lambda_1, \dots, \lambda_4) = (3, 3, 2, 1).$

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau *T*. Rows weakly, columns strictly increasing. $\lambda_i = \text{length of } i\text{th row}$ Shape sh(*T*) = $\lambda = (\lambda_1, \dots, \lambda_4) = (3, 3, 2, 1)$.

 $x_{i,j}$ = number of entries $\leq i$ in row j of the P-tableau.

Classic RSK maps an $n \times N$ weight matrix $Y = (Y_{i,j})$ with nonnegative integer entries bijectively to a pair (P, Q) of Young tableaux with common shape.



Example of (semistandard) Young tableau *T*. Rows weakly, columns strictly increasing. $\lambda_i = \text{length of } i\text{th row}$ Shape sh(*T*) = $\lambda = (\lambda_1, \dots, \lambda_4) = (3, 3, 2, 1)$.

 $x_{i,j}$ = number of entries $\leq i$ in row j of the P-tableau.

$$x_{i,j} \leq x_{i-1,j-1} \leq x_{i,j-1}$$

$\{x_{i,j}\}$ can be arranged

in a Gelfand-Tsetlin pattern:

$\{x_{i,j}\}$	can	be	arranged
---------------	-----	----	----------

in a Gelfand-Tsetlin pattern:

 $\begin{array}{c} x_{1,1} \\ x_{2,2} \\ \dots \\ \end{array} \begin{array}{c} x_{2,1} \\ \dots \end{array}$

 $X_{N,N}$... $X_{N,2}$ $X_{N,1}$

$\{x_{i,j}\}$	can	be	arranged
---------------	-----	----	----------

in a Gelfand-Tsetlin pattern:

Bottom row = shape.

 $\{x_{i,j}\}$ can be arranged $x_{1,1}$ in a Gelfand-Tsetlin pattern: $x_{2,2}$ $x_{2,1}$ Bottom row = shape. $x_{N,N}$ \dots $x_{N,2}$

$$x_{k,1} + \cdots + x_{k,\ell} = \max_{\pi_1, \dots, \pi_\ell \text{ disjoint }} \sum_{(i,j) \in \pi_1 \cup \cdots \cup \pi_\ell} Y_{i,j}$$

where π_m are up-right lattice paths in the weight matrix.

 $\{x_{i,j}\}$ can be arranged $x_{1,1}$ in a Gelfand-Tsetlin pattern: $x_{2,2}$ $x_{2,1}$ Bottom row = shape. $x_{N,N}$ \dots $x_{N,2}$

$$x_{k,1} + \cdots + x_{k,\ell} = \max_{\pi_1, \dots, \pi_\ell \text{ disjoint }} \sum_{(i,j) \in \pi_1 \cup \cdots \cup \pi_\ell} Y_{i,j}$$

where π_m are up-right lattice paths in the weight matrix.

Compare this with tropical formula

$$z_{k,1}(n) \cdots z_{k,\ell}(n) = \sum_{\pi \in \prod_{n,k}^{\ell}} wt(\pi).$$

 $\{x_{i,j}\}$ can be arranged $x_{1,1}$ in a Gelfand-Tsetlin pattern: $x_{2,2}$ $x_{2,1}$ Bottom row = shape. $x_{N,N}$ \dots $x_{N,2}$

$$x_{k,1} + \cdots + x_{k,\ell} = \max_{\pi_1, \dots, \pi_\ell \text{ disjoint }} \sum_{(i,j) \in \pi_1 \cup \cdots \cup \pi_\ell} Y_{i,j}$$

where π_m are up-right lattice paths in the weight matrix.

Compare this with tropical formula

$$z_{k,1}(n) \cdots z_{k,\ell}(n) = \sum_{\pi \in \prod_{n,k}^{\ell}} wt(\pi).$$

Difference is $(+, \cdot)$ vs. (max, +).



Consequences of Burke property: variance identity



Exit point of path from x-axis

$$\xi_x = \max\{k \ge 0 : x_i = (i, 0) \text{ for } 0 \le i \le k\}$$

Consequences of Burke property: variance identity



Exit point of path from x-axis $\xi_x = \max\{k \ge 0 : x_i = (i, 0) \text{ for } 0 \le i \le k\}$

For $\theta, x > 0$ define positive function

$$L(\theta, x) = \int_0^x (\Psi_0(\theta) - \log y) x^{-\theta} y^{\theta-1} e^{x-y} \, dy$$

Consequences of Burke property: variance identity



Exit point of path from x-axis $\xi_x = \max\{k \ge 0 : x_i = (i, 0) \text{ for } 0 \le i \le k\}$

For $\theta, x > 0$ define positive function

$$L(\theta, x) = \int_0^x (\Psi_0(\theta) - \log y) x^{-\theta} y^{\theta-1} e^{x-y} dy$$

Theorem. For the stationary case,

$$\operatorname{Var}\left[\log Z_{m,n}\right] = n\Psi_1(\mu - \theta) - m\Psi_1(\theta) + 2 E_{m,n}\left[\sum_{i=1}^{\xi_x} L(\theta, Y_{i,0}^{-1})\right]$$

Variance identity leads to fluctuation bounds for $\log Z$

With $0 < \theta < \mu$ fixed and $N \nearrow \infty$ assume

$$|\mathit{m}-\mathit{N}\Psi_1(\mu- heta)|\leq \mathit{CN}^{2/3}$$
 and $|\mathit{n}-\mathit{N}\Psi_1(heta)|\leq \mathit{CN}^{2/3}$ (1)

Variance identity leads to fluctuation bounds for $\log Z$

With 0 < θ < μ fixed and N $\nearrow \infty$ assume

$$|m - N\Psi_1(\mu - heta)| \leq CN^{2/3}$$
 and $|n - N\Psi_1(heta)| \leq CN^{2/3}$ (1)

Theorem: Variance bounds along characteristic

For (m, n) as in (1), $C_1 N^{2/3} \leq Var(\log Z_{m,n}) \leq C_2 N^{2/3}$

Variance identity leads to fluctuation bounds for $\log Z$

With 0 < θ < μ fixed and N $\nearrow \infty$ assume

 $|m - N\Psi_1(\mu - \theta)| \leq CN^{2/3}$ and $|n - N\Psi_1(\theta)| \leq CN^{2/3}$ (1)

Theorem: Variance bounds along characteristic

For (m, n) as in (1), $C_1 N^{2/3} \leq \mathbb{V}ar(\log Z_{m,n}) \leq C_2 N^{2/3}$.

Theorem: Off-characteristic CLT

Suppose $n = \Psi_1(\theta)N$ and $m = \Psi_1(\mu - \theta)N + \gamma N^{\alpha}$ with $\gamma > 0$, $\alpha > 2/3$. Then

$$N^{-lpha/2} \Big\{ \log Z_{m,n} - \mathbb{E} \big(\log Z_{m,n} \big) \Big\} \ \Rightarrow \ \mathcal{N} \big(0, \gamma \Psi_1(heta) \big)$$