

Practical unbiased Monte Carlo for Uncertainty Quantification


Sergios Agapiou


Department of Statistics, University of Warwick

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Enabling Quantification of
EQUIP
Uncertainty for Inverse Problems

<http://www.sergiosagapiou.com/>

-  S. Agapiou, G. O. Roberts and S. J. Vollmer, *Unbiased Monte Carlo: posterior estimation for intractable/infinite dimensional models*, <http://arxiv.org/abs/1411.7713>

-  C. H. Rhee, *Unbiased estimation with biased samples*, PhD thesis, Stanford University, 2013 (supervisor P. W. Glynn).

Outline

- 1 Introduction - General theory
- 2 UQ example
- 3 Removing specific sources of bias
- 4 Performance/Optimization
- 5 Conclusions

Problem

Want to estimate expectations of functions f wrt an intractable measure μ ,
$$\mathbb{E}_\mu[f] := \mathbb{E}_\mu[f(\cdot)].$$

- Would like to use Monte Carlo estimator: for $X^{(m)} \stackrel{iid}{\sim} \mu$ let

$$Y_M := \frac{1}{M} \sum_{m=1}^M f(X^{(m)}).$$

For all M

$$\mathbb{E}[Y_M] = \mathbb{E}_\mu[f] \quad (Y_M \text{ unbiased})$$

and

$$Y_M \xrightarrow{M} \mathbb{E}_\mu[f], \text{ almost surely} \quad (Y_m \text{ consistent})$$

- Intractability of μ forces the use of approximations μ_i introducing **bias**.

Debiasing idea - John von Neumann, Stanislaw Ulam

- We study **unbiased** estimation of $\mathbb{E}_\mu[f]$ using biased samples.
- Assume $\mathbb{E}_{\mu_i}[f] \xrightarrow{i} \mathbb{E}_\mu[f]$.
- Let $X_i \sim \mu_i$ and define $\Delta_i := f(X_i) - f(X_{i-1})$.
- **If** Fubini applies

$$\mathbb{E}_\mu[f] = \sum_{i=1}^{\infty} (\mathbb{E}_{\mu_i}[f] - \mathbb{E}_{\mu_{i-1}}[f]) = \sum_{i=1}^{\infty} \mathbb{E}\Delta_i \stackrel{?}{=} \mathbb{E} \sum_{i=1}^{\infty} \Delta_i.$$

- $\sum_{i=1}^{\infty} \Delta_i$ is **unbiased** but requires **infinite computing time**.

Debiasing idea - John von Neumann, Stanislaw Ulam

$$Z := \sum_{i=0}^N \frac{\Delta_i}{\mathbb{P}(N \geq i)},$$

N integer-valued r.v. independent of Δ_i , s.t. $\mathbb{P}(N \geq i) > 0, \forall i$.

- If Fubini applies then Z unbiased

$$\mathbb{E}[Z] = \mathbb{E} \left[\sum_{i=0}^{\infty} \frac{\mathbb{1}_{\{N \geq i\}} \Delta_i}{\mathbb{P}(N \geq i)} \right] \stackrel{?}{=} \sum_{i=0}^{\infty} \frac{\mathbb{E}[\mathbb{1}_{\{N \geq i\}} \Delta_i]}{\mathbb{P}(N \geq i)} = \sum_{i=0}^{\infty} \mathbb{E} \Delta_i = \mathbb{E}_{\mu}[f].$$

- To be practical, Z needs to have finite variance and finite expected computing time.

Unbiasing theory of Glynn and Rhee

Proposition (GR13)

Assume

$$\sum_{i \leq \ell} \frac{\|\Delta_i\|_2 \|\Delta_\ell\|_2}{\mathbb{P}(N \geq i)} < \infty.$$

Then $Z := \sum_{i=0}^N \frac{\Delta_i}{\mathbb{P}(N \geq i)}$ is an unbiased estimator for $\mathbb{E}_\mu[f]$ with finite variance.

Can use $\tilde{\Delta}_i$ copy of Δ_i s.t. $\{\tilde{\Delta}_i\}$ mutually independent.

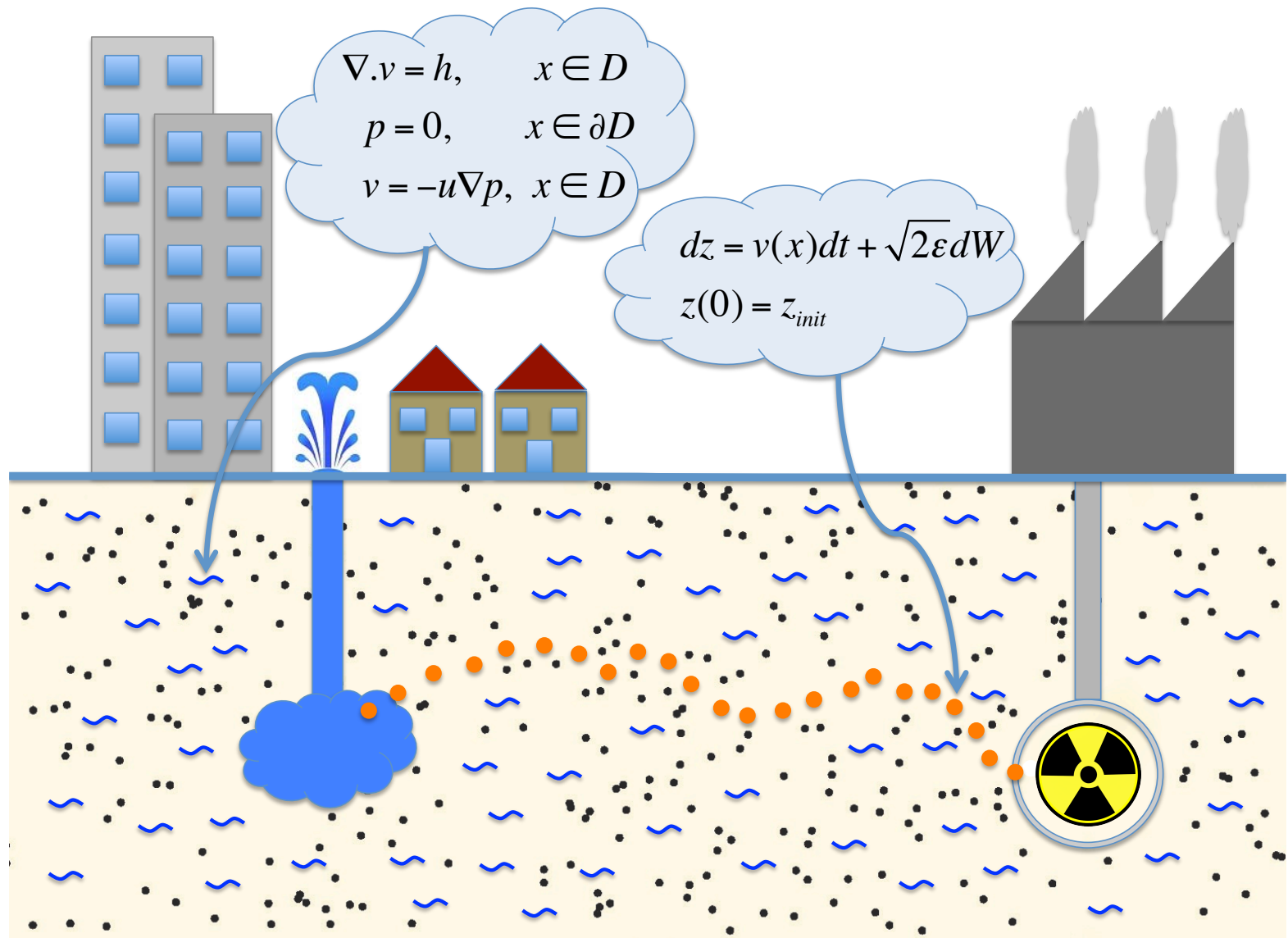
- t_i expected cost of generating Δ_i . Expected computing time of Z

$$\mathbb{E}(\tau) = \mathbb{E} \sum_{i=0}^N t_i = \sum_{i=0}^{\infty} t_i \mathbb{P}(N \geq i).$$

- To be possible to choose $\mathbb{P}(N \geq i)$ s.t. Z practical, suffices to generate Δ_i 's with correct expectation s.t. $\|\Delta_i\|_2^2$ decays sufficiently faster than t_i blows-up.

Example - Contamination scenario

- u permeability field
- p pressure
- v Darcy velocity
- $p = G(u)$



Quantity of interest: $f(u) = \mathbb{E}[\inf_{t \geq 0} \{|z(t)| > R\}]$

Example - UQ in contamination scenario

Permeability field u **unknown**, have prior information $u \sim \mu_0$.

- **Vanilla-UQ**: probe $\mu_0 \circ f^{-1}$, e.g. estimate $\mathbb{E}_{\mu_0}[f(u)]$.
- Have noisy indirect measurements of pressure: data model in \mathbb{R}^J

$$y = \mathcal{G}(u) + \eta, \quad \eta \sim N(0, \Gamma).$$

Formulate Bayesian inverse problem (see [DS13](#)), μ^y posterior on $u|y$

$$\frac{d\mu^y}{d\mu_0}(u; y) \propto \exp\left(-\frac{1}{2}\|y - \mathcal{G}(u)\|^2\right).$$

BIP-UQ: probe $\mu^y \circ f^{-1}$, e.g. estimate $\mathbb{E}_{\mu^y}[f(u)]$.

- μ_0 is ∞ -dim, needs to be approximated by $\mu_{0,i}$ in \mathbb{R}^i introducing **discretization bias** ([ARV14](#)).
- cannot sample μ^y directly, construct Markov chain targeting μ^y , use finite-time distributions $\mu^{y,k}$ **burn-in time issues** ([GR13](#), [ARV14](#)).
- to implement in computer construct Markov chain targeting approximation μ_i^y in \mathbb{R}^i , use finite-time distributions $\mu_i^{y,k}$ introducing **discretization bias and burn-in time issues** ([ARV14](#)).

Removing discretization bias

- $\mathcal{X} = L^2[0, 1]$, $\{\varphi_\ell\}$ complete orthonormal basis.
- μ Gaussian measure in \mathcal{X} given via the **Karhunen-Loeve** expansion:

$$\mu = \mathcal{L} \left(\sum_{\ell=1}^{\infty} l^{-a} \xi_\ell \varphi_\ell \right), \quad \xi_\ell \stackrel{iid}{\sim} N(0, 1), \quad a > \frac{1}{2}.$$

- To estimate $\mathbb{E}_\mu[f]$, need to truncate introducing **discretization bias** in MC estimators.
(Vanilla-UQ example)

Aim: unbiasedly estimate $\mathbb{E}_\mu[f]$ in finite time.

- Approximations $\mu_i = \mathcal{L} \left(\sum_{\ell=1}^{j_i} l^{-a} \xi_\ell \varphi_\ell \right)$, $\{j_i\}$ increasing.
- $\Delta_i = f(u_i) - f(u_{i-1})$, $u_i \sim \mu_i$.

Removing discretization bias

Theorem 1 (ARV14)

Assume $a > 1$ and f Lipschitz. Then \exists choices j_i and $\mathbb{P}(N \geq i)$, s.t. $Z = \sum_{i=1}^N \frac{\Delta_i}{\mathbb{P}(N \geq i)}$ is unbiased estimator of $\mathbb{E}_\mu[f]$ with finite variance and finite expected computing time.

Proof.

- Consider $j_i = 2^i$. Use Proposition.
- Cost of Δ_i , $t_i = \mathcal{O}(j_i) = \mathcal{O}(2^i)$ (# $N(0, 1)$ draws).
- Bound

$$\|\Delta_i\|_2^2 = \mathbb{E}(|f(u_i) - f(u_{i-1})|^2) \leq \|f'\|_\infty^2 \mathbb{E}(\|u_i - u_{i-1}\|^2) = \mathcal{O}(2^{i(1-2a)}).$$

- $\|\Delta_i\|_2^2$ decays sufficiently faster than t_i blows-up.
- Can choose $\mathbb{P}(N \geq i)$ s.t. $\mathbb{E}(\tau), \text{Var}(Z) < \infty$.



Removing burn-in time bias

- \mathcal{X} general state space, d distance in \mathcal{X} , f d -Lipschitz.
- Measure μ intractable, cannot be sampled directly but can construct $\mathbb{X} = (X_n)_{n \in \mathbb{N}}$ Markov chain with stationary distribution μ .
- $\{a_i\}$ increasing sequence of positive integers.
- To estimate $\mathbb{E}_\mu[f]$, use finite-time distributions $\mu_i = \mathcal{L}(X_{a_i})$ introducing **burn-in issues**.

Aim: unbiasedly estimate $\mathbb{E}_\mu[f]$ in finite time.

Removing burn-in time bias

- Weak convergence of μ_j not enough to get convergence of Δ_j .
- **Contracting coupling assumption**: we can simultaneously generate chains started at different states s.t. they come together in d geometrically quickly.
- Use **top level** chain \mathcal{T}^i running for a_i steps and **bottom level** chain \mathcal{B}^i running for a_{i-1} steps, coupled as follows:

$$\begin{array}{cccccc}
 x_0 = & \mathcal{B}_{-a_{i-1}}^i & \cdots & \mathcal{B}_{-a_0}^i & \cdots & \mathcal{B}_0^i \\
 & | & & | & & | \\
 x_0 = & \mathcal{T}_{-a_i}^i & \cdots & \mathcal{T}_{-a_{i-1}}^i & \cdots & \mathcal{T}_0^i
 \end{array}
 \} \Delta_i = f(\mathcal{T}_0^i) - f(\mathcal{B}_0^i)$$

Removing burn-in time bias

Theorem 2 (ARV14)

\exists choices a_i and $\mathbb{P}(N \geq i)$, s.t. $Z = \sum_{i=1}^N \frac{\Delta_i}{\mathbb{P}(N \geq i)}$ is unbiased estimator of $\mathbb{E}_\mu[f]$ with finite variance and finite expected computing time.

Proof.

- Use Proposition.
- Using assumptions, can show $\|\Delta_i\|_2^2 \leq \|f'\|_\infty^2 \mathbb{E}d^2(\mathcal{T}_0^i, \mathcal{B}_0^i) \leq cr^{a_i}$.
- Cost of Δ_i , $t_i = \mathcal{O}(a_i)$ (# steps).
- $\|\Delta_i\|_2^2$ decays sufficiently faster than t_i blows-up.
- Can choose $\mathbb{P}(N \geq i)$ s.t. $\mathbb{E}(\tau), \text{Var}(Z) < \infty$.



UE for BIP-UQ in function space

- Combining can perform UE of $\mathbb{E}_\mu[f]$ for μ both ∞ -dim and only accessible in the limit of a Markov chain (BIP-UQ example).
- Approximation using finite-time distributions and discretizing space: **top chain** \mathcal{T}^i more steps **and** higher discretization level than **bottom chain** \mathcal{B}^i .

$$\begin{array}{rcc}
 j_{i-1} : & & x_0 = \mathcal{B}_{-a_{i-1}}^i \cdots \mathcal{B}_{-a_0}^i \cdots \mathcal{B}_0^i \\
 & & \quad \quad \quad | \quad \quad | \quad \quad | \quad \quad | \quad \quad | \\
 j_i : & x_0 = & \mathcal{T}_{-a_i}^i \cdots \mathcal{T}_{-a_{i-1}}^i \cdots \cdots \mathcal{T}_0^i
 \end{array}
 \left. \vphantom{\begin{array}{rcc} j_{i-1} : & & x_0 = \mathcal{B}_{-a_{i-1}}^i \cdots \mathcal{B}_{-a_0}^i \cdots \mathcal{B}_0^i \\ j_i : & x_0 = & \mathcal{T}_{-a_i}^i \cdots \mathcal{T}_{-a_{i-1}}^i \cdots \cdots \mathcal{T}_0^i \end{array}} \right\} \Delta_i = f(\mathcal{T}_0^i) - f(\mathcal{B}_0^i)$$

- In [ARV14](#), achieve this:
 - in non-linear BIPs (e.g. groundwater flow example) with uniform priors, using [independence sampler](#);
 - for targets μ which have Lipschitz log-density wrt Gaussian, using [pCN algorithm](#).
(MH with proposal $X_{k+1} = \lambda X_k + \sqrt{1 - \lambda^2} \xi$)

Comparison of Unbiased Estimator vs Ergodic Average

- 1d Gaussian autoregression

$$X_{n+1} = \rho X_n + \sqrt{1 - \rho^2} \xi_{n+1},$$

$\rho \in (0, 1)$, ξ_n i.i.d. $N(0, 1)$.

- Ergodic with invariant distribution $\mu = N(0, 1)$. Estimate $\mathbb{E}_\mu[\text{Id}] = 0$.
- Compare MSE-work product of Monte Carlo estimator based on UE vs EA.
- For EA

$$\lim_{n \rightarrow \infty} \text{MSE-work} = \frac{1 + \rho}{1 - \rho} T_{\text{step}}.$$

- For UE have non-asymptotic expression for MSE-work product, depending on a_i and $\mathbb{P}(N \geq i)$. Optimize by minimizing wrt a_i and $\mathbb{P}(N \geq i)$: **hard!**

Comparison of Unbiased Estimator vs Ergodic Average

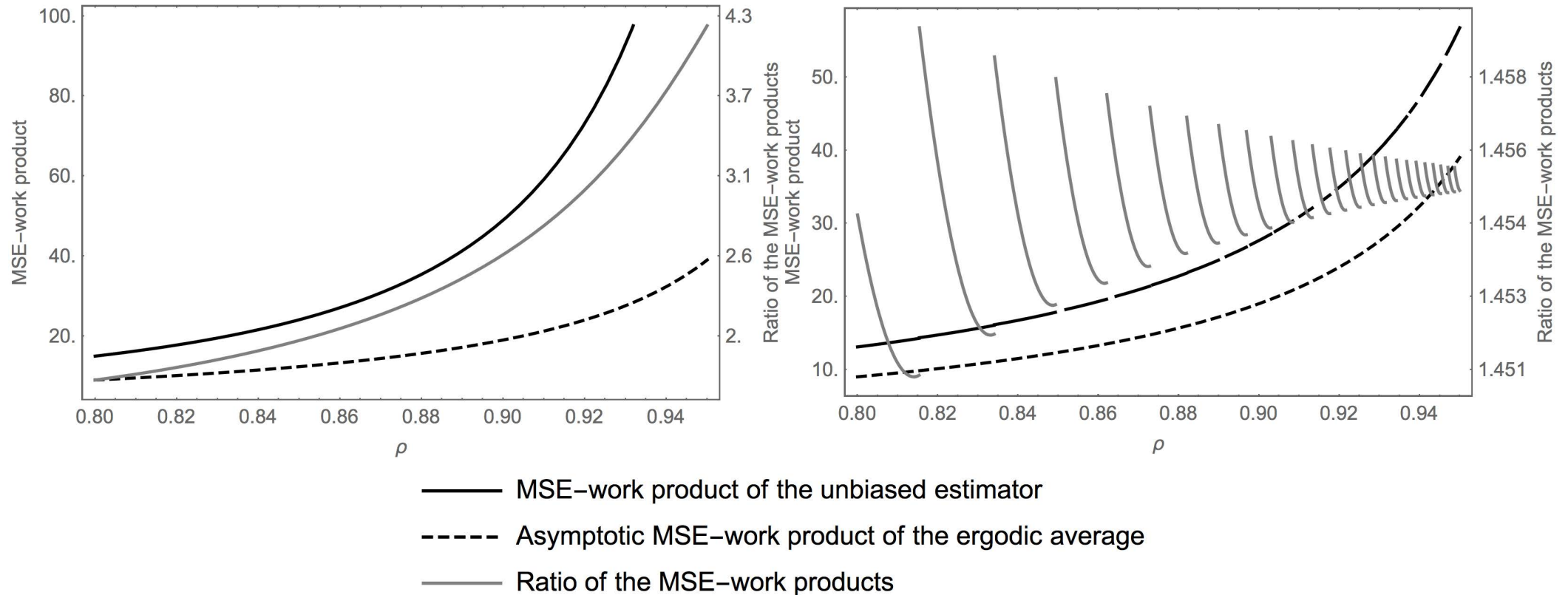







Figure: 1) Left: optimized $\mathbb{P}(N \geq i)$ for fixed $a_i = 4(i + 1)$ (as in [GR13](#)), 2) Right: optimized $\mathbb{P}(N \geq i)$ and a_i over subclass $a_i = m(i + 1)$.

Conclusions - further work

- UE is often feasible.
- Optimization wrt parameters is **crucial** especially in function space setting.
- UE easily **parallelizable**: a) use independent copies of Z , b) Δ_i 's independent.
- UE seems competitive. Looking forward to comparisons in problems of higher complexity (e.g. BIP-UQ example).

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