# Quantifying Uncertainty in Differential Equation Models: Manifolds, Metrics and Russian Roulette 

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Enabling Quantification of

Uncertainty for Inverse Problems

## Talk Outline

- Why statistical inference for mechanistic models is hugely challenging


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- How to account for uncertainty induced by implicit definition of dynamics via PDE and ODE representation
- Oftentimes under fine spatial mesh refinement computing a likelihood exactly may be infeasible
- Forward and Inverse inference can still progress exploiting pseudo-marginal constructions in general form of Russian Roulette


## Simple Dynamics

$$
\begin{aligned}
& \frac{d x}{d t}=\theta_{1} y z \\
& \frac{d y}{d t}=-x z \\
& \frac{d z}{d t}=-\theta_{2} x y
\end{aligned}
$$



## Simple Dynamics



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\frac{d x}{d t} & =\frac{72}{36+y}-\alpha \\
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- Induces a data density posing many challenges for simulation based inference


## Systems Identification - Posterior Inference



## Mixing of Markov Chains



## Geometric Concepts in MCMC

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- Rao, 1945 to first order

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\chi^{2}(\delta \boldsymbol{\theta})=\int \frac{|p(\mathbf{y} ; \boldsymbol{\theta}+\delta \boldsymbol{\theta})-p(\mathbf{y} ; \boldsymbol{\theta})|^{2}}{p(\mathbf{y} ; \boldsymbol{\theta})} d \mathbf{y} \approx \delta \boldsymbol{\theta}^{\top} \mathbf{G}(\boldsymbol{\theta}) \delta \boldsymbol{\theta}
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\boldsymbol{\theta}_{d}^{\prime}=\boldsymbol{\theta}_{d}+\frac{\epsilon^{2}}{2}\left(\boldsymbol{G}^{-1}(\boldsymbol{\theta}) \nabla_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\theta})\right)_{d}-\epsilon^{2} \sum_{i, j}^{D} \boldsymbol{G}(\boldsymbol{\theta})_{i j}^{-1} \Gamma_{i j}^{d}+\epsilon\left(\sqrt{\mathbf{G}^{-1}(\boldsymbol{\theta})} \mathbf{z}\right)_{d}
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- Non-Euclidean geometry can exploit geodesics equations to devise sampling schemes (RMHMC)


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## Infinite and finite dimensional mismatch

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- Observations, $\mathbf{y}(t) \in \mathbb{R}^{R}$, via transformation, $\mathcal{G}: \mathbb{R}^{P} \rightarrow \mathbb{R}^{R}$, of the true or exact solution, $\mathbf{x}(t, \theta)$, of system at $T$ time points.


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\end{aligned}
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where $\delta^{(r)}(\mathbf{t})=\mathcal{G}^{(r)}(\mathbf{x}(\mathbf{t}))-\mathcal{G}^{(r)}\left(\mathbf{x}^{N}(\mathbf{t})\right)$.

## Full Bayesian posterior measure for model uncertainty

- Define a probability measure over functions

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- Uncertainty in the probabilistic solution $\mathbf{x}(\mathbf{t}, \boldsymbol{\theta})$ is made explicit taking into account the mismatch between state solution and a finite approximation


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- Radon-Nikodym derivative of posterior measure with respect to GP prior

$$
\frac{d \mu^{f}}{d \mu_{0}^{f}}(\dot{x}(\mathbf{s})) \propto \exp \left(-\frac{1}{2}\left\|\dot{x}(\mathbf{s})-\mathbf{f}_{1: N}\right\|_{\Lambda_{N}}^{2}\right)
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- Gaussian measure, $\mu_{0}^{f}=\mathcal{N}\left(m_{0}^{f}, \mathcal{C}_{0}^{f}\right)$, on a Hilbert space, $\mathcal{H}$, mean function $m_{0}^{f}$, covariance operator $\mathcal{C}_{0}^{f}$ well defined

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- Eigenfunctions of covariance operator form basis for derivative space
- $m_{0}^{f}\left(t_{1}\right)=\ell\left(t_{1}\right), \quad \mathcal{C}_{0}^{f}\left(t_{1}, t_{2}\right)=\alpha^{-1} \int_{\mathbb{R}} \mathrm{R}_{\lambda}\left(t_{1}, z\right) \mathrm{R}_{\lambda}\left(t_{2}, z\right) \mathrm{d} z=\mathrm{RR}\left(t_{1}, t_{2}\right)$.


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- Posterior measure for $\dot{x}(\mathbf{t})$ denoted by $\mu_{n}^{f}=\mathcal{N}\left(m_{n}^{f}(\mathbf{t}), \mathcal{C}_{n}^{f}(\mathbf{t}, \mathbf{t})\right)$


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- Suggests probabilistic construction (integration), sampling of solutions


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- Fifteen solution samples illustrate uncertainty over domain propagates through system resulting in noticeably distinct dynamics, not captured by deterministic numerical solvers.


## Kuramoto-Sivashinsky model of reaction-diffusion



Figure: Side view and top view of a probabilistic solution realization of the Kuramoto-Sivashinsky PDE with initial function $u(0, x)=\cos (x / 16)\{1+\sin (x / 16)\}$ and domain $x \in[0,32 \pi], t \in[0,150]$.

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Figure: Fifteen realizations of the probabilistic solution of the Kuramoto-Sivashinsky PDE using a fixed initial function. The solution is known to exhibit temporal chaos. Deterministic numerical solutions only capture one type of behaviour given a fixed initial function, which can lead to bias when used in conjunction with data-based inference.

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## Full Bayesian Uncertainty Quantification

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Propose $\boldsymbol{\theta}^{\star} \sim q\left(\boldsymbol{\theta}^{\star} \mid \boldsymbol{\theta}\right)$;
Conditionally simulate a solution realisation $\mathbf{x}^{\star}(\mathbf{t})$ from $p\left(\mathbf{x}(\mathbf{t}), \mathbf{f}_{1: N} \mid \boldsymbol{\theta}, \mathbf{x}_{0}, \Psi\right)$ Compute:

$$
\rho\left(\boldsymbol{\theta}, \mathbf{x}(\mathbf{t}) \rightarrow \boldsymbol{\theta}^{\star}, \mathbf{x}^{\star}(\mathbf{t})\right)=\frac{q\left(\boldsymbol{\theta} \mid \boldsymbol{\theta}^{\star}\right)}{q\left(\boldsymbol{\theta}^{\star} \mid \boldsymbol{\theta}\right)} \frac{p\left(\boldsymbol{\theta}^{\star}\right)}{p(\boldsymbol{\theta})} \frac{p\left(\mathbf{y}(\mathbf{t}) \mid \mathcal{G}\left(\mathbf{x}^{\star}(\mathbf{t})\right), \Sigma\right)}{p(\mathbf{y}(\mathbf{t}) \mid \mathcal{G}(\mathbf{x}(\mathbf{t})), \Sigma)}
$$

if $\min \left[1, \rho\left(\boldsymbol{\theta} \rightarrow \boldsymbol{\theta}^{\star}\right)\right]>\mathrm{U}[0,1]$ then
Update $\boldsymbol{\theta}, \mathbf{x}(\mathbf{t})=\boldsymbol{\theta}^{\star}, \mathbf{x}^{\star}(\mathbf{t})$;
end if
Return $\boldsymbol{\theta}, \mathbf{x}(\mathbf{t})$.
end for

## Inference for model of cellular signal transduction



Figure: Experimental data and sample paths of the observation processes obtained by transforming a sample from marginal posterior state distribution by observation function

## Inference for model of cellular signal transduction



Figure: Marginal parameter posterior based on sample of size 100 K generated by a parallel tempering algorithm utilizing seven chains, with the first 10K samples removed. Prior densities are shown in red.

## Intractable Likelihoods under Mesh Refinement

What can we do?

## Large Scale GMRF Ozone Column Model



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p(\mathbf{x} \mid \theta)=\frac{1}{(2 \pi)^{\frac{N}{2}}} \exp \left\{-\frac{1}{2}\left[\log \left|\mathbf{C}_{\theta}\right|+\mathbf{x}^{\top} \mathbf{C}_{\theta}^{-1} \mathbf{x}\right]\right\}
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& =\frac{1}{(2 \pi)^{\frac{N}{2}}} \exp \left\{\frac{1}{2} \sum_{n=1}^{\infty} \frac{1}{n} \mathrm{E}\left\{\mathbf{z}^{\top}\left(\mathbf{I}-\mathbf{C}_{\theta}\right)^{n} \mathbf{z}\right\}-\frac{1}{2} \sum_{m=0}^{\infty} \mathbf{x}^{\top}\left(\mathbf{I}-\mathbf{C}_{\theta}\right)^{m} \mathbf{x}\right\}
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$$

- Exploit Pseudo-Marginal construction - Andrieu \& Roberts, 2009 Russian Roulette unbiased truncation of infinite series - MCMC based inference can proceed..... in principle ;-)


## Conclusions and Discussion

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