# A Decision Theoretic Approach to Optimize Clinical Trial Designs for Targeted Therapies

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### Full Population F



- Assume the treatment effects in the subgroups satisfy  $\delta_S \geq \delta_{S'}.$
- The treatment effect in F is

$$\delta_F = \lambda_S \delta_S + (1 - \lambda_S) \delta_{S'}$$

where  $\lambda_S$  is the prevalence of subgroup *S*.

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Recruitment from population F. No Biomarker is determined. Test of  $H_F$ .

#### Stratification Design:

Recruitment from population F. Stratified randomization by Biomarker. Test of  $H_F$  and  $H_S$ .

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Recruitment only from population *S* . Test of *H<sub>S</sub>* 



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 $H_F$  is tested with a z-test.

**Stratification Design:** 

H<sub>S</sub> and H<sub>F</sub> are tested with a closed
 Spiessens-Debois (2010) test at levels α<sub>S</sub>, α<sub>F</sub>.
 If a hypothesis is rejected, the other is tested at level α.

• To reject  $H_F$ , also the consistency condition

 $p_S \leq \tau_S$  and  $p_{S'} \leq \tau_{S'}$ ,

for parameters  $\tau_S, \tau_{S'}$ , must be satisfied.

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**Enrichment Design:** 

# Optimizing Clinical trial designs

- When is a biomarker (BM) design beneficial compared to a classical design?
- When to choose stratified, when an enrichment design?
- Which sample size?
- Which significance levels α<sub>F</sub> and α<sub>S</sub> for H<sub>F</sub> and H<sub>S</sub> in the weighted multiple test for the stratified design are optimal?

We apply a utility based approach, (cf. Beckman et al., 2011; Rosenblum et al., 2014; Graf et al., 2015), to model the expected utilities of a particular trial design from a sponsor's and a public health view.

$$U(d) = -C(d) + \begin{cases} \varphi_{F,d} & \text{if } H_F \text{ is rejected} \\ \varphi_{S,d} & \text{if only } H_S \text{ is rejected} \\ 0 & \text{if no hypothesis is rejected} \end{cases}$$

#### Sponsor view

$$\varphi_{F,d} = \mathbf{N} \cdot \mathbf{r}_F \cdot (\hat{\delta}_{F,d} - \mu_F)^+$$
$$\varphi_{S,d} = \lambda_S \cdot \mathbf{N} \cdot \mathbf{r}_S \cdot (\hat{\delta}_{S,d} - \mu_S)^+$$

- *N* . . . number of future patients (market size).
- *r<sub>F</sub>*, *r<sub>S</sub>*... revenue parameters.
- $\hat{\delta}_{F,d}, \hat{\delta}_{S,d} \dots$  efficacy estimates.
- $\mu_F, \mu_S \ldots$  clinically relevant thresholds.

Public health view

$$\varphi_{F,d} = N \cdot r_F \cdot (\delta_F - \mu_F)$$
  
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 $c_{\text{setup}} + 2n c_{\text{per-patient}}.$ 

Stratified Design

 $c_{\text{setup}} + c_{\text{BM development}} + 2n(c_{\text{per-patient}} + c_{\text{BM determination}}).$ 

Enrichment Design

 $c_{\text{setup}} + c_{\text{BM development}} + 2n(c_{\text{per-patient}} + \frac{c_{\text{BM determination}}}{\lambda_S}).$ 

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# The Expected Utility of Trial Designs in the Planning Stage Identifying optimized trial designs

Expected Utility:

$$E_{\pi} \{ E_{\Delta}[U(d)] \}$$

The expectation is taken over

- the prior  $\pi$  on the effect sizes  $\Delta = (\delta_S, \delta_{S'})$  and
- the sampling distribution

Optimal design: Choose the design with maximal expected utility optimizing over

- Type of design (classical/stratified/enrichment)
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# Prior Distributions $\pi$ On the Effects $\delta_{S}, \delta_{S'}$

$\delta_S$	0	$\theta$	$\theta$	$\theta$
δς'	0	0	$\theta/2$	$\theta$
Weak Biomarker Prior	0.2	0.2	0.3	0.3
Strong Biomarker Prior	0.2	0.6	0.1	0.1

 $\theta > 0 \dots$  effect size parameter.

# Scenario

# • Effect size parameter in the prior $\theta = 0.3$

- Reward parameters  $Nr_F = Nr_S = 1000MUSD$  $\mu_F = \mu_S = 0.1.$
- Cost Parameters in (MUSD)
  - $c_{\text{setup}} = 1$
  - $c_{\text{per-patient}} = 0.05$
  - $c_{\rm BM \ development} = 1$
  - $c_{\text{BM determination}} = 0.005.$
- Consistency parameters  $\tau_S = \tau_{S'} = 0.3$ .

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Optimized Expected Utilities Weak Biomarker Prior





Optimized Alpha Allocation Weak Biomarker Prior



# Optimal Designs for the Strong Biomarker Prior



#### Optimal Trial Designs Weak Biomarker Prior



#### Optimal Trial Designs Strong Biomarker Prior



- The optimal sample size under the public health view is typically larger than in the sponsor view.
- For the considered priors, the enrichment design is never optimal for the sponsor view
- The optimal design depends strongly on the particulars of the situation: Subgroup prevalence, trial costs and initial beliefs.

# Two Extensions of the Trial Designs

Partial Enrichment Design

The prevalence of the subgroup in the trial  $\lambda_T$  is a design parameter and may differ from  $\lambda_S$ , the prevalence in the population. Special cases are the stratified design ( $\lambda_T = \lambda_S$ ) and the (full) enrichment design ( $\lambda_T = 1$ ).

E.g., Zhao et al. (2010)

#### Adaptive Enrichment Designs

Two stage design, where the second stage sample size and second stage trial prevalence may depend on the first stage outcome.

E.g., Brannath et al. (2009); Beckman et al. (2011); Jenkins et al. (2011); Friede

et al. (2012); Simon and Simon (2013); Krisam and Kieser (2015); Graf et al.

(2015); Fisher and Rosenblum (2016)

For simplicity we use as multiple testing procedure a single step unweighted Bonferroni test. Because trial and population prevalence do not match, the standard z-test is not a valid test for  $H_F$ .

• *H<sub>F</sub>* is tested with a reweighted z-statistics

$$\tilde{Z}_{F} = \xi \left( \frac{\lambda_{S}}{\sqrt{\lambda_{T}}} Z_{S} + \frac{1 - \lambda_{S}}{\sqrt{1 - \lambda_{T}}} Z_{S'} \right),$$

where  $Z_S, Z_{S'}$  denote the z-statistics for the subgroups S, S' and  $\xi$  is a normalizing constant.

e.g., Zhao et.al. 2010

• As above, to reject  $H_F$ , in addition the consistency condition needs to be fulfilled.





# Optimal Subgroup Prevalence $\lambda_{\mathcal{T}}$ Weak Biomarker Prior





First stage

Sample size  $n_1$ , subgroup trial prevalence  $\lambda_T^1$ Second Stage Second stage sample size  $n_2$  and subgroup trial

Second stage sample size  $n_2$  and subgroup trial prevalence  $\lambda_T^2$  are chosen based on first stage data.

## Testing procedure

• Overall test statistics computed with combination function:

$$Z_{S} = \sqrt{\frac{1}{2}} Z_{S}^{1} + \sqrt{\frac{1}{2}} Z_{S}^{2}$$
$$\tilde{Z}_{F} = \sqrt{\frac{1}{2}} \tilde{Z}_{F}^{1} + \sqrt{\frac{1}{2}} \tilde{Z}_{F}^{2}$$

where  $Z_S^1, Z_S^2$  and  $\tilde{Z}_F^1, \tilde{Z}_F^2$  are stage wise z-statistics.

• Unweighted Bonferroni test boundaries applied to  $Z_S$  and  $\tilde{Z}_F$  (if  $\lambda_T^2 < 1$ ).

## **Optimal Adaptation Rule**

- In the interim analysis choose  $n_2$ ,  $\lambda_T^2$  such that the expected utility conditional on the first stage data is maximized.
- Especially,  $n_2 = 0$  corresponds to stopping for futility,  $\lambda_T = 1$  to a second stage enrichment design.

## Optimizing first stage parameters

Choose n<sub>1</sub>, λ<sup>1</sup><sub>T</sub> such that the expected utility (given the optimal adaptation rule is applied at interim) is maximized.

#### Example for the Optimal Decision Rule Weak Biomarker Prior ( $n_1 = 100, \lambda_S = \lambda_T^1 = 0.5$ )







- Partial Enrichment Designs can increase the utility mainly for the sponsor utility function.
- Adaptive Enrichment Designs further increase the expected utility, also for the public health utility function.
- Extensions: weighted, stepwise Spiessens Debois test for the partial enrichment design, optimized weights in combination function...

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# Backup Slides

# Optimal Sample Size of the Partial Enrichment Design Weak Biomarker Prior



#### Example for the Optimal Decision Rule Weak Biomarker Prior ( $n_1 = 100, \lambda_S = \lambda_T^1 = 0.5$ )



#### Example for the Optimal Decision Rule Strong Biomarker Prior $(n_1 = 100, \lambda_S = \lambda_T^1 = 0.5)$

