Optimizing Confirmatory Clinical Trial Designs for Targeted Therapies A Decision Theoretic Approach

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Joint work with Thomas Ondra, Sebastian Jobjörnsson, Robert Beckman, Carl-Fredrik Burman, Franz König, Nigel Stallard







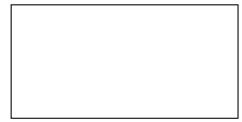
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For the development of targeted therapies, clinical trials with complex objectives, confirming treatment effects in sub-populations and/or in the overall populations are required.

- Knowledge on the genetic basis of many diseases is rapidly increasing and therapies that target underlying molecular mechanisms are developed.
- Patients' responses are predicted to targeted treatments based on genetic features or other biomarkers.

Objective: Identify subgroups (based on biomarkers) where the treatment has a positive benefit risk balance.

Full Population F

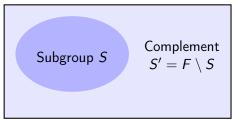


• The overall treatment effect is

$$\delta_F = \lambda_S \delta_S + (1 - \lambda_S) \delta_{S'}$$

- Assume $\delta_{S'} \leq \delta_S$.
- Test of hypotheses $H_F : \delta_F \leq 0$ and $H_S : \delta_S \leq 0$.

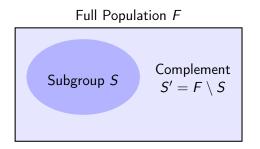
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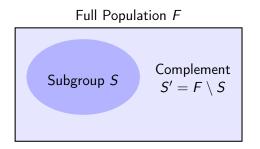
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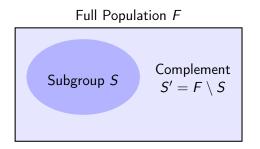
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Recruitment from population F. No Biomarker is determined. Test of H_F .

Stratification Design:

Recruitment from population F. Stratified randomization by Biomarker. Test of H_F and H_S .

Enrichment Design:

Recruitment only from population *S* . Test of *H_S*



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- When is a biomarker (BM) design beneficial compared to a classical design?
- When to choose stratified, when an enrichment design?
- Which sample size?
- Which multiple test for the stratified design is optimal?

Perform a Trial in the Full or Only the Subpopulation?

- The power to reject at least one hypothesis depends on the effect sizes $\delta_S, \delta_{S'}$ and the prevalences
- Assume we suspect that $\delta_{S'} \leq \delta_S$ but believe that $\delta_{S'} > \delta_S$ is not plausible.
- Then, the design in the subpopulation (recruiting only patients in *S*) always leads to the highest power:
 - If $\delta_{S'} = \delta_S$ the enrichment design has a larger power than the stratification design (which requires multiple testing) and equal power as the classical design.
 - If $\delta_{S'} < \delta_S$ the enrichment design has larger power than stratification & classical design.

Thus, is enrichment always preferable?

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Different Gains, different Costs...

The argument results from an oversimplification. The Power to reject any null hypotheses is not the only criteria.

- The stratification design tests the full population H_F : δ_F ≤ 0, demonstrating that the treatment works "on average".
- The enrichment design tests a limited null hypothesis $H_S: \delta_S \leq 0$ leading to a limited indication.
- Ethical problem if patients that may benefit are excluded.
- Enrichment maybe costly (e.g. due to longer patient recruitment, ...).

We apply a utility based approach, (cf. Beckman et al., 2011; Graf et al., 2015), to model the expected utilities of a particular trial design from a sponsor's and a public health view.

$$U(d) = -C(d) + \begin{cases} \varphi_{F,d} & \text{if } H_F \text{ is rejected} \\ \varphi_{S,d} & \text{if only } H_S \text{ is rejected} \\ 0 & \text{if no hypothesis is rejected} \end{cases}.$$

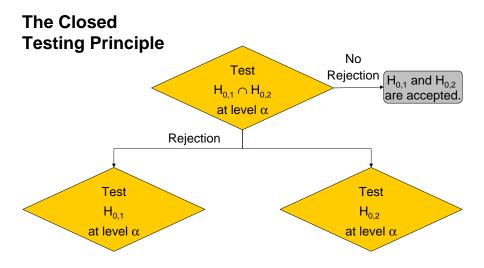
where, $d \in D$ denotes a trial design (Classical, Stratification, Enrichment) together with a specific sample size, and testing procedure.

- In the stratification design two hypotheses H_F , H_S are tested.
- Regulatory authorities require the control of the Familywise Error Rate defined as

The probability to reject at least one true null hypothesis.

• A general, powerful approach is the Closed Testing Principle.

Marcus, R; Peritz, E; Gabriel, KR (1976)



- Let p_S , p_F denote p-values for z-tests of H_F and H_S
- The Spiessens and Debois test rejects the intersection hypothesis $H = H_S \cap H_F$ if

 $p_S \leq \alpha_S$ or $p_F \leq \alpha_F$

- If the intersection hypothesis is rejected, the elementary hypotheses are tested at level α .
- α_F, α_S are chosen such that

 $\mathbb{P}_{H_F \cap H_S}(p_F \leq \alpha_F \text{ or } p_S \leq \alpha_S) = \alpha.$

- The correlation of p_S, p_F and therefore also the levels α_S, α_F depend on λ .
- For fixed α , λ_S and α_S , the level α_F is determined by the above equation.
- If the prevalence increases, the correlation increases and for α_S fixed, α_F increases.

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- H_S and H_F are tested with a closed Spiessens-Debois (2010) test at levels α_S, α_F.
 If a hypothesis is rejected, the other is tested at level α.
- To reject H_F , also the consistency condition

 $p_S \leq \tau_S$ and $p_{S'} \leq \tau_{S'}$,

for parameters $\tau_S, \tau_{S'}$, must be satisfied.

Classical Design:

 H_F is tested with a z-test.

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Sponsor view

$$\varphi_{F,d} = \mathbf{N} \cdot \mathbf{r}_F \cdot (\hat{\delta}_{F,d} - \mu_F)^+$$
$$\varphi_{S,d} = \lambda_S \cdot \mathbf{N} \cdot \mathbf{r}_S \cdot (\hat{\delta}_{S,d} - \mu_S)^+$$

- *N* . . . number of future patients (market size).
- *r_F*, *r_S* ... revenue parameters.
- $\hat{\delta}_{F,d}, \hat{\delta}_{S,d} \dots$ efficacy estimates.
- $\mu_F, \mu_S \ldots$ clinically relevant thresholds.

Public health view

$$\varphi_{F,d} = N \cdot r_F \cdot (\delta_F - \mu_F)$$

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The Rewards

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• $\delta_S, \delta_F \dots$ true effect sizes.

• Classical Design

 $c_{\text{setup}} + 2n c_{\text{per-patient}}.$

Stratified Design

 $c_{\text{setup}} + c_{\text{BM development}} + 2n(c_{\text{per-patient}} + c_{\text{BM determination}}).$

• Enrichment Design

 $c_{\text{setup}} + c_{\text{BM development}} + 2n(c_{\text{per-patient}} + \frac{c_{\text{BM determination}}}{\lambda_S}).$

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Expected Utility:

$E_{\pi}\left\{ E_{\Delta}[U(d)]\right\}$

The expectation is taken over

- the prior π on the effect sizes $\Delta = (\delta_{\mathcal{S}}, \delta_{\mathcal{S}'})$ and
- the sampling distribution

Optimal design: Choose the design with maximal expected utility with

- Optimal type of design (classical/stratified/enrichment)
- Optimal sample size
- Optimal α allocation (for the stratified design)

Prior Distributions π on the Effect Sizes $\delta_{\mathcal{S}}, \delta_{\mathcal{S}'}$

δ_S		θ	θ	θ
δς'	0	0	$\theta/2$	θ
Weak Biomarker Prior	0.2	0.2	0.3	0.3
Strong Biomarker Prior	0.2	0.6	0.1	0.1

where $\theta > 0$ is an effect size parameter.

Scenario

• Effect size parameter in the prior $\theta = 0.3$

- Reward parameters $Nr_F = Nr_S = 1000MUSD$ $\mu_F = \mu_S = 0.1.$
- Cost Parameters in (MUSD)
 - $c_{\text{setup}} = 1$
 - $c_{\text{per-patient}} = 0.05$
 - $c_{\rm BM \ development} = 1$
 - $c_{\text{BM determination}} = 0.005.$
- Consistency parameters $\tau_S = \tau_{S'} = 0.3$.

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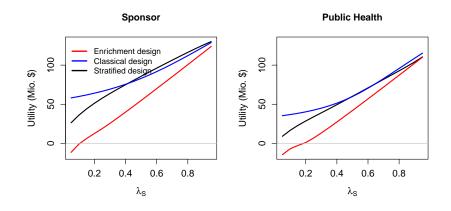


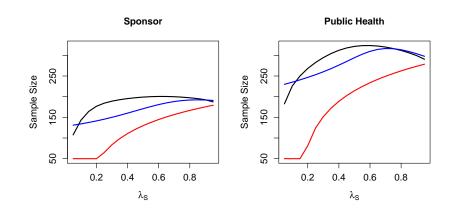
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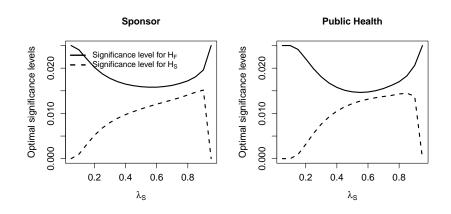
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Optimized Expected Utilities Weak Biomarker Prior



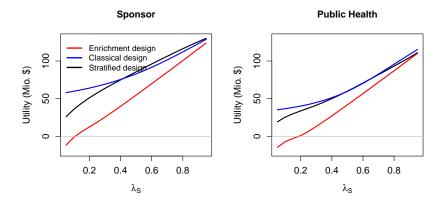


Optimized Alpha Allocation Weak Biomarker Prior



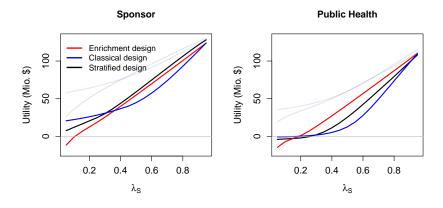
Optimized Expected Utilities - Impact of the Prior

Weak Biomarker Prior

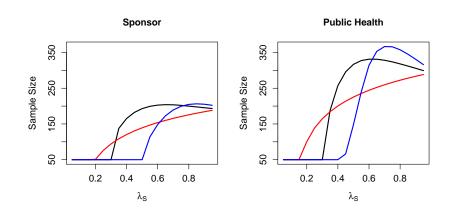


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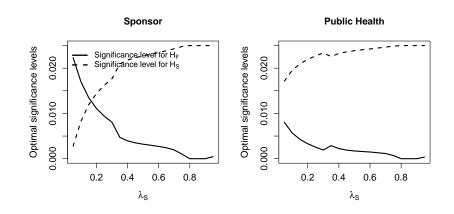
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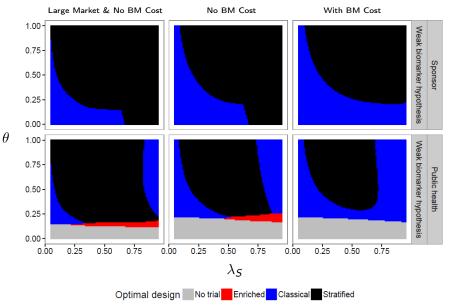
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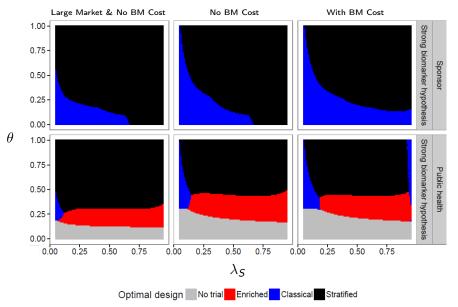
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Optimal Trial Designs Weak Biomarker Prior



Optimal Trial Designs Strong Biomarker Prior



Some General Observations and Conclusion

- The decision theoretic model can inform the choice of
 - the type of trial design,
 - the sample size and the weights in the multiple test.
- The optimal sample size under the public health view is typically larger than in the sponsor view.
- The enrichment design is never optimal for the sponsor view
- For some scenarios the optimized designs differ, but the expected utilities of different design options are often small.
- The optimal design depends strongly on the particulars of the situation: Subgroup prevalence, trial costs and initial beliefs.

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